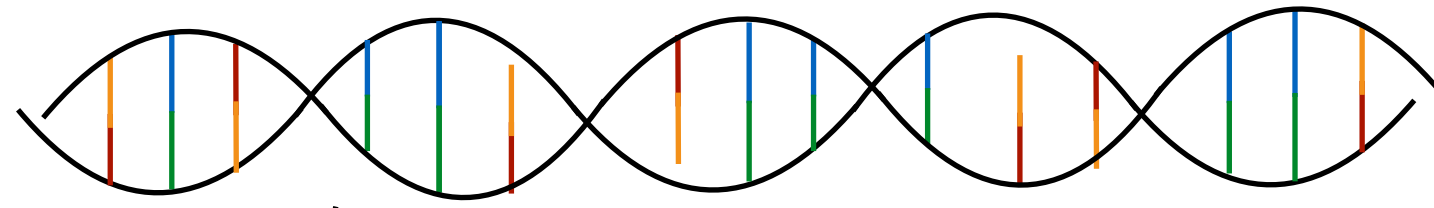


Analyzing gene and transcript expression using RNA-seq

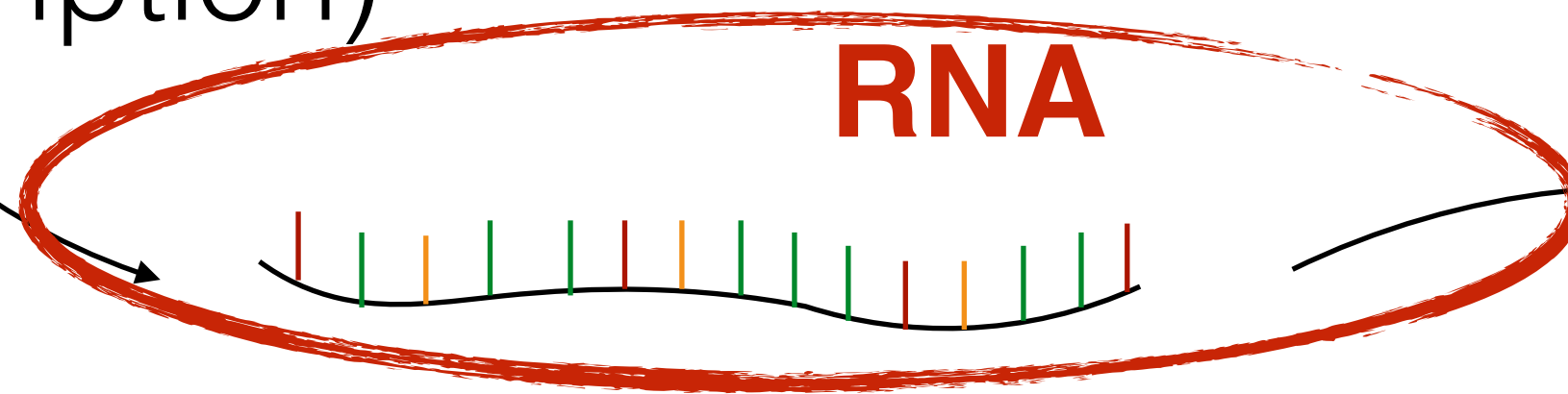
“Flow” of information in the cell

DNA



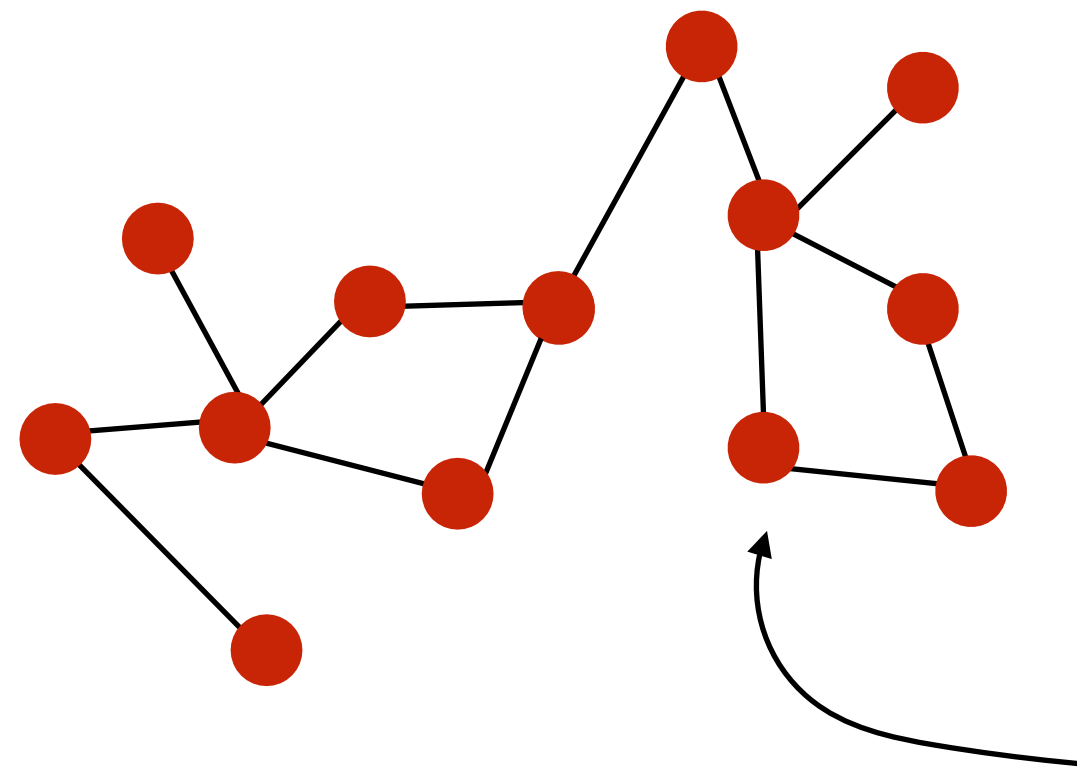
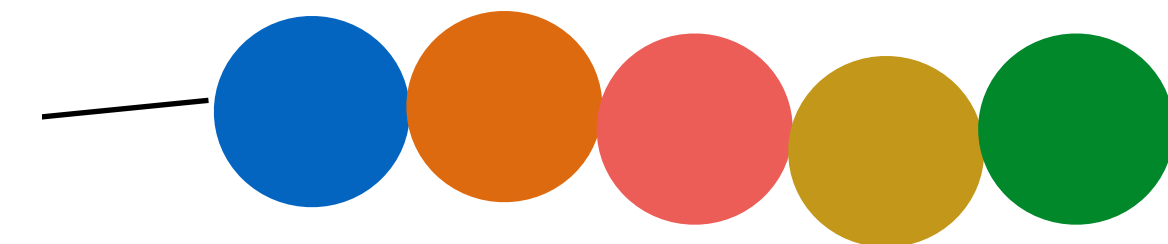
RNA Polymerase
(transcription)

RNA



Ribosomes
(translation)

Protein



Form networks &
pathways; perform a
vast set of cellular
functions

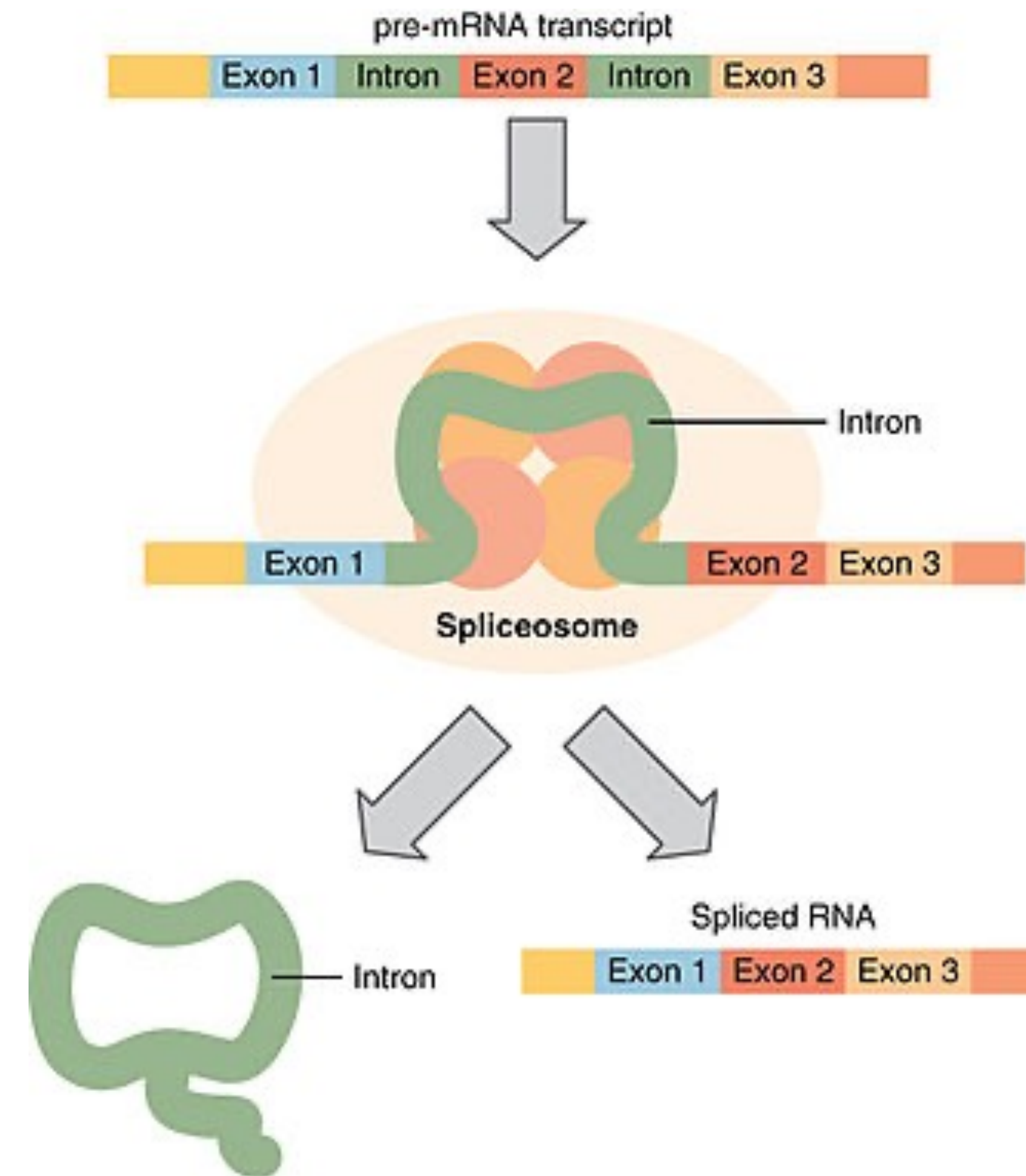
RNA Splicing

DNA transcribed into pre-mRNA

Some “processing occurs”
capping & **polyadenylation**

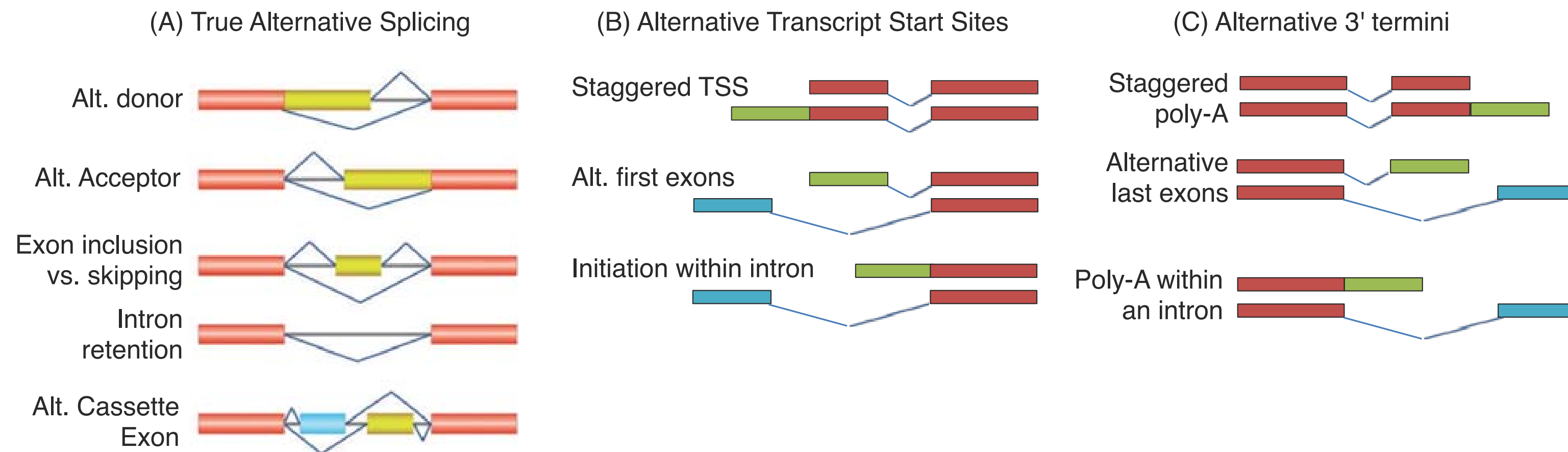
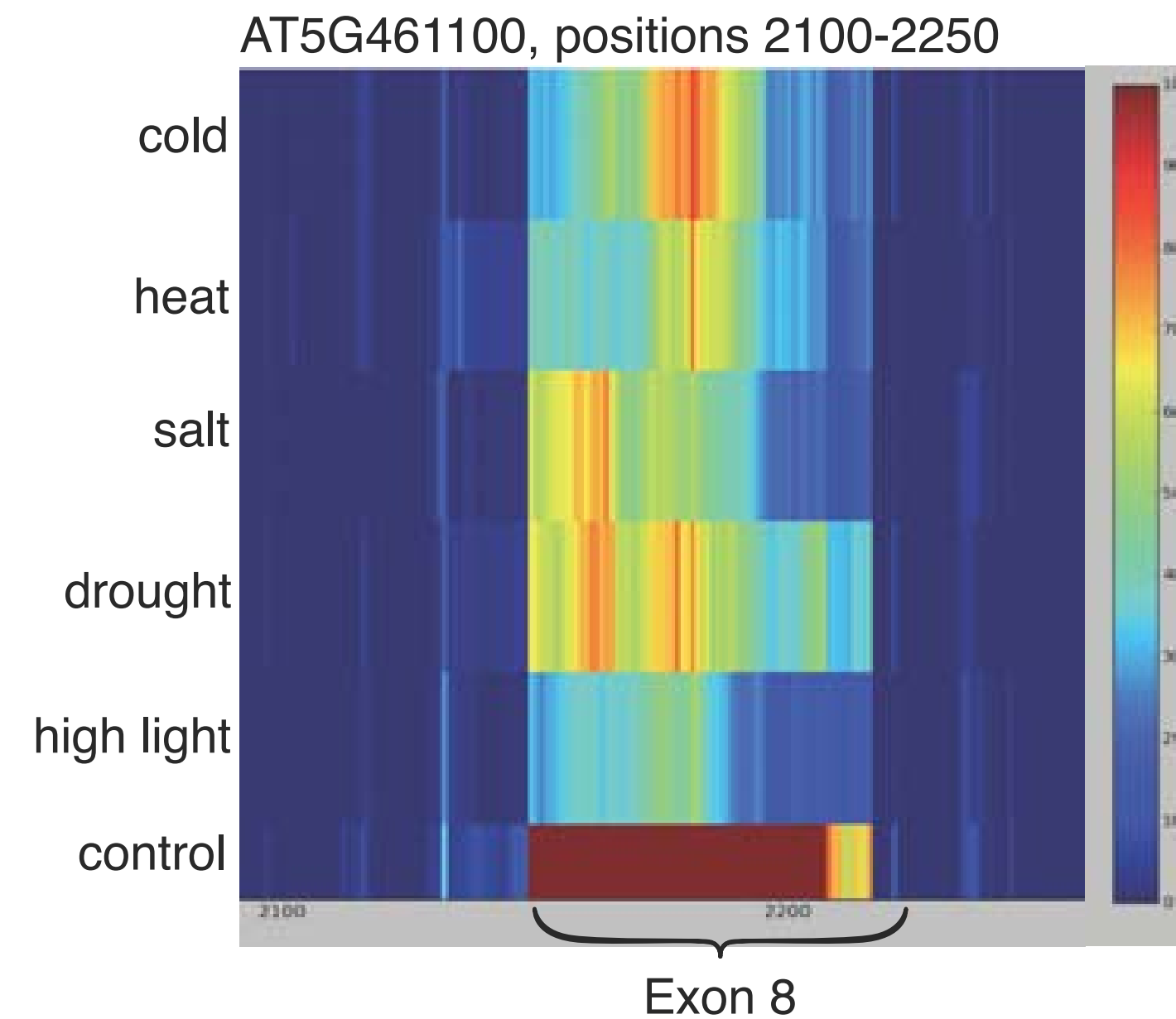
Introns removed from pre-mRNA

Introns removed resulting in
mature mRNA

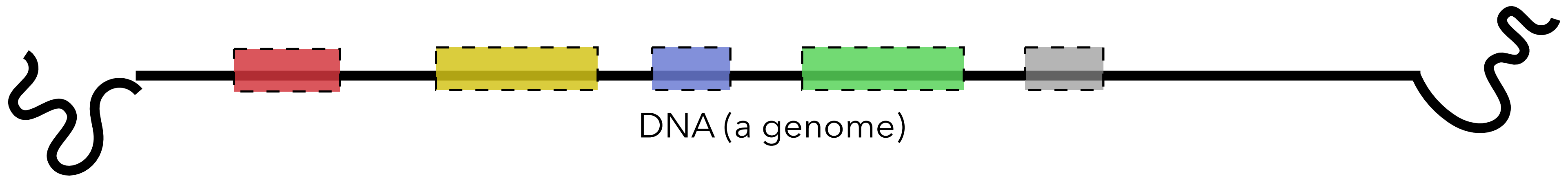


Alternative Splicing & Isoform Expression

- Expression of genes can be measured via RNA-seq (sequencing transcripts)
- Sequencing gives you short (35-300bp length reads)

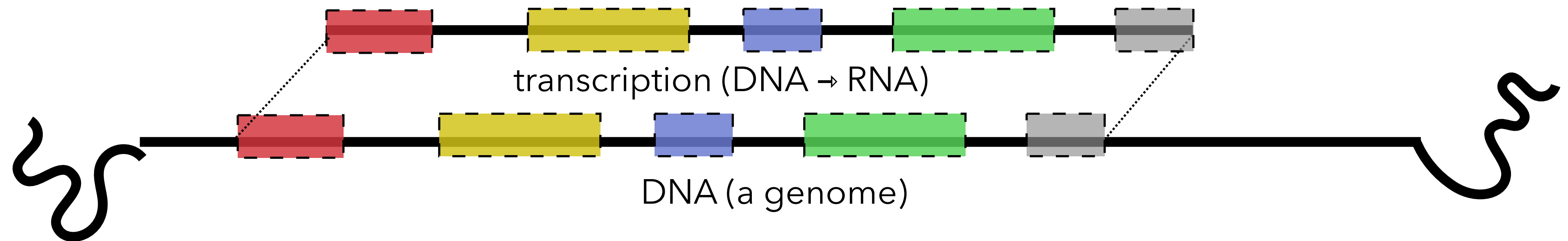


What is RNA sequencing



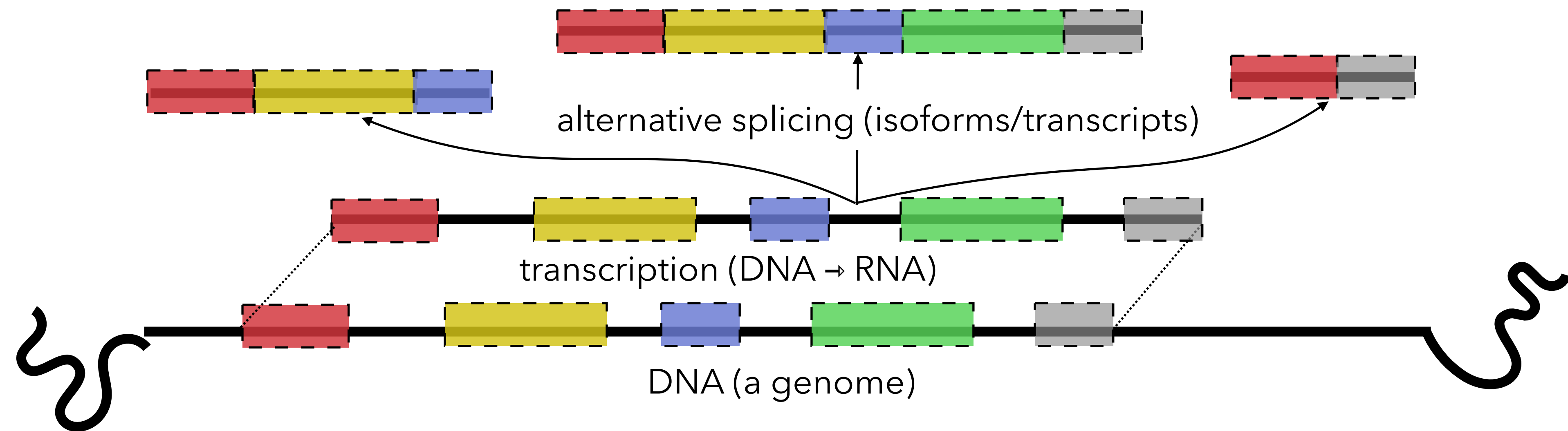
* most protocols actually sequence complementary DNA (cDNA), not RNA directly

What is RNA sequencing



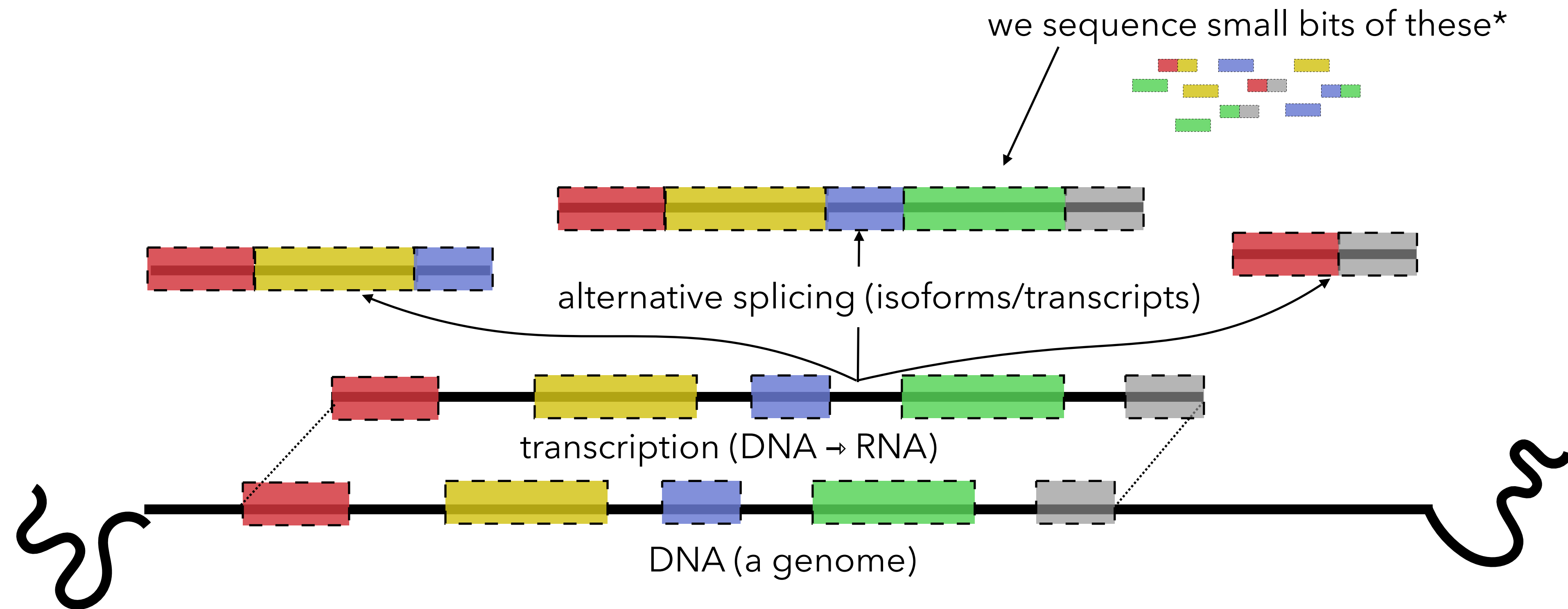
* most protocols actually sequence complementary DNA (cDNA), not RNA directly

What is RNA sequencing



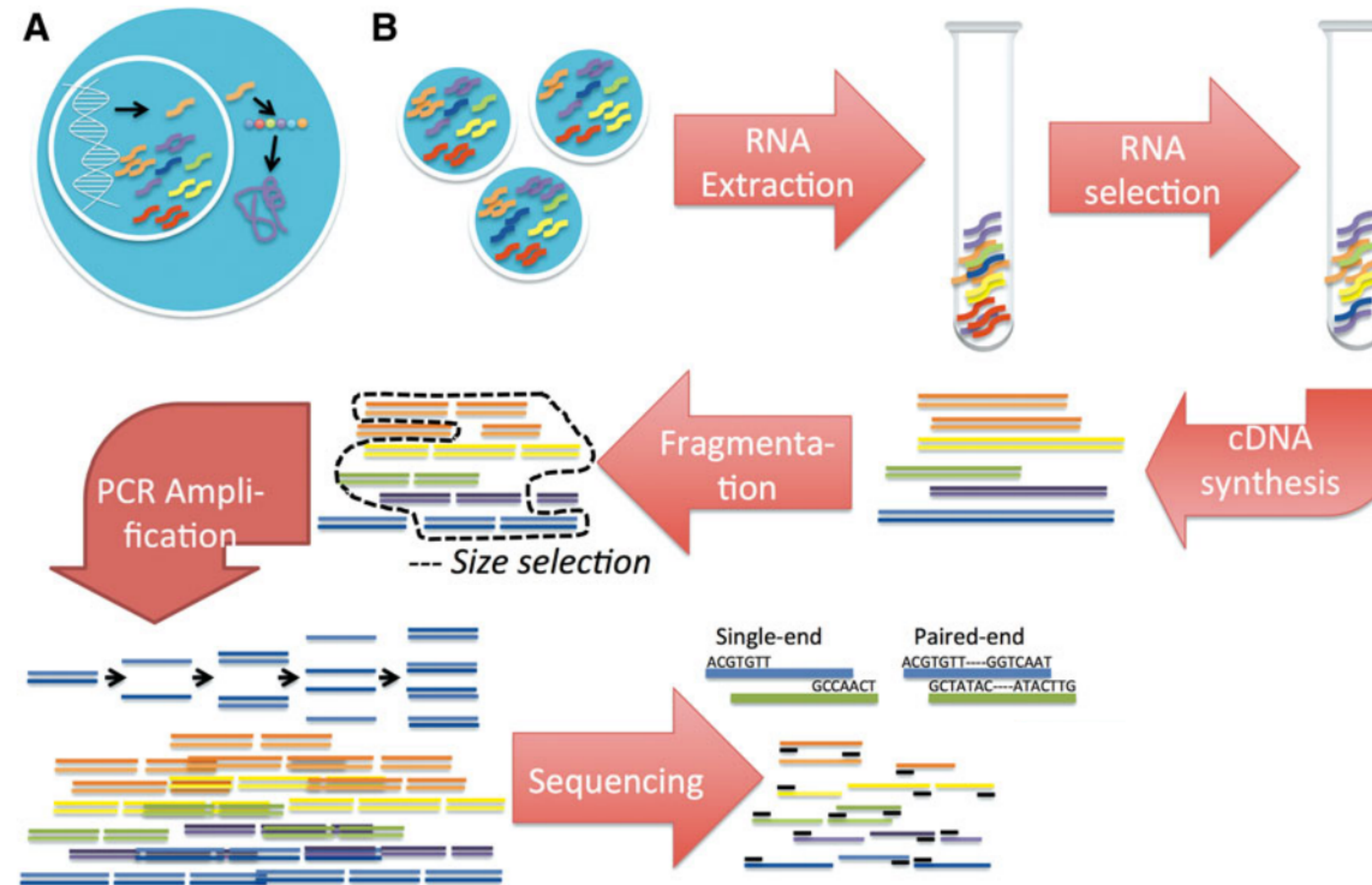
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What is RNA sequencing

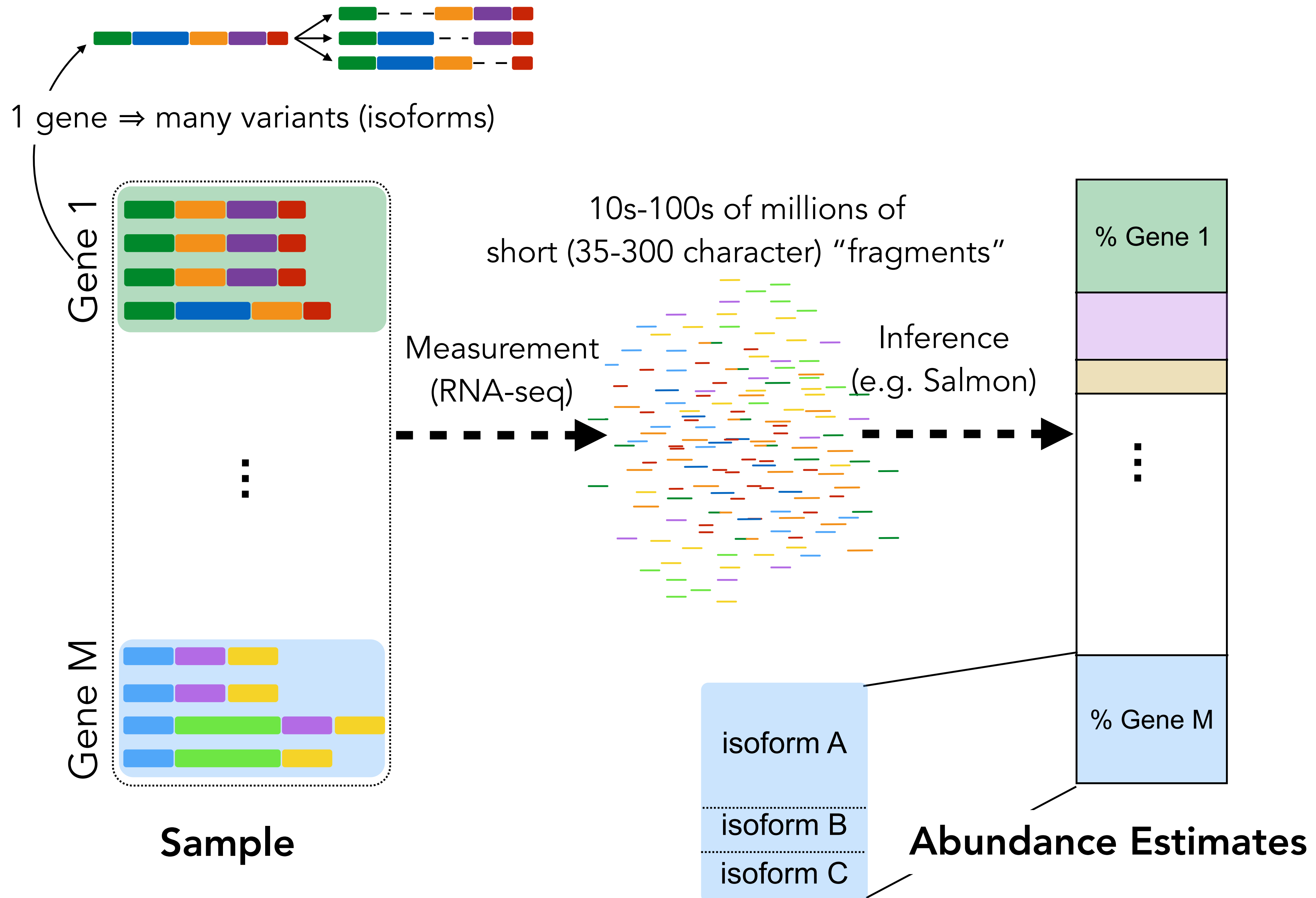


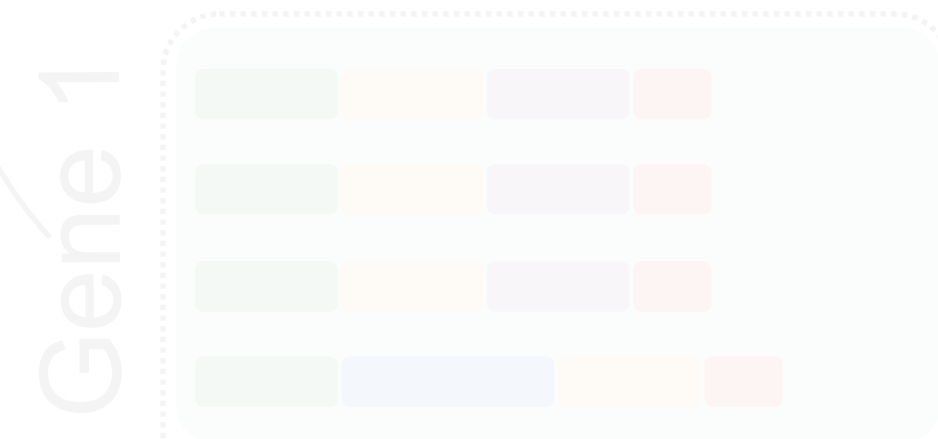
* most protocols actually sequence complementary DNA (cDNA), not RNA directly

Actual protocols are much more involved



Transcript Quantification: An Overview

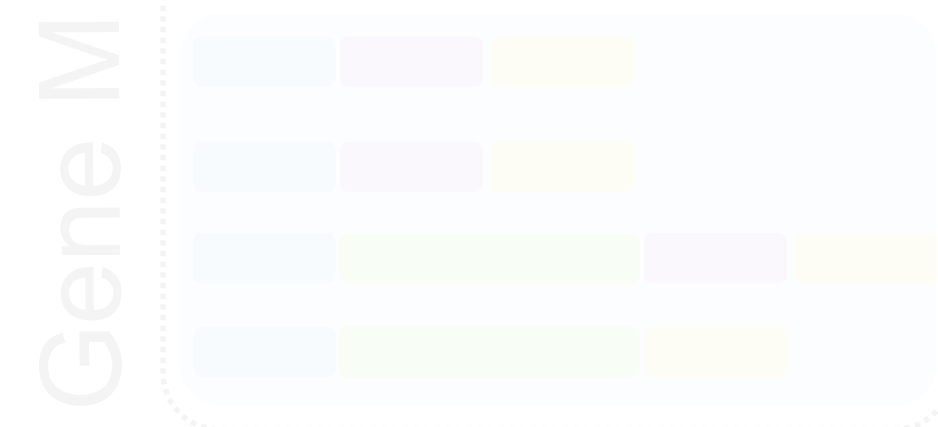




10s-100s of millions of
short (35-300 character) "reads"

Given: (1) Collection of RNA-Seq fragments
(2) A set of known (or assembled) transcript sequences

Estimate: The relative abundance of each transcript



Sample

isoform A

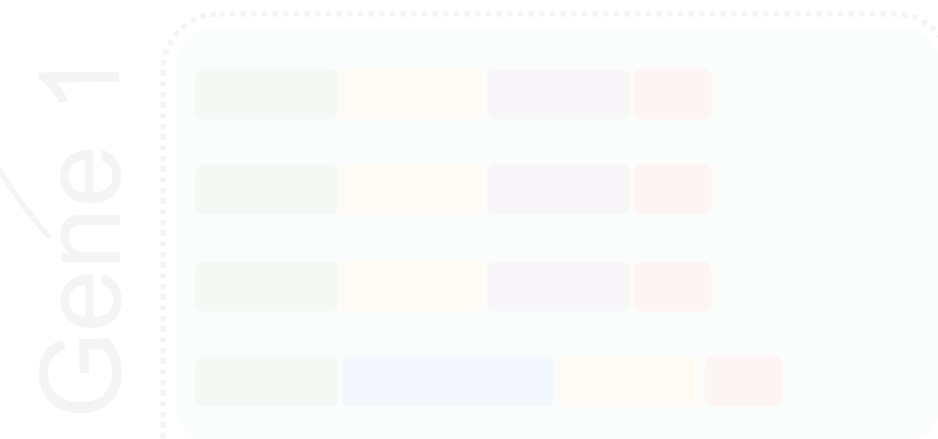
isoform B

isoform C

% Gene 1

% Gene M

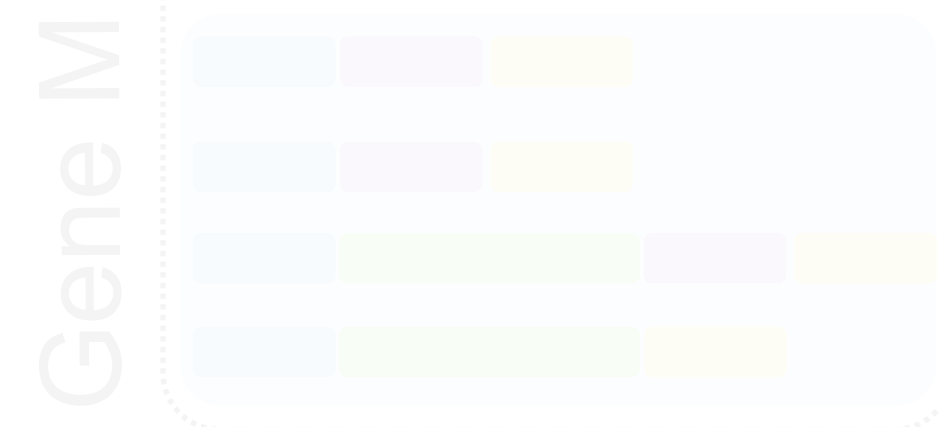
Abundance Estimates



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Sample

isoform A

isoform B

isoform C

% Gene 1

% Gene M

Abundance Estimates

Why not simply “count” reads

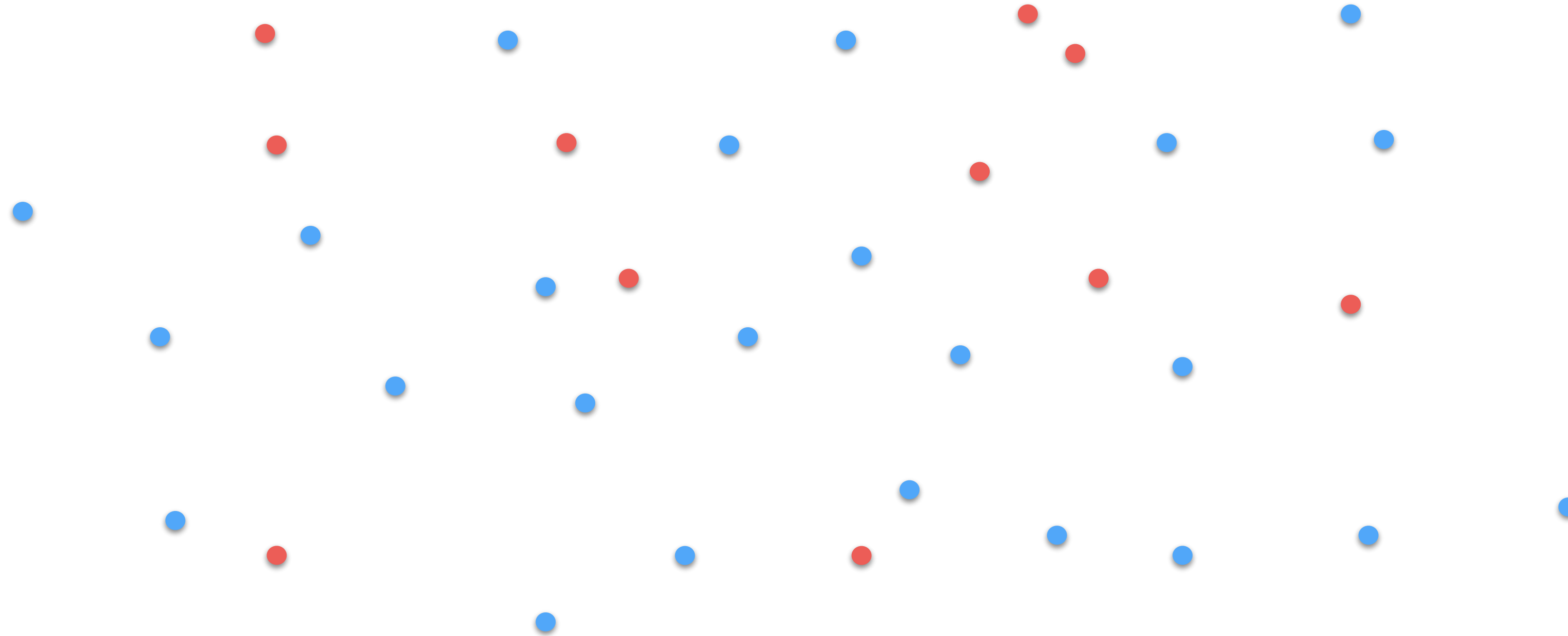
The RNA-seq reads are drawn from transcripts, and our (spliced) aligners let us map them back to the transcripts on the genome from which they originate.

Problem: How do you handle reads that align equally-well to multiple isoforms / or multiple genes?

- Discarding multi-mapping reads leads to incorrect and biased quantification
- Even at the gene-level, the transcriptional output of a gene should depend on what isoforms it is expressing.

First, consider this non-Biological example

Imagine I have two colors of circle, **red** and **blue**. I want to estimate the **fraction of circles** that are **red** and **blue**. I'll *sample* from them by tossing down darts.



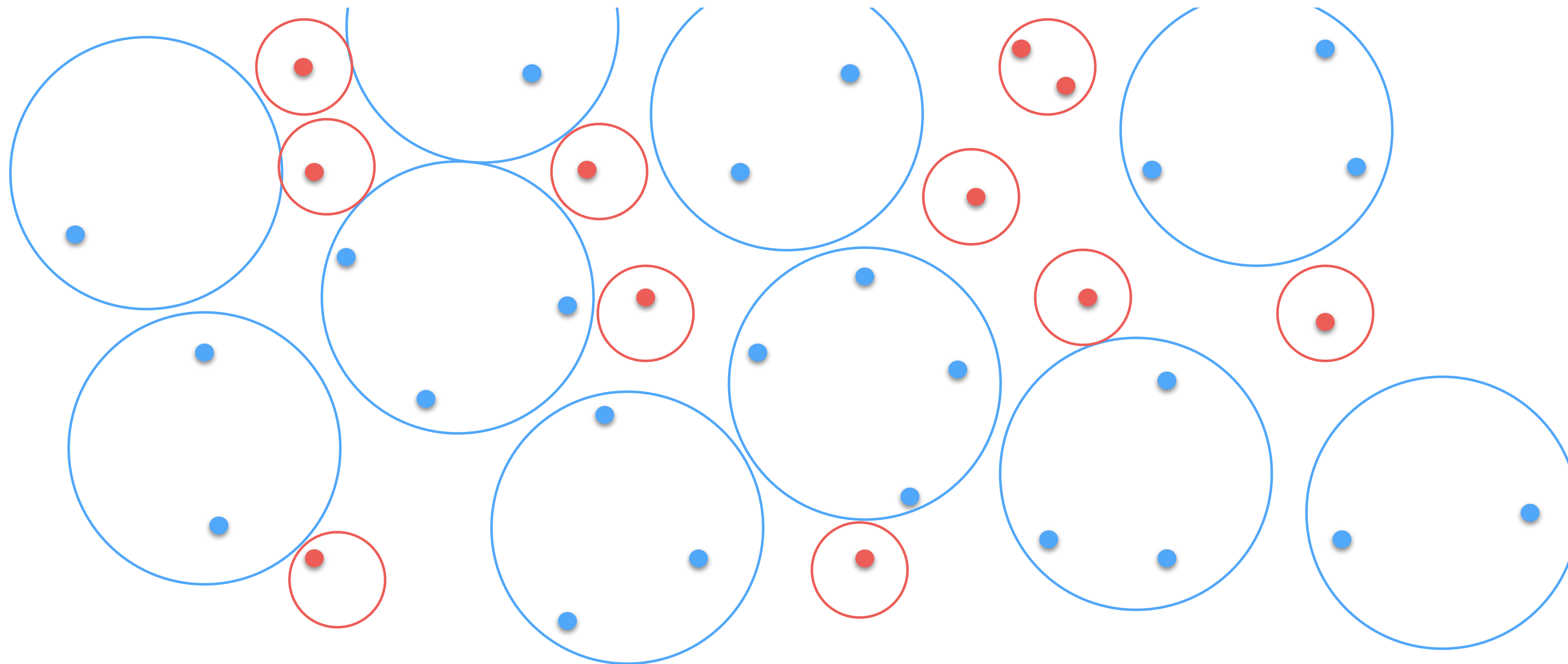
Here, a dot of a color means I hit a circle of that color.

What type of circle is more prevalent?

What is the fraction of red / blue circles?

First, consider this non-Biological example

Imagine I have two colors of circle, **red** and **blue**. I want to estimate the **fraction of circles** that are **red** and **blue**. I'll *sample* from them by tossing down darts.



You're missing a **crucial piece of information!**

The areas!

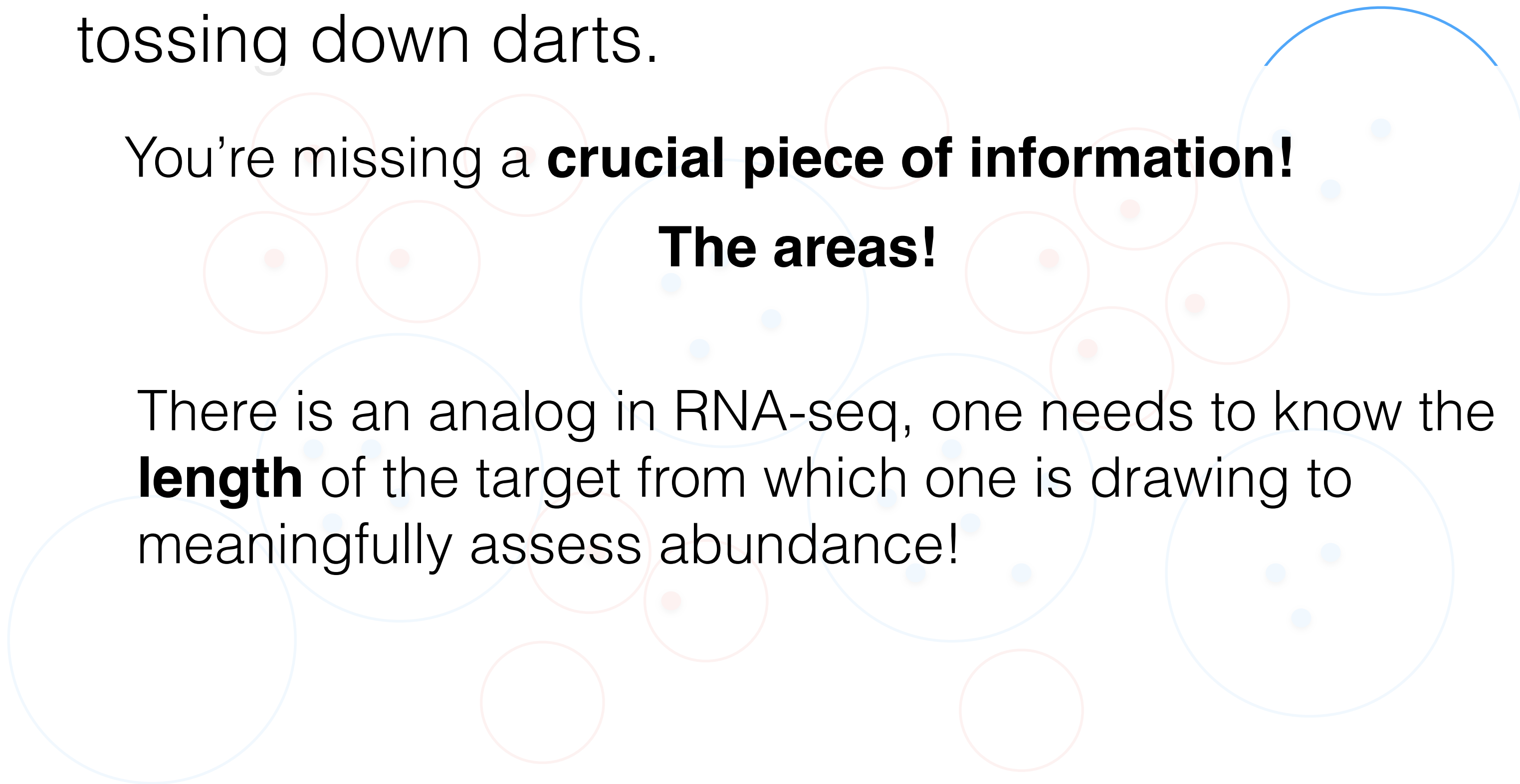
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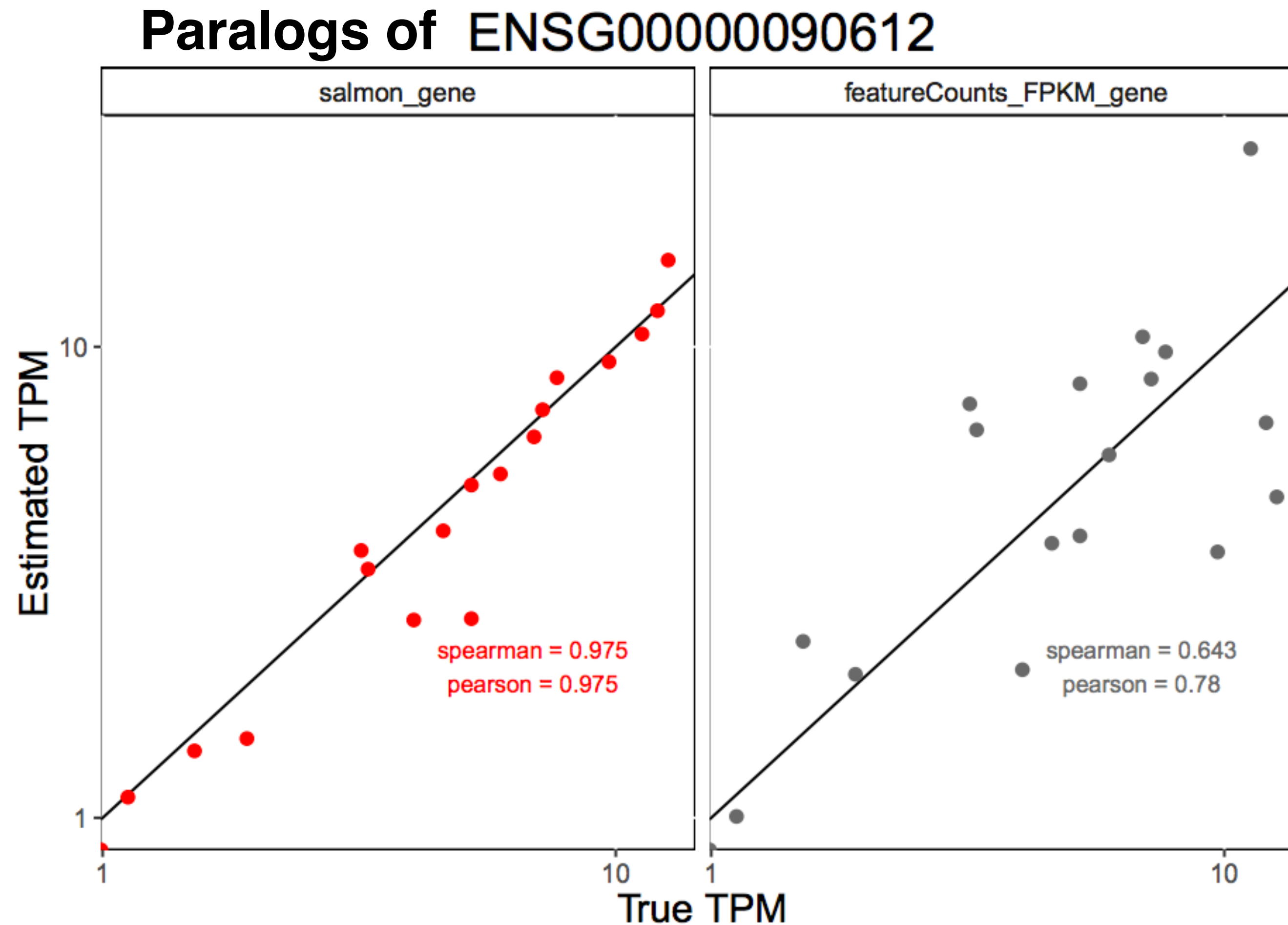
The areas!

There is an analog in RNA-seq, one needs to know the **length** of the target from which one is drawing to meaningfully assess abundance!



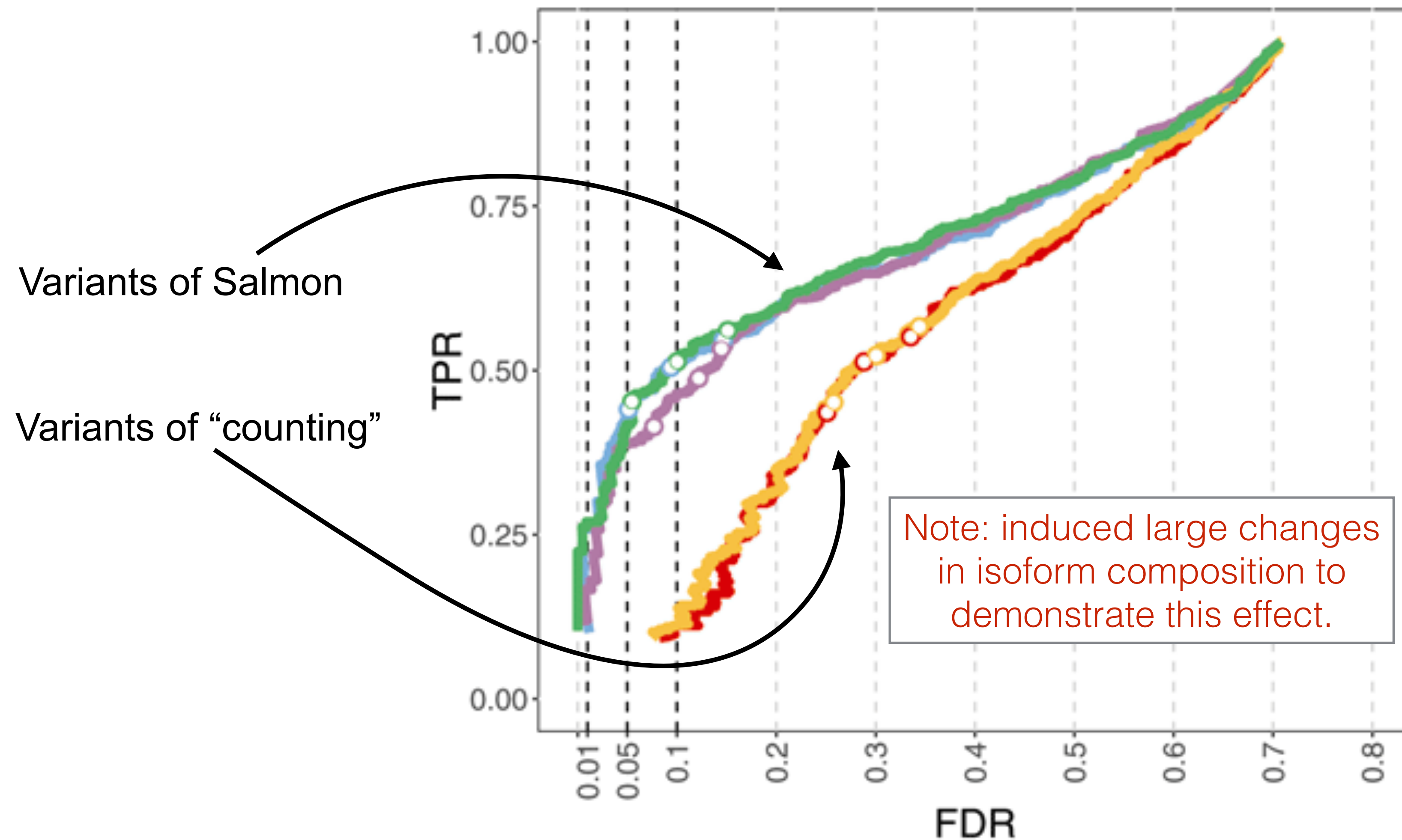
Resolving multi-mapping is fundamental to quantification

Can even affect abundance estimation in **absence** of alternative-splicing
(e.g. paralogous genes)



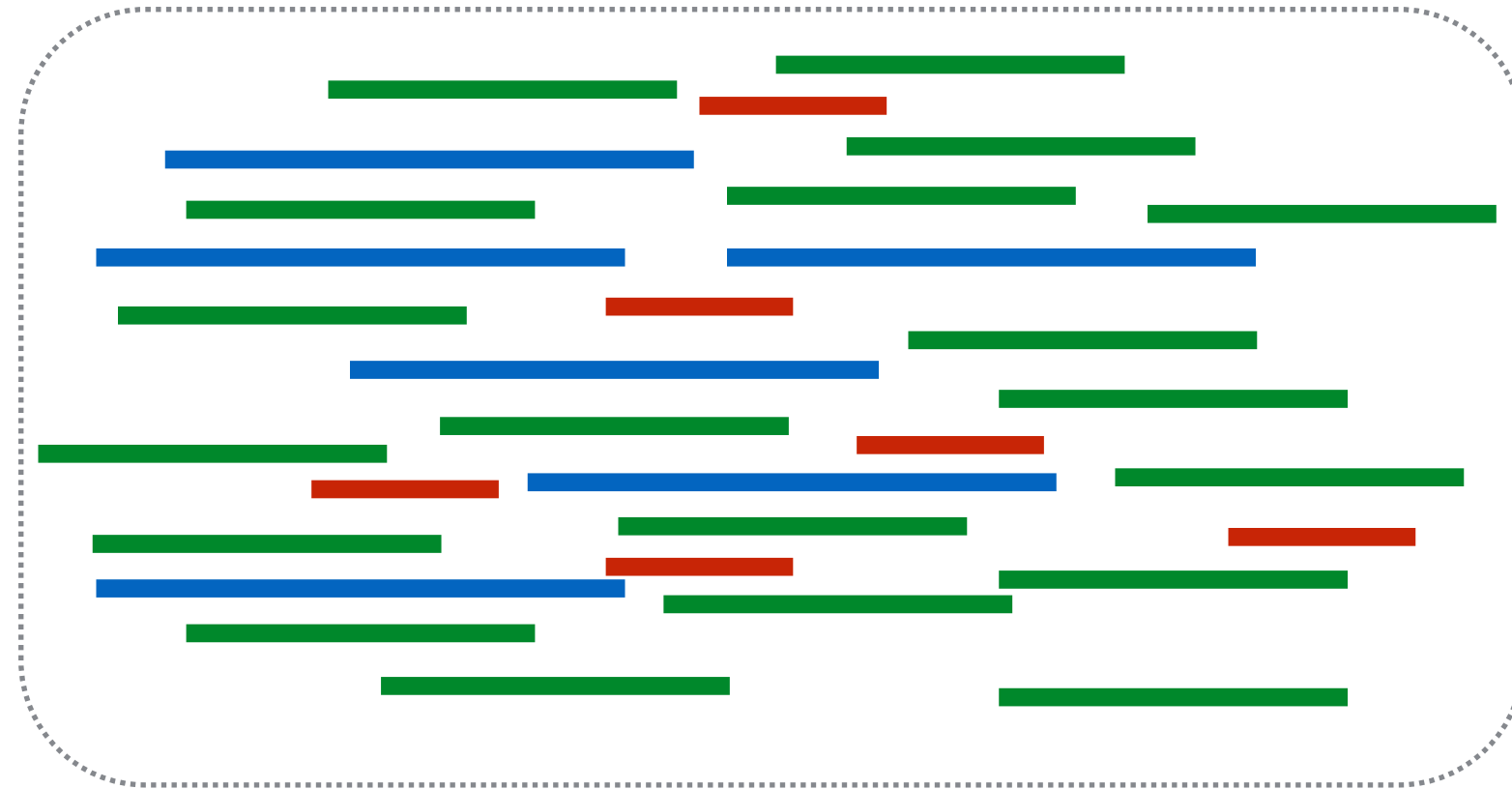
Resolving multi-mapping is fundamental to quantification

These errors can affect DGE calls



How can we perform inference from sequenced fragments?

Experimental Mixture

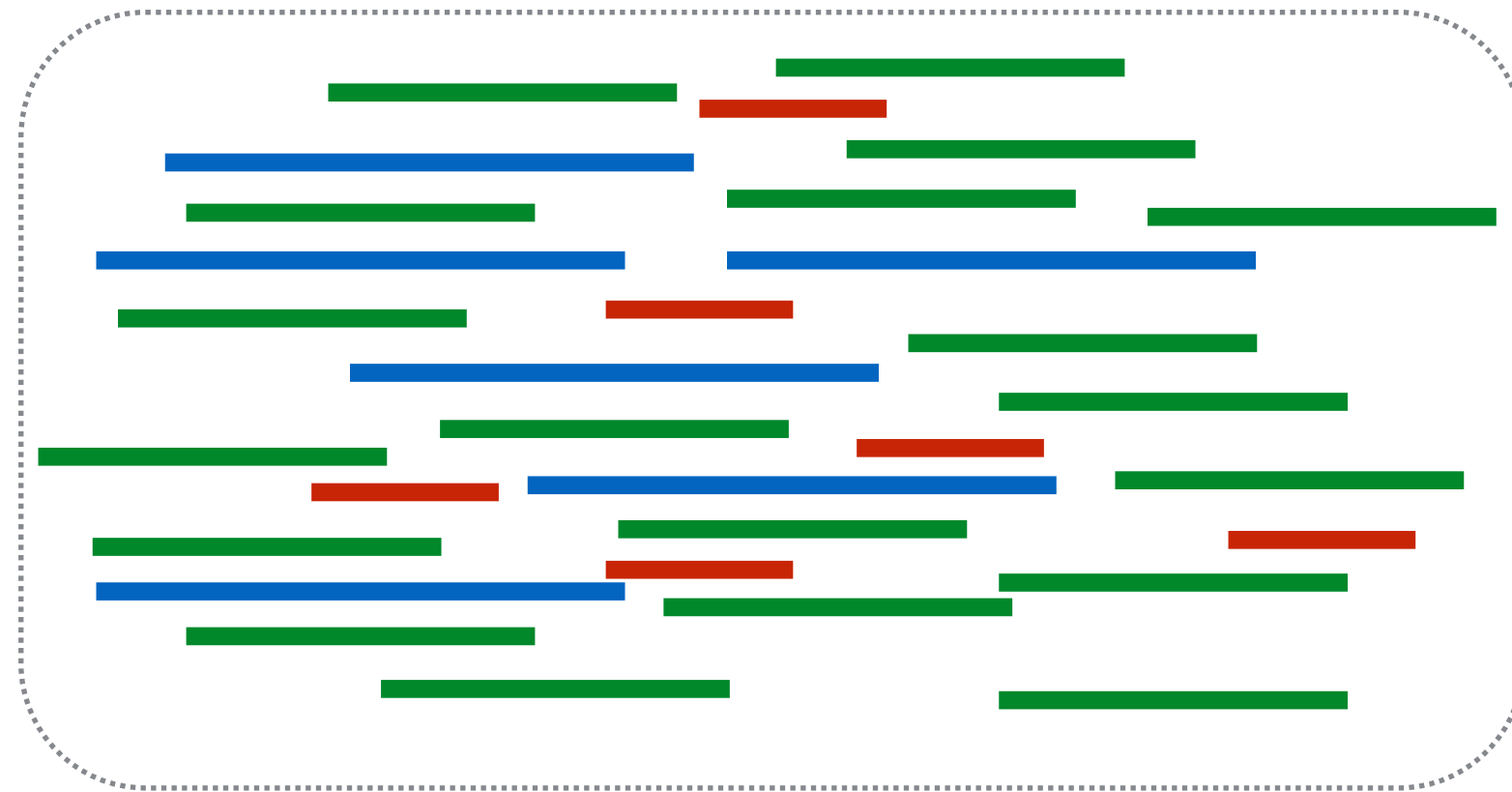


In an unbiased experiment, sampling fragments depends on:

- # of copies of each txp type
- length of each txp type

How can we perform inference from sequenced fragments?

Experimental Mixture



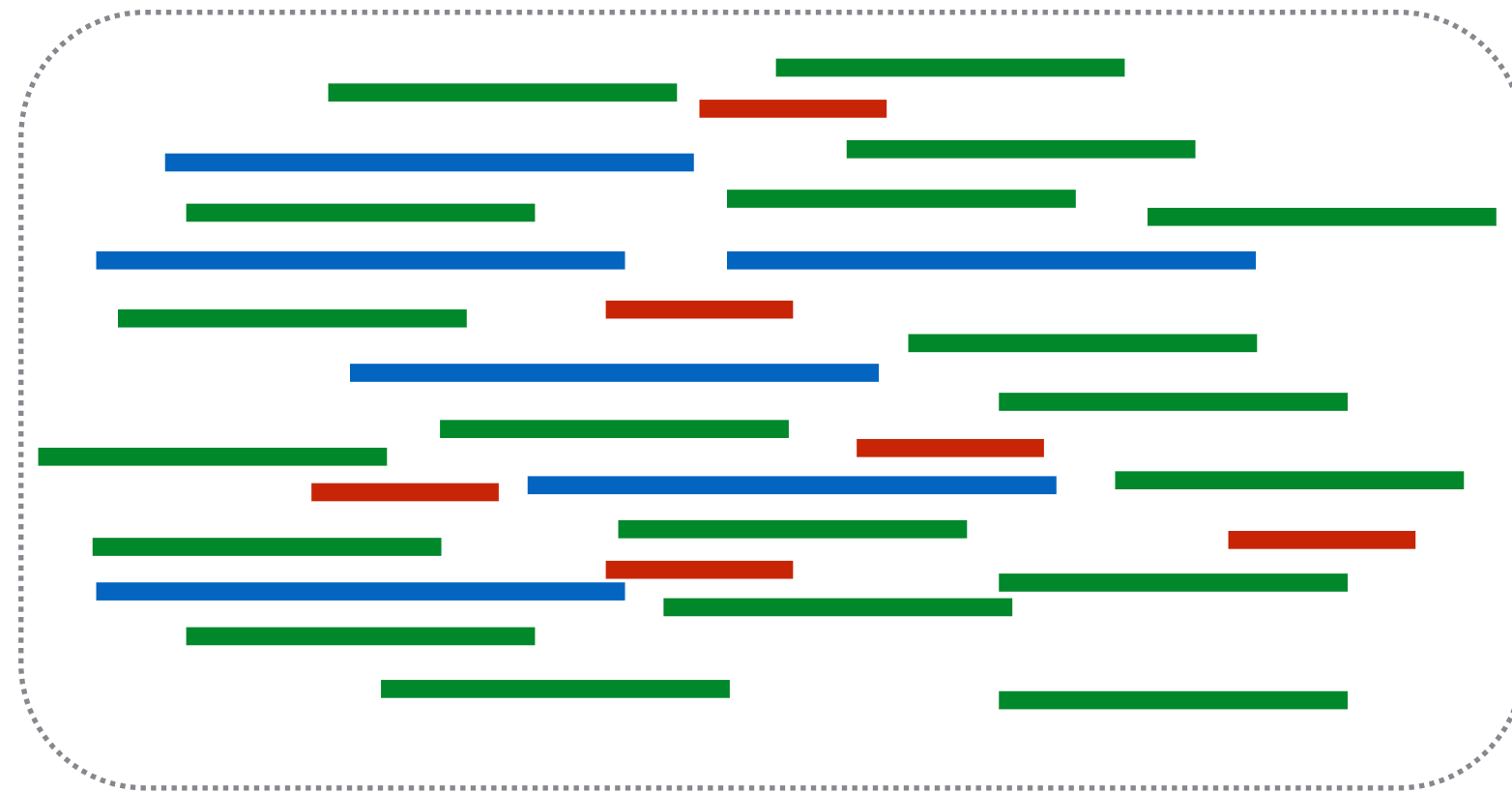
length() = 100

In an unbiased experiment, sampling fragments depends on:

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How can we perform inference from sequenced fragments?

Experimental Mixture



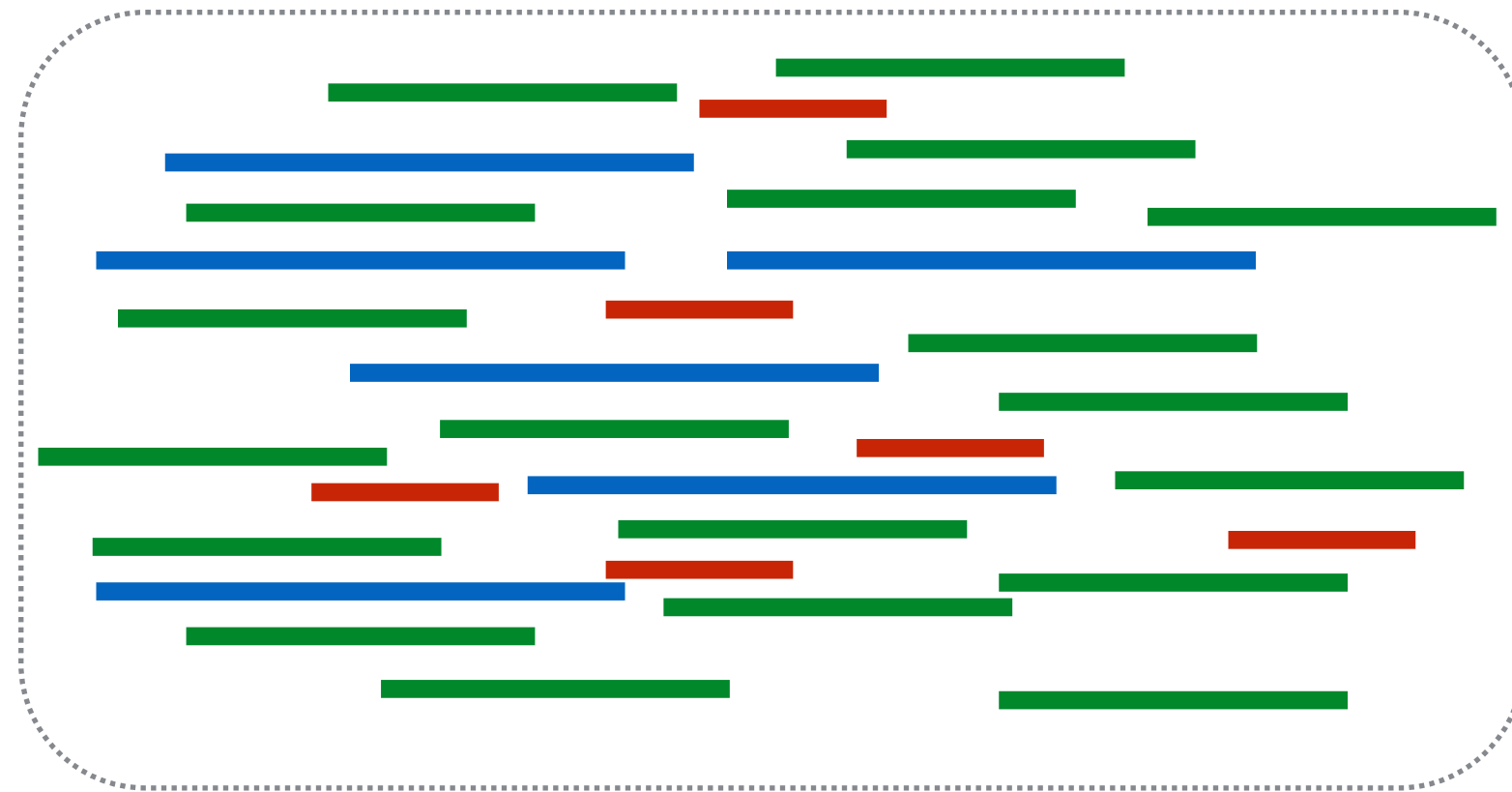
In an unbiased experiment, sampling fragments depends on:

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- length of each txp type

length() = 100 x 6 copies

How can we perform inference from sequenced fragments?

Experimental Mixture



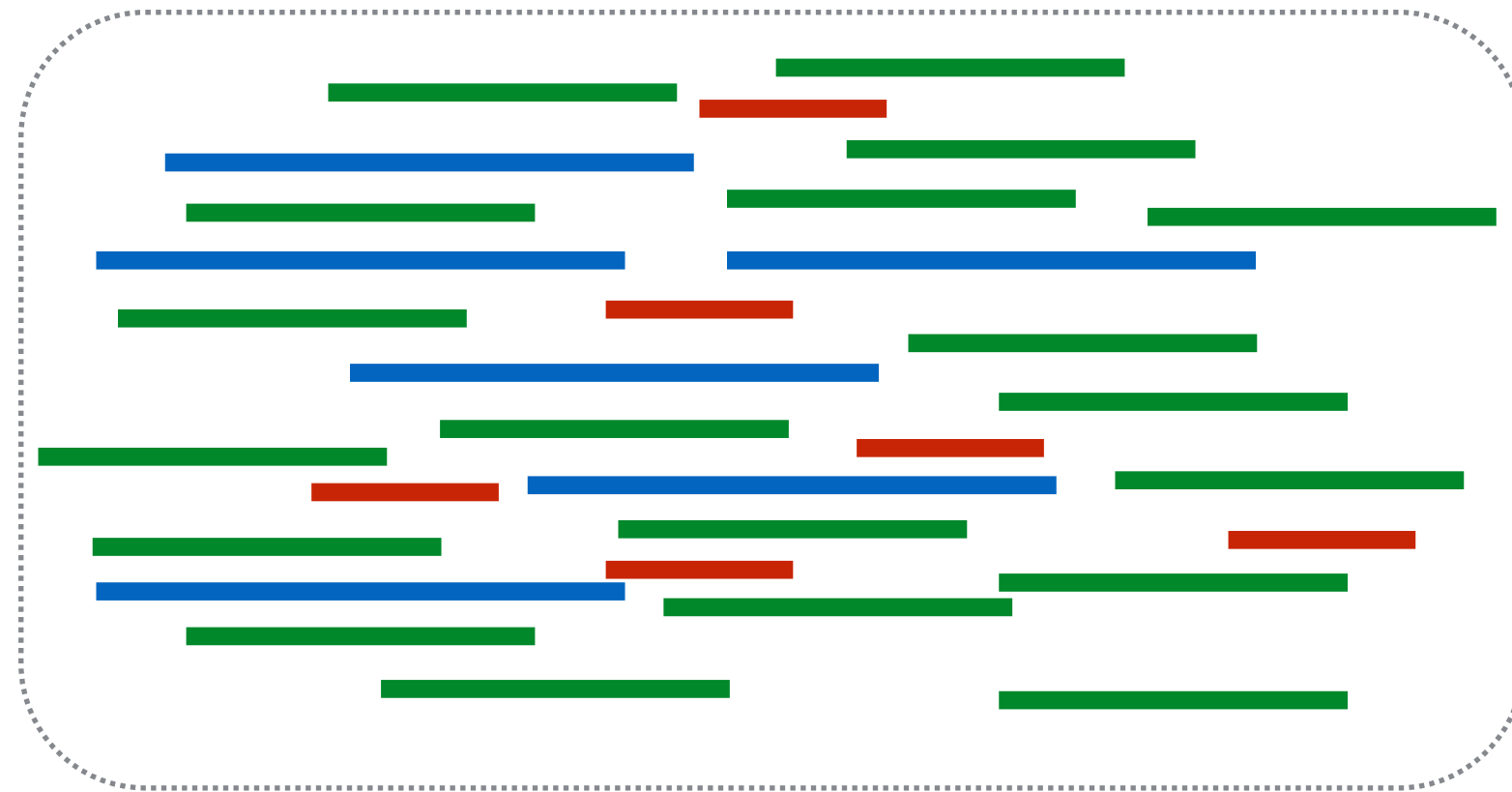
In an unbiased experiment, sampling fragments depends on:

- # of copies of each txp type
- length of each txp type

$$\text{length}(\text{—————}) = 100 \times 6 \text{ copies} = 600 \text{ nt}$$

How can we perform inference from sequenced fragments?

Experimental Mixture



In an unbiased experiment, sampling fragments depends on:

- # of copies of each txp type
- length of each txp type

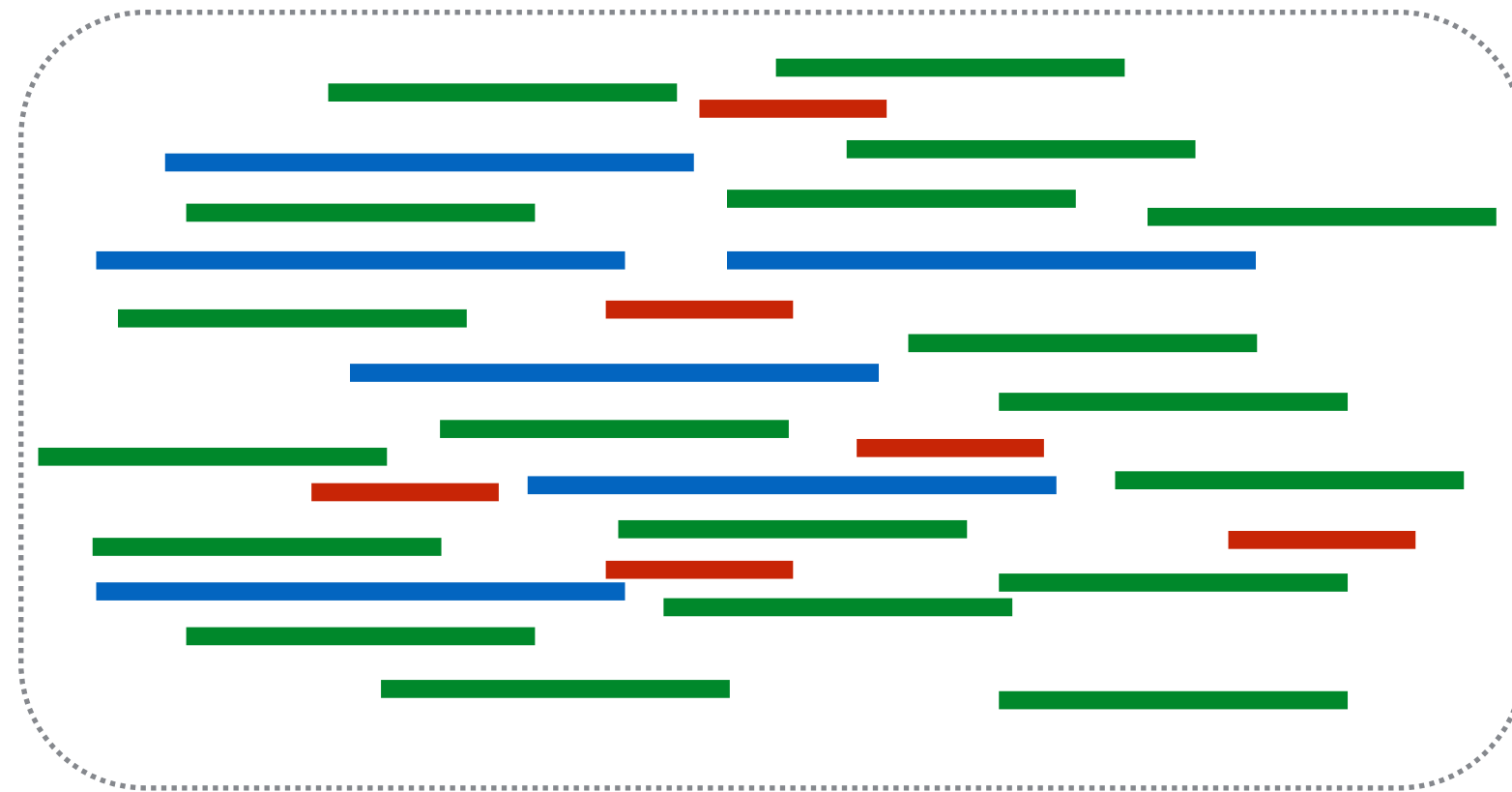
length() = 100 x 6 copies = 600 nt

length() = 66 x 19 copies = 1254 nt

length() = 33 x 6 copies = 198 nt

How can we perform inference from sequenced fragments?

Experimental Mixture



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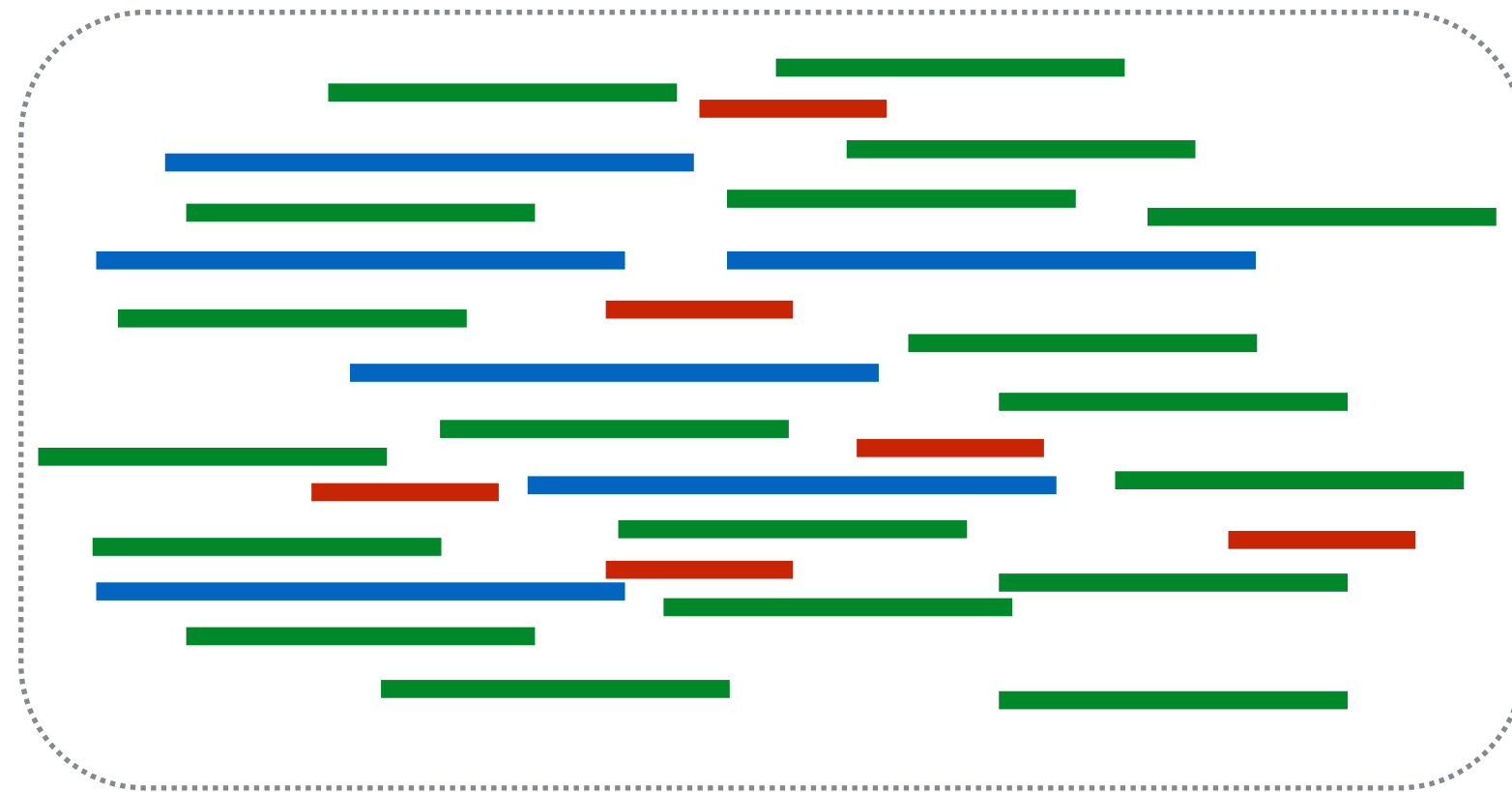
length() = 100 x 6 copies = 600 nt ~ 30% blue

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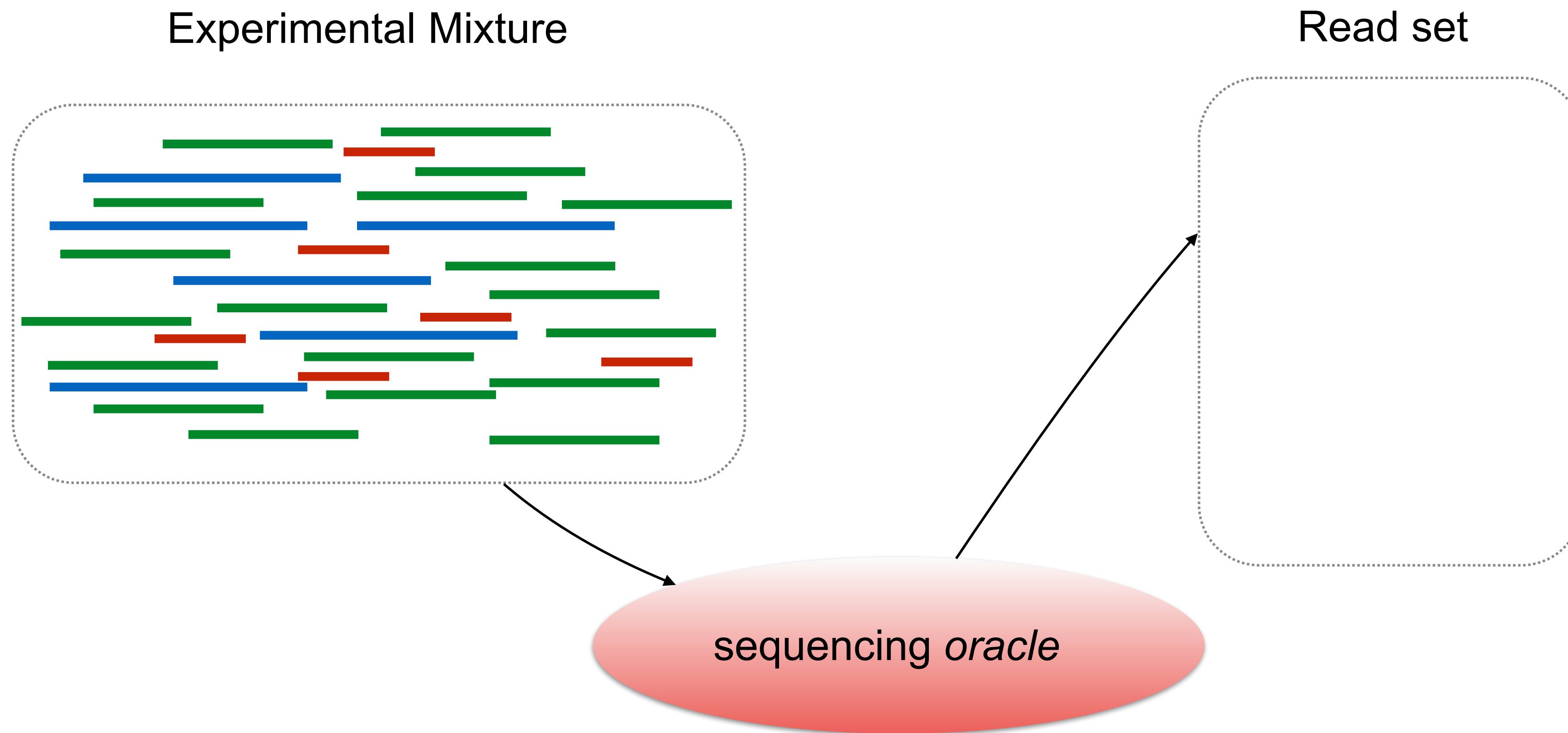
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We call these values $\eta = [0.3, 0.6, 0.1]$ the nucleotide fractions, they become the primary quantity of interest

How can we perform inference from sequenced fragments?

Think about the “ideal” RNA-seq experiment . . .

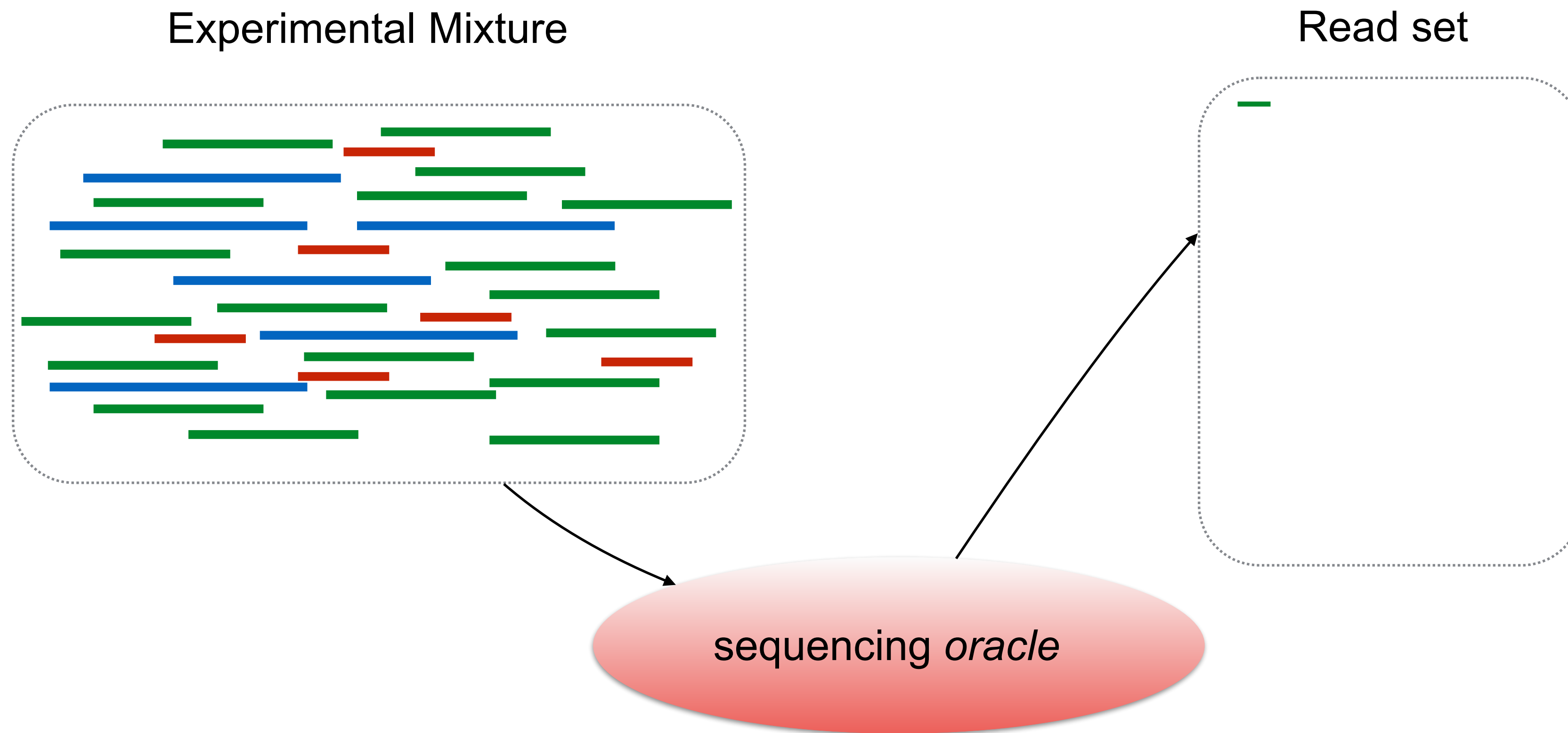


(1) Pick transcript $\mathbf{t} \propto$ total available nucleotides = count * length

(2) Pick a position $\mathbf{p}$ on $\mathbf{t}$ “uniformly at random”

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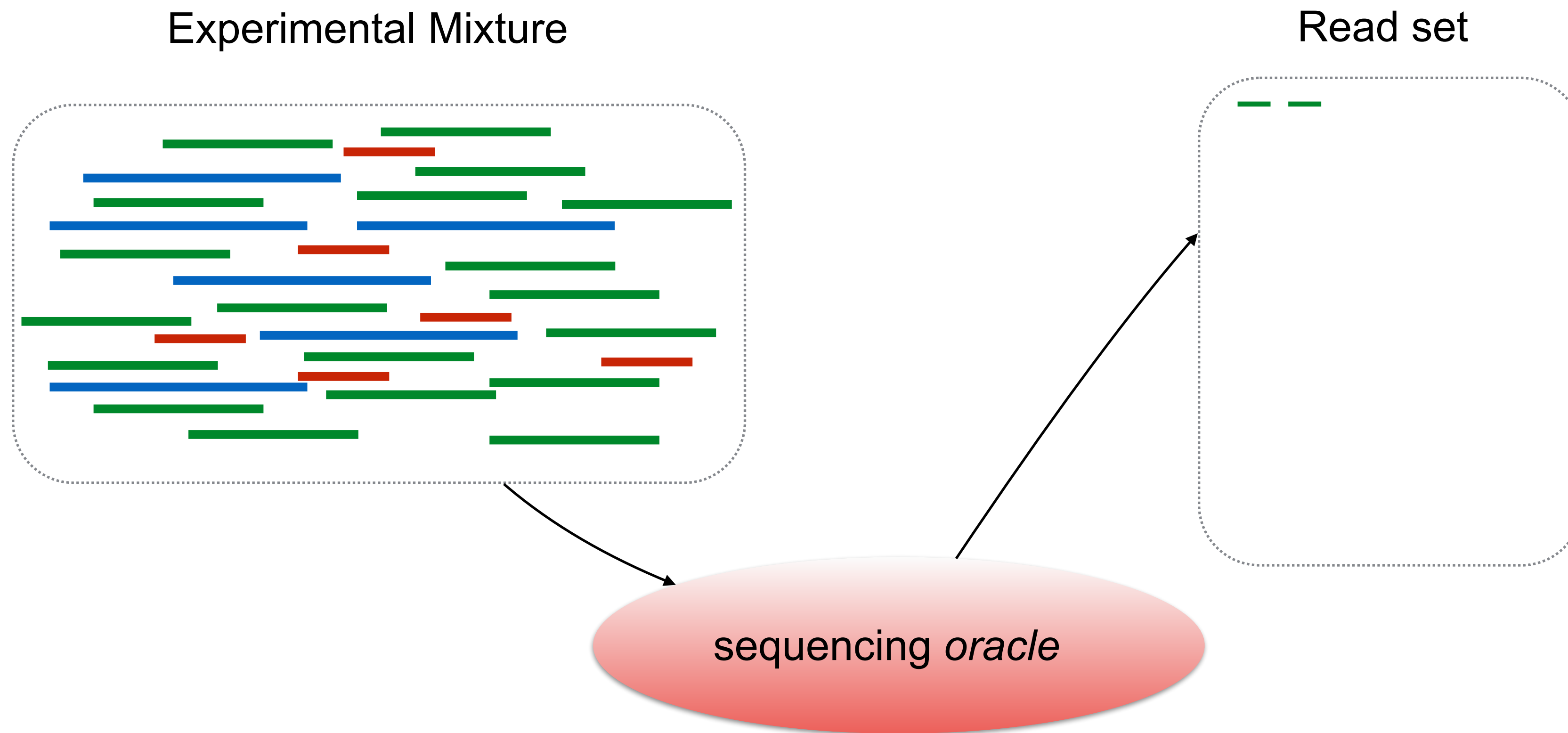


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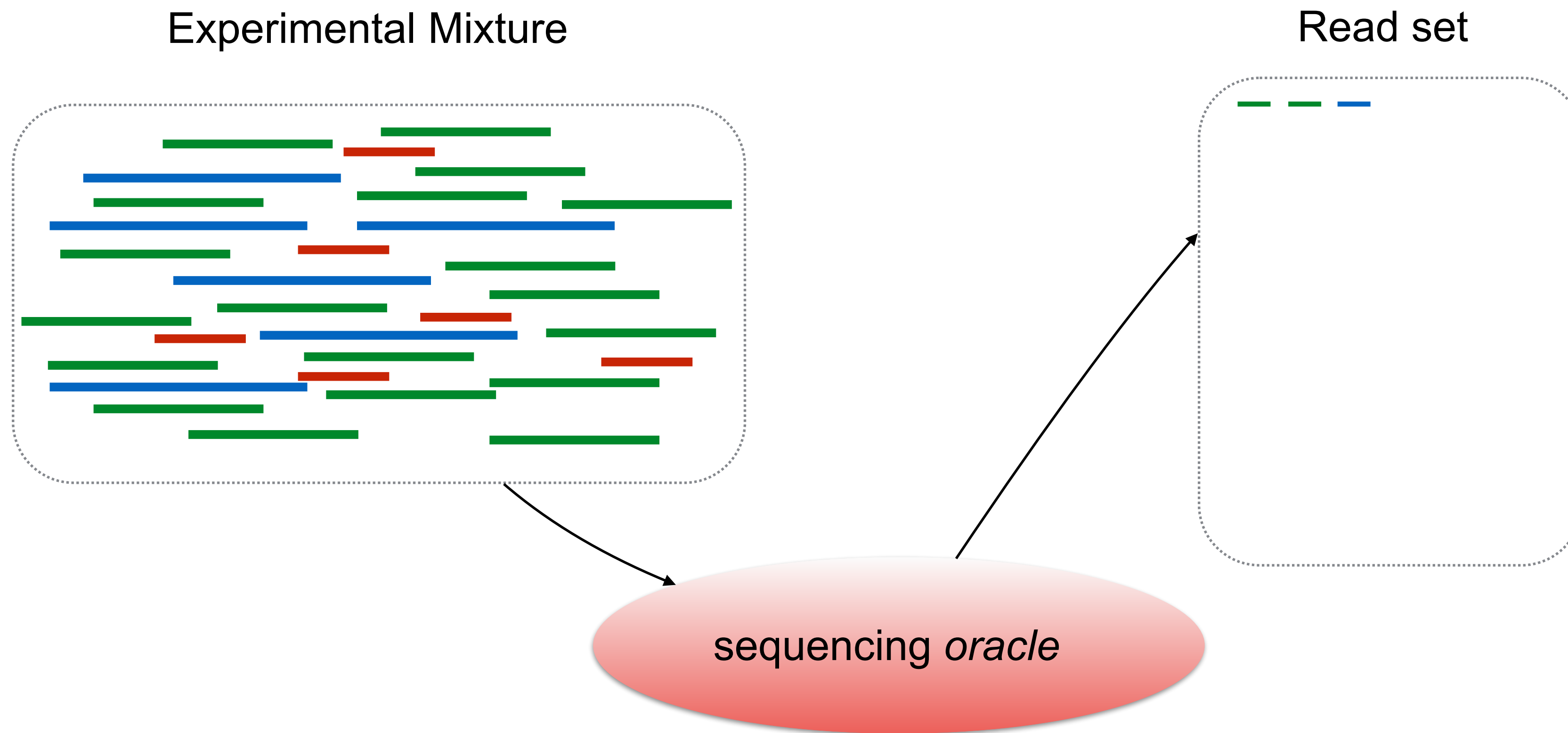


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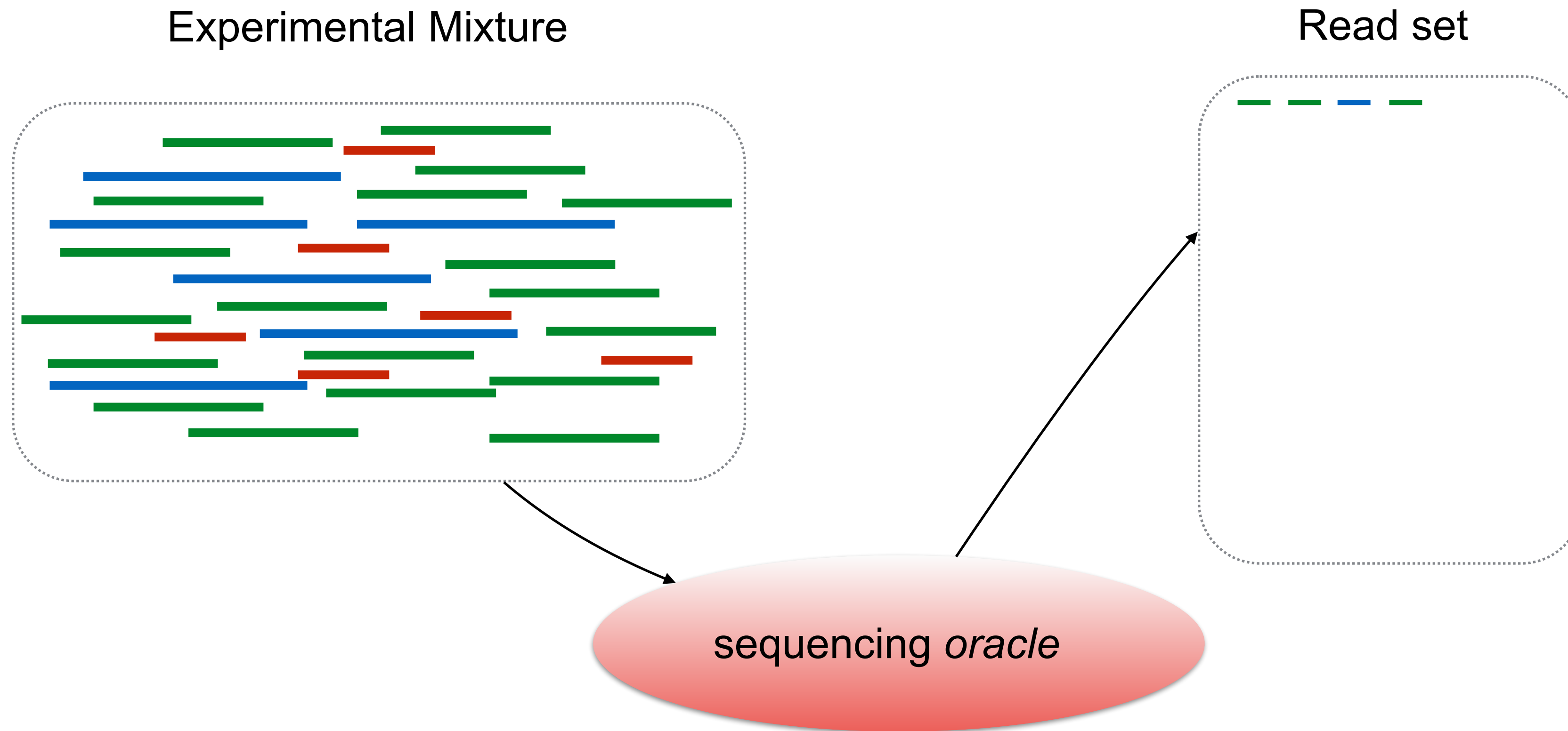


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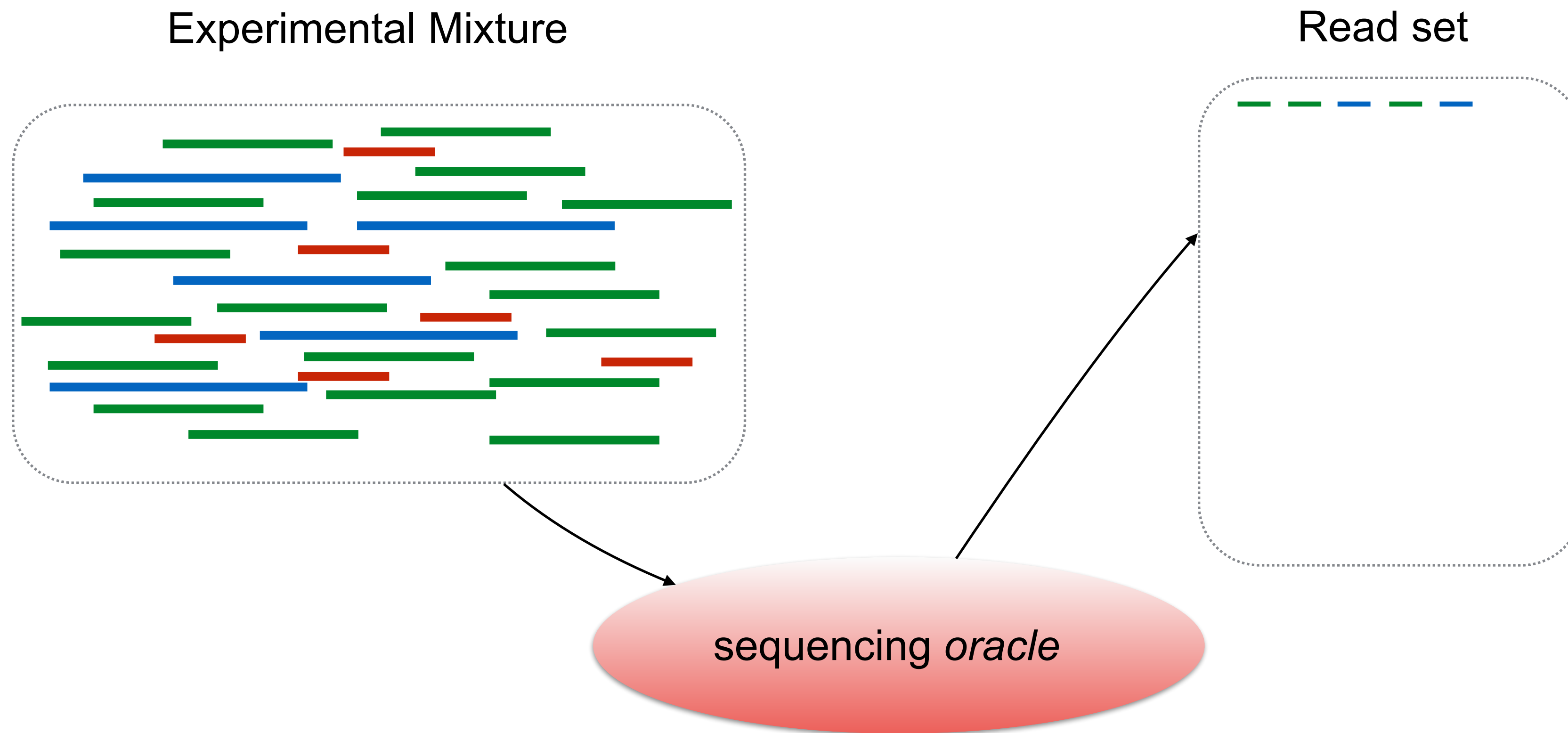


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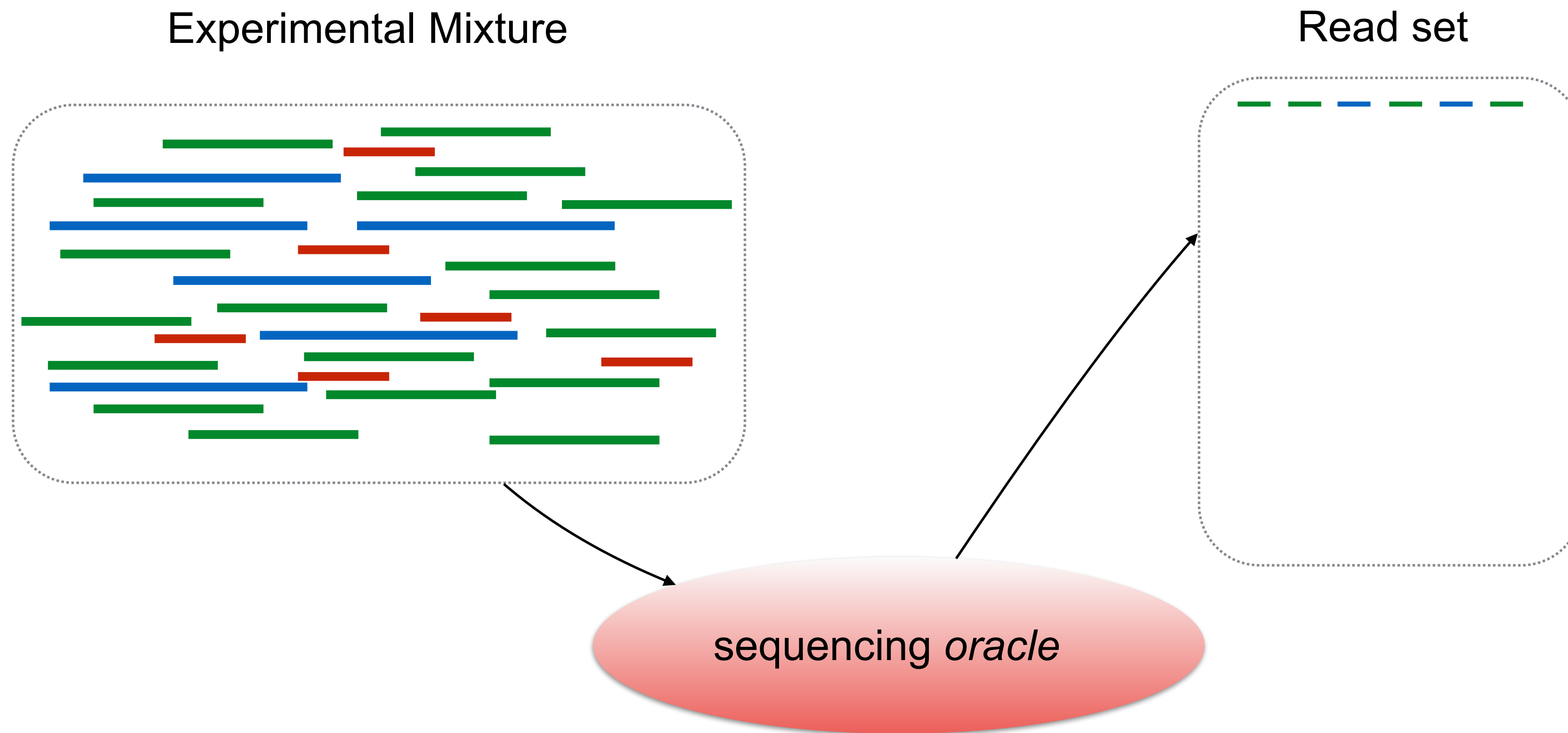


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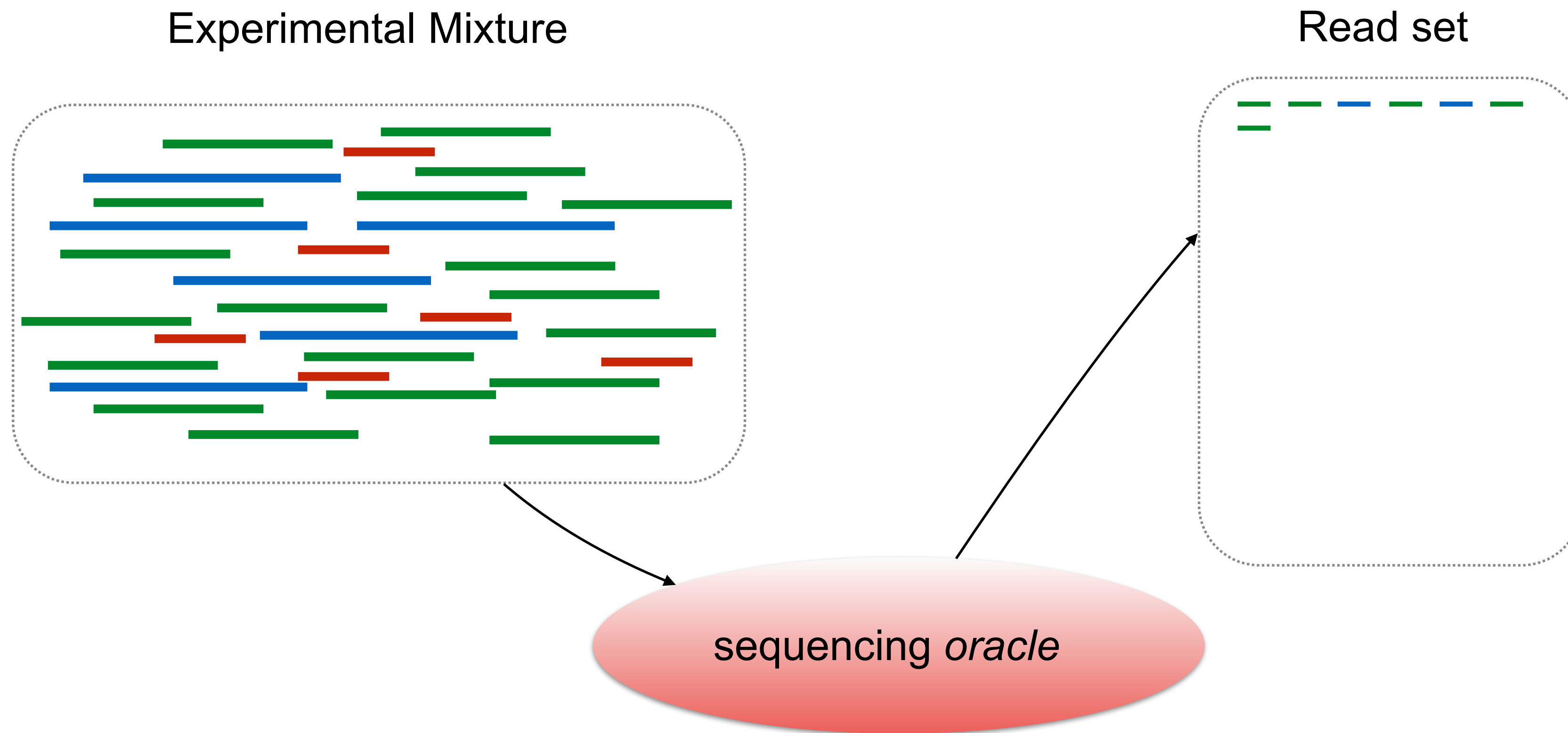


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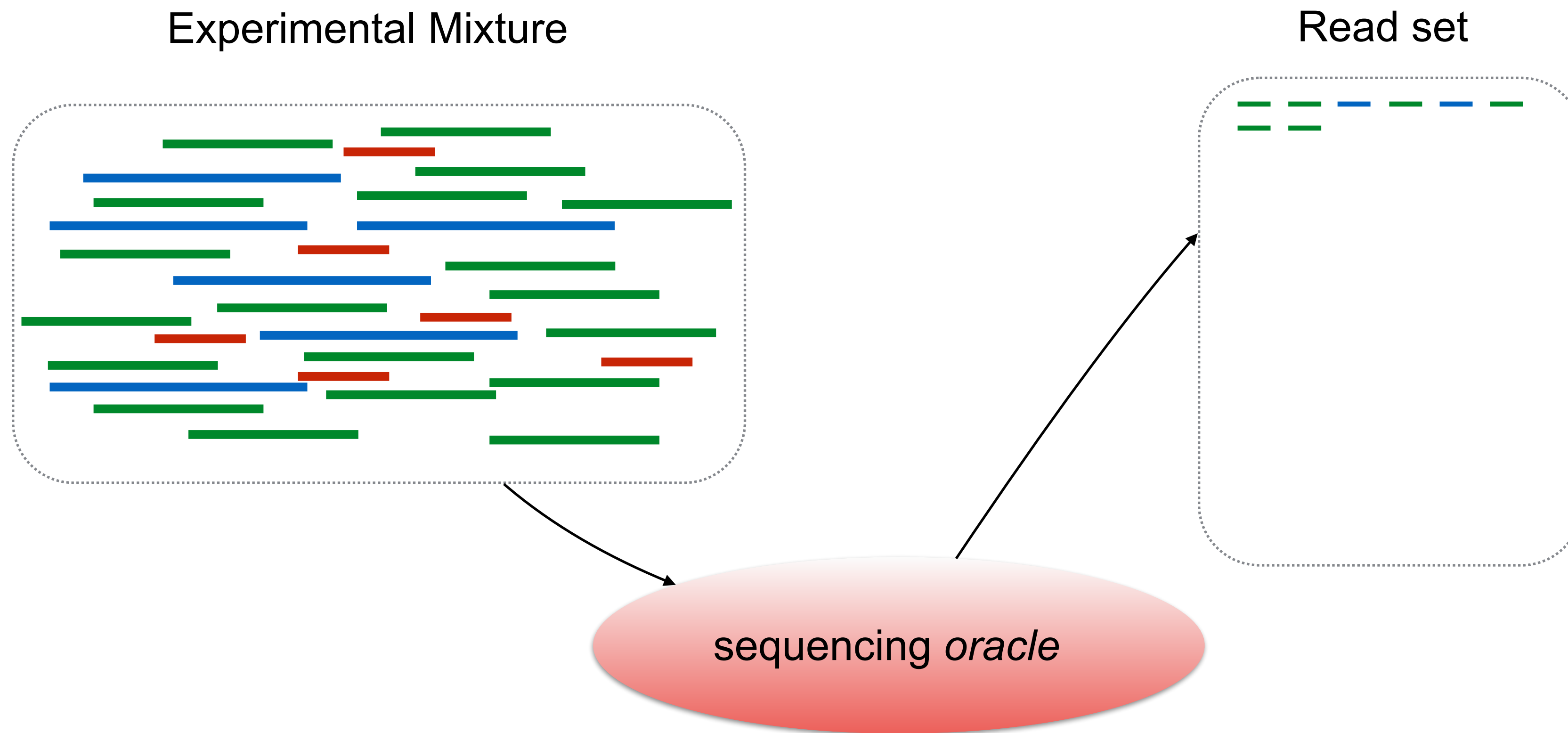


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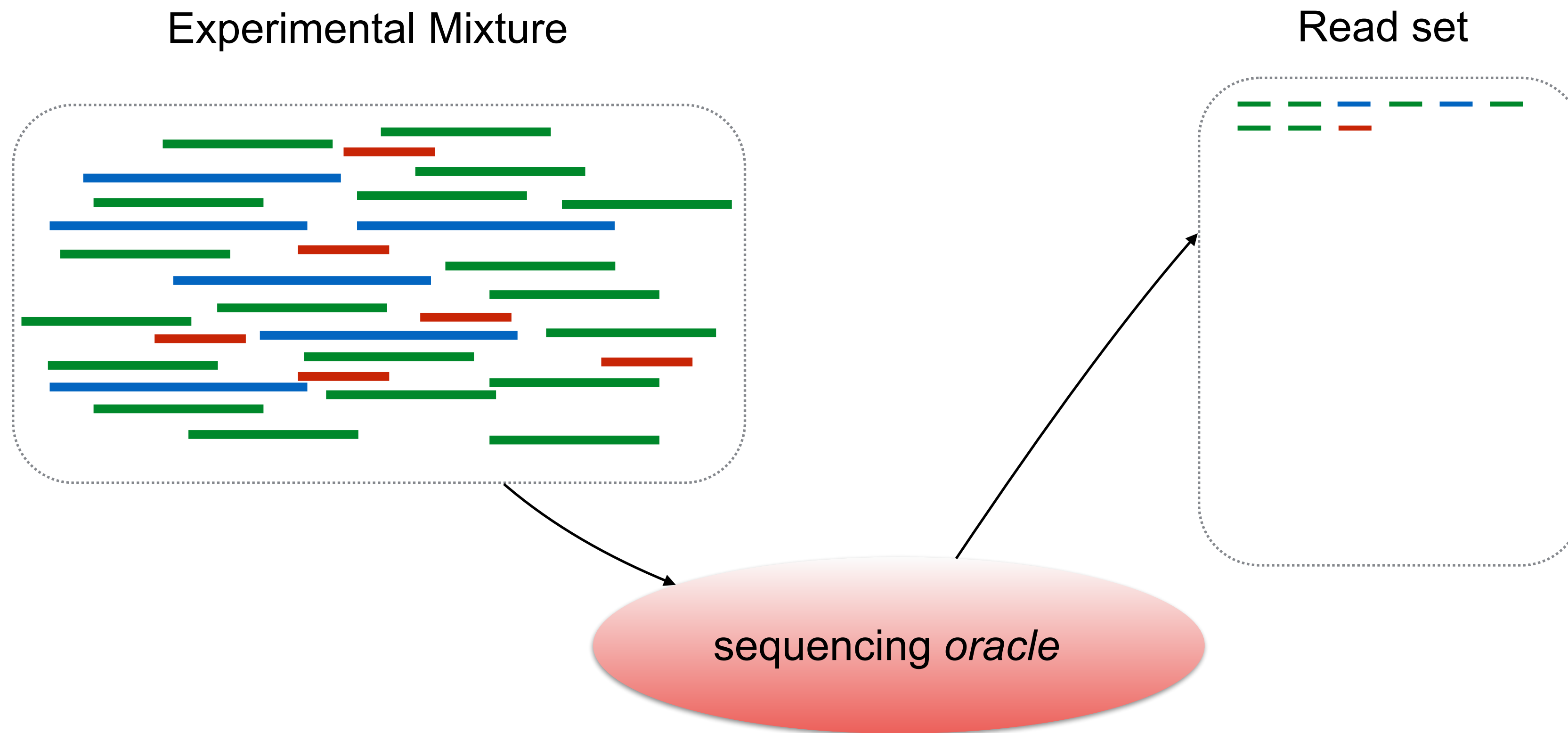


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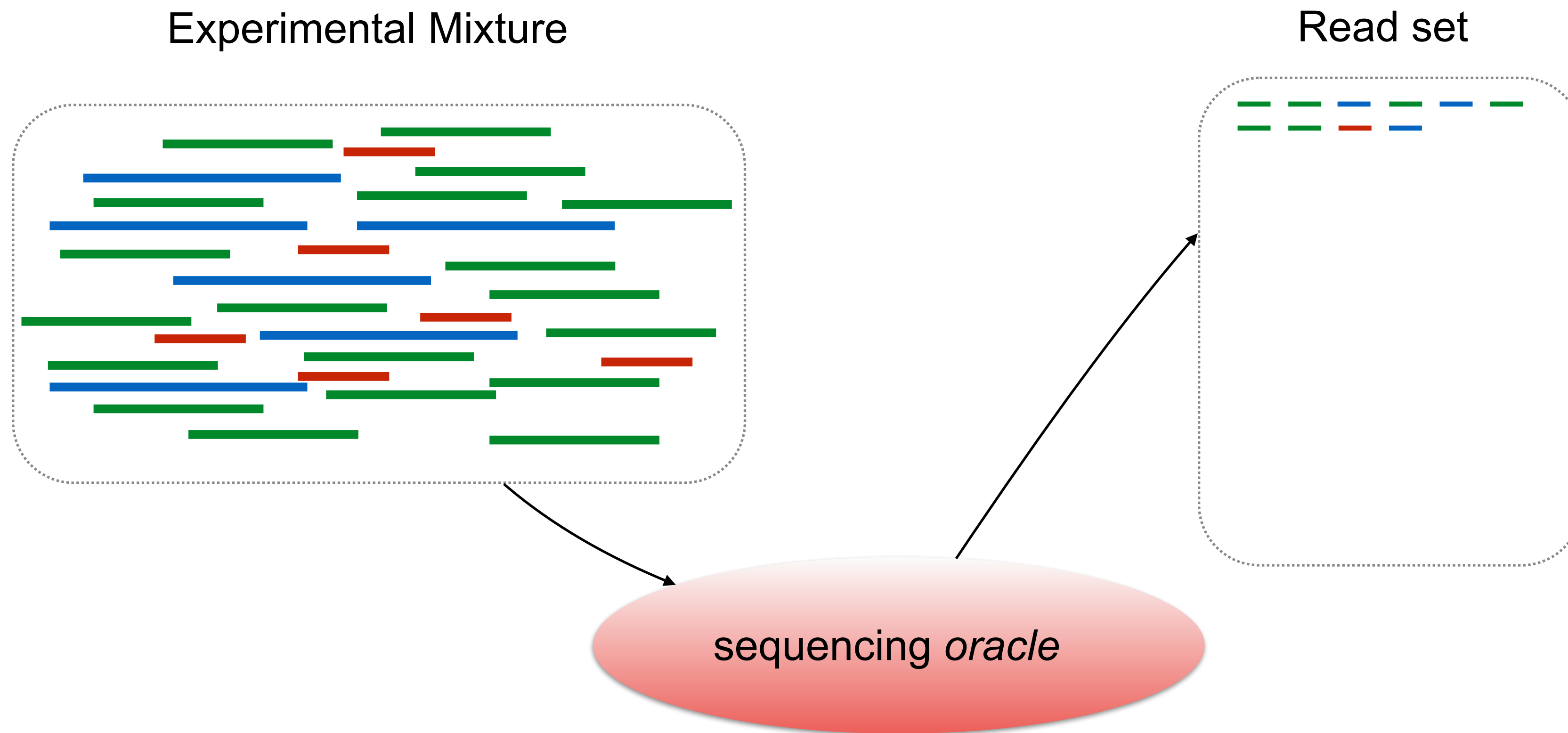


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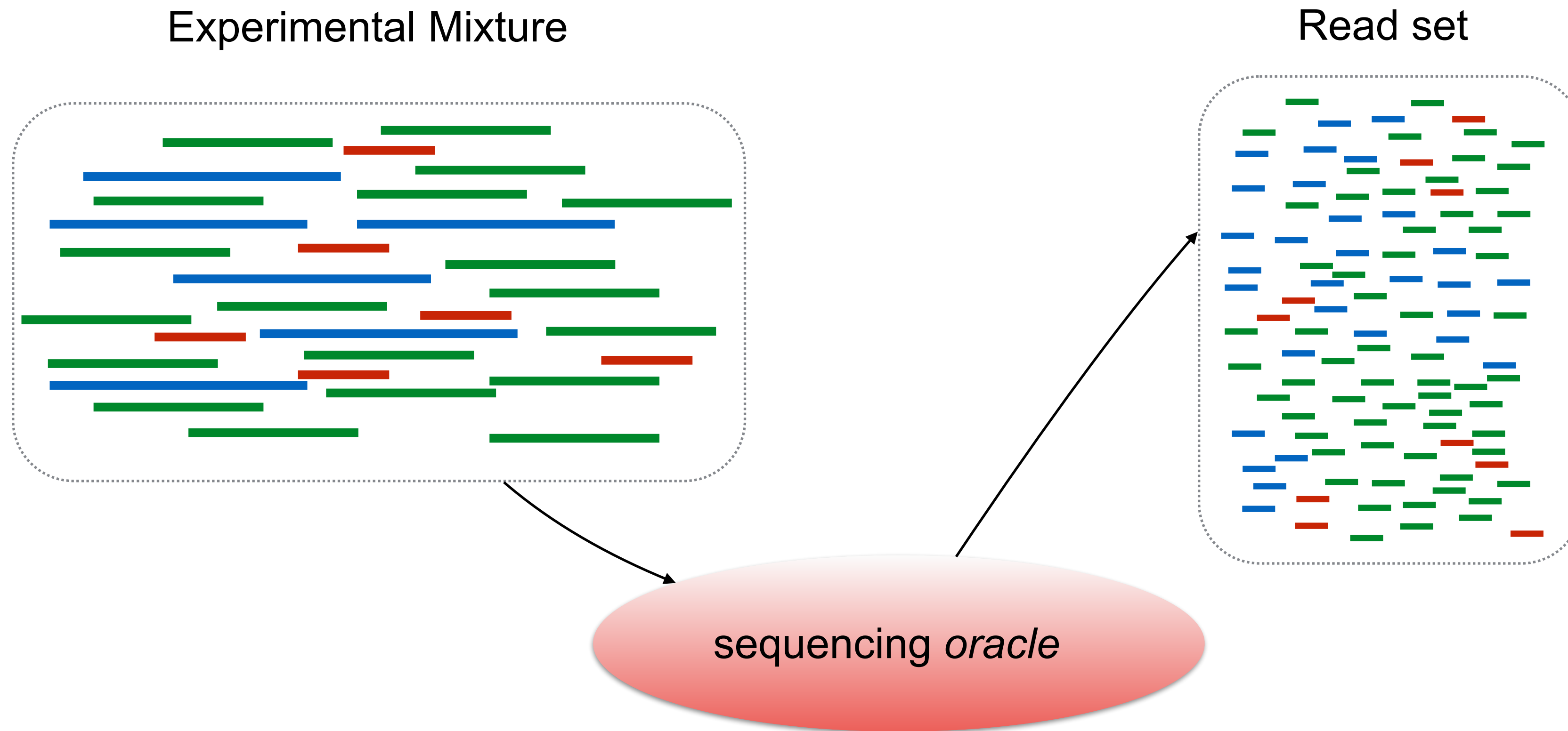


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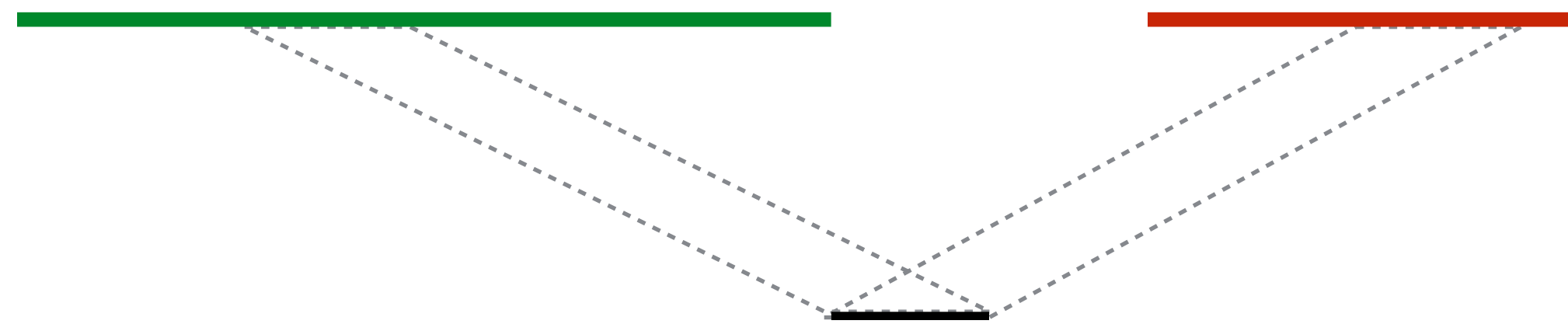
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(1) Pick transcript $\mathbf{t} \propto$ total available nucleotides = count * length

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Resolving a single multi-mapping read



Say we *knew* the η , and observed a *single* read that mapped ambiguously, as shown above.

What is the probability that it truly originated from **G** or **R**?

$$\Pr \{r \text{ from } G\} = \frac{\frac{\eta_G}{\text{length}(G)}}{\frac{\eta_G}{\text{length}(G)} + \frac{\eta_R}{\text{length}(R)}} = \frac{\frac{0.6}{66}}{\frac{0.6}{66} + \frac{0.1}{33}} = 0.75$$

$$\Pr \{r \text{ from } R\} = \frac{\frac{\eta_R}{\text{length}(R)}}{\frac{\eta_G}{\text{length}(G)} + \frac{\eta_R}{\text{length}(R)}} = \frac{\frac{0.1}{33}}{\frac{0.6}{66} + \frac{0.1}{33}} = 0.25$$

normalization factor

length() = 100 x 6 copies = 600 nt ~ 30% **blue**

length() = 66 x 19 copies = 1254 nt ~ 60% **green**

length() = 33 x 6 copies = 198 nt ~ 10% **red**

Units for Relative Abundance

TPM (Transcripts Per Million)

$$\text{TPM}_i = \rho_i \times 10^6 \text{ where } 0 \leq \rho_i \leq 1 \text{ and } \sum_i \rho_i = 1$$

$$\rho_i = \frac{\frac{X_i}{\ell_i}}{\sum_j \frac{X_j}{\ell_j}}$$

← Reads coming from transcript i

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Reads coming from transcript i

Length of transcript i

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abundance of i
as fraction of all
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Reads coming from
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Reads coming from
transcript i

Length of transcript i

Interlude: Maximum Likelihood Estimation & the EM-algorithm

Maximum Likelihood & EM slides taken from UW CSE312 (winter '17)

A probabilistic view of RNA-Seq quantification

$$\Pr\{\mathcal{F} \mid \boldsymbol{\eta}, \mathcal{T}\} = \prod_{j=1}^N \Pr\{f_j \mid \boldsymbol{\eta}, \mathcal{T}\}$$

$$= \prod_{j=1}^N \sum_{i=1}^M \Pr\{t_i \mid \boldsymbol{\eta}\} \cdot \Pr\{f_j \mid t_i, \mathbf{z}_{ji} = 1\}$$

nucleotide fractions

known transcriptome

assumes independence of fragments

observed fragments (reads)

Prob. of selecting t_i given $\boldsymbol{\eta}$

Prob. of generating fragment f_j given that it originates from t_i

Depends on abundance estimate

Independent of abundance estimate

We want to find the values of $\boldsymbol{\eta}$ that **maximize** this probability.
We can do this (at least locally) using the EM algorithm.

A probabilistic view of RNA-Seq quantification

nucleotide fractions known transcriptome assumes independence of fragments

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A probabilistic view of RNA-Seq quantification

nucleotide fractions known transcriptome assumes independence of fragments

$$\Pr\{\mathcal{F} \mid \boldsymbol{\eta}, \mathcal{T}\} = \prod_{j=1}^N \Pr\{f_j \mid \boldsymbol{\eta}, \mathcal{T}\}$$

observed fragments (reads)

We can safely truncate $\Pr\{t_i \mid \boldsymbol{\eta}\}$ to 0 for transcripts where a fragment doesn't map/align.

$$= \prod_{j=1}^N \sum_{i=1}^M \Pr\{t_i \mid \boldsymbol{\eta}\} \cdot \Pr\{f_j \mid t_i, z_{ji} = 1\}$$

Prob. of selecting t_i given $\boldsymbol{\eta}$ Prob. of generating fragment f_j given that it originates from t_i

Depends on abundance estimate Independent of abundance estimate

We want to find the values of $\boldsymbol{\eta}$ that **maximize** this probability.
We can do this (at least locally) using the EM algorithm.

A probabilistic view of RNA-Seq quantification

E-step: (what is the “soft assignment” of each read to the transcripts where it aligns)

$$E_{Z|\mathcal{F},\eta^{(t)}}[Z_{nij}] = P(Z_{nij} = 1 \mid \mathcal{F}, \eta^{(t)}) = \frac{(\eta_i^{(t)} / \ell_i) P(f_n \mid Z_{nij} = 1)}{\sum_{i',j'} (\eta_{i'}^{(t)} / \ell_{i'}) P(f_n \mid Z_{ni'j'} = 1)}$$

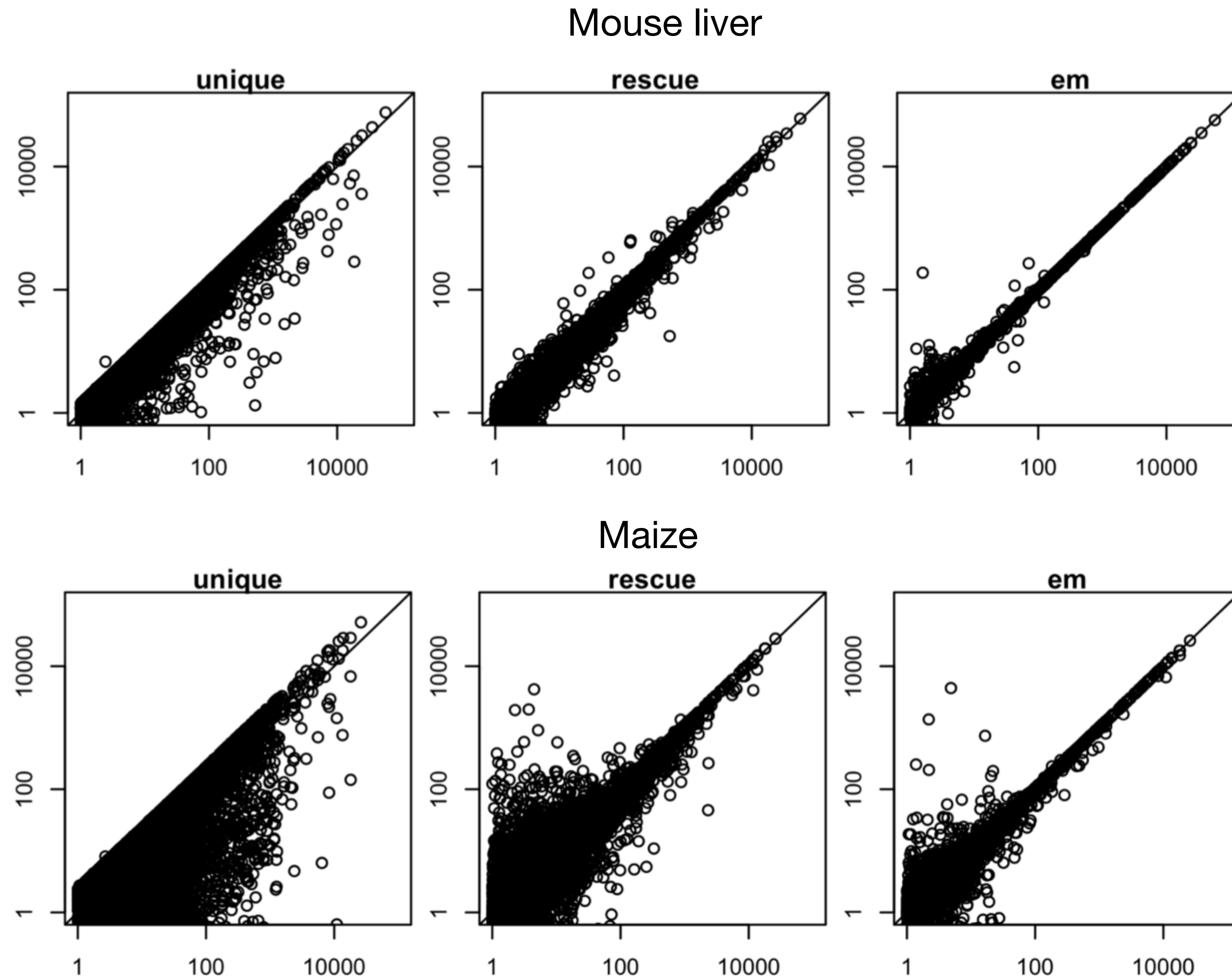
M-step: Given these soft assignments, how abundant is each transcript?

$$\eta_i^{(t+1)} = \frac{E_{Z|\mathcal{F},\eta^{(t)}}[C_i]}{N},$$

$$\text{where } C_i = \sum_{n,j} Z_{nij}$$

This approach is quite effective. Unfortunately, it's also quite slow.

Gene expression estimation accuracy in simulated data

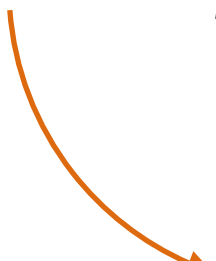


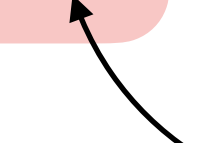
A probabilistic view of RNA-Seq quantification

We want to find the values of η that **maximize** this probability.
We can do this (at least locally) using the EM algorithm.

but

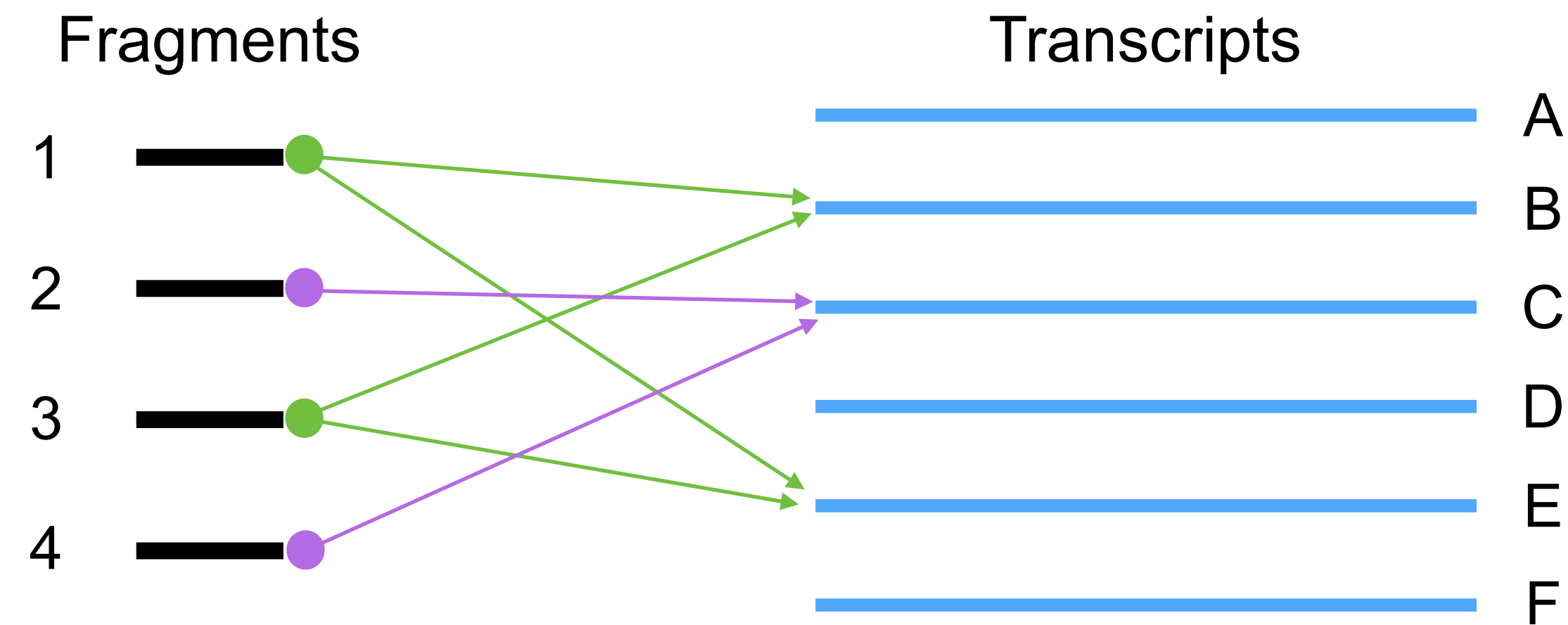
This leads to an iterative EM algorithm where *each iteration* scales in the total number of **alignments** in the sample (typically on the order of 10^7 — 10^8), and typically 10^2 — 10^3 **iterations**


$$\mathcal{L}(\eta; \mathcal{F}, \mathcal{T}) = \prod_{f \in \mathcal{F}} \sum_{t_i \in \Omega(f)} \Pr(t_i \mid \eta) \Pr(f \mid t_i)$$



Set of transcripts where f maps/aligns

Fragment Equivalence Classes



Reads **1** & **3** both map to transcripts B & E

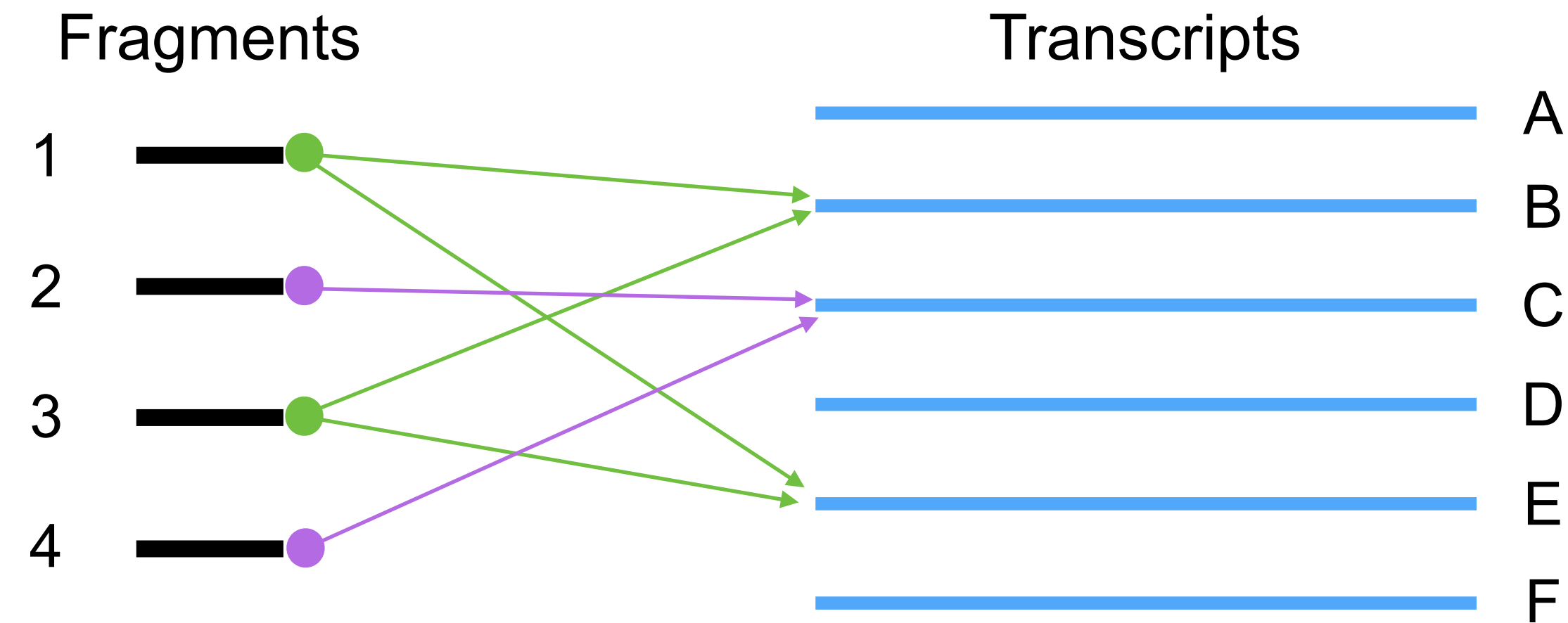
Reads **2** & **4** both map to transcript C

We have 4 reads, but only 2 eq. classes of reads

eq. Label	Count	Aux weights
{B,E}	2	$w^{\{B,E\}}_B, w^{\{B,E\}}_E$
{C}	2	$w^{\{C\}}_C$

This idea goes quite far back in the RNA-seq literature; at least to MMSeq (Turro et al. 2011)

Fragment Equivalence Classes



Reads **1** & **3** both map to transcripts B & E
 Reads **2** & **4** both map to transcript C

w_{ij} encodes the “affinity” of class j to transcript i according to the model. This is $P\{f_j | t_i\}$, aggregated for all fragments in a class.

We have 4 reads, but only 2 eq. classes of reads

eq. Label	Count	Aux weights
{B,E}	2	$w^{\{B,E\}}_B, w^{\{B,E\}}_E$
{C}	2	$w^{\{C\}}_C$

This idea goes quite far back in the RNA-seq literature; at least to MMSeq (Turro et al. 2011)

The number of equivalence classes is **small**

	Yeast	Human	Chicken
# contigs	7353	107,389	335,377
# samples	6	6	8
Total (paired-end) reads	~36,000,000	~116,000,000	~181,402,780
Avg # eq. classes (across samples)	5197	100,535	222,216

The **# of equivalence classes grows with the complexity of the transcriptome** — independent of the # of sequence fragments.

Typically, ***two or more orders of magnitude*** fewer equivalence classes than sequenced fragments.

The offline **inference** algorithm **scales in # of fragment equivalence classes**.

This naturally handles different types of multi-mapping *without* having to rely on the annotation

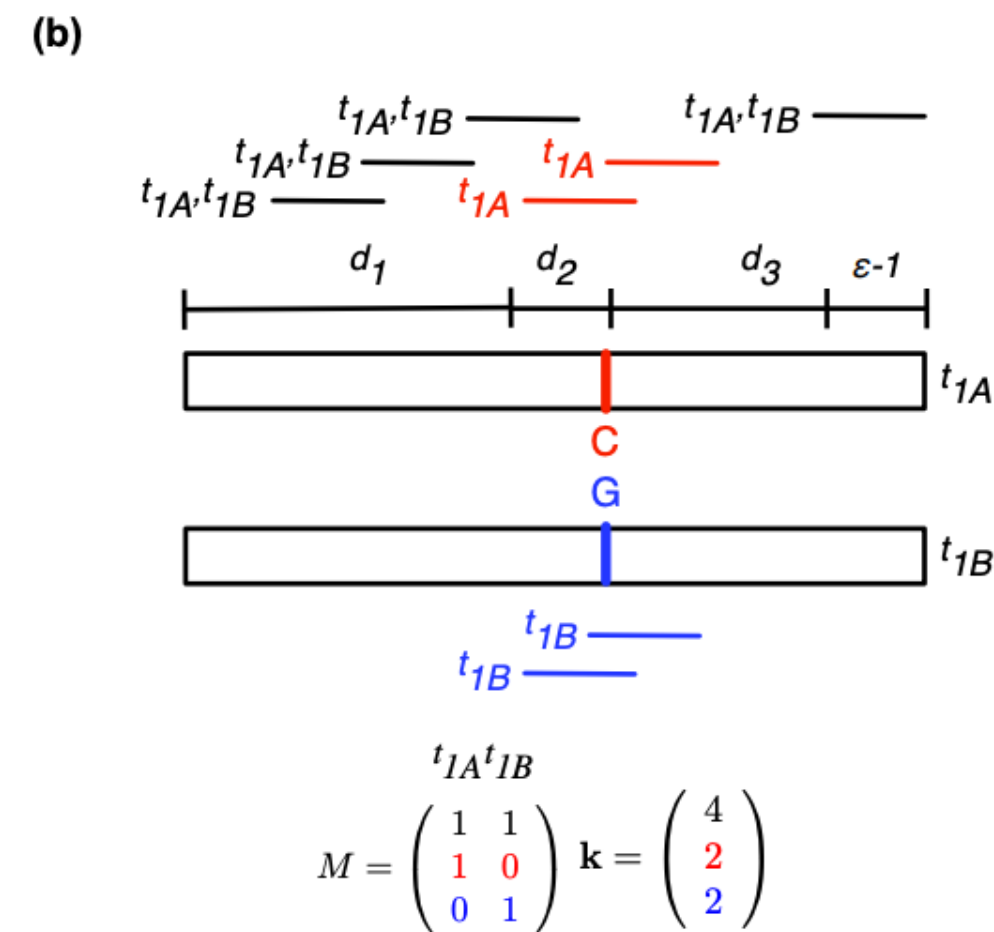
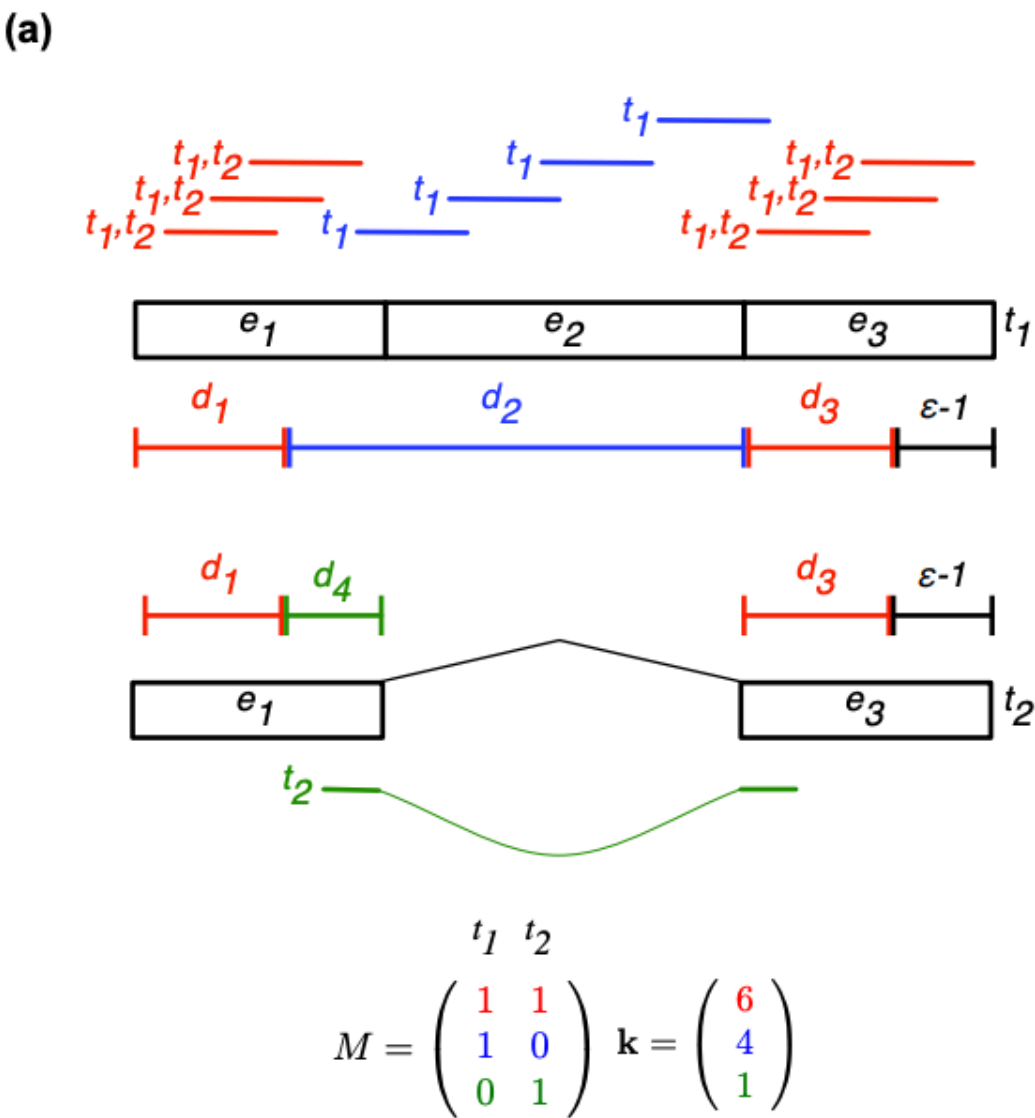


Figure 2 from Turro, Ernest, et al. "Haplotype and isoform specific expression estimation using multi-mapping RNA-seq reads." *Genome biology* 12.2 (2011): R13.

This lets us approximate the likelihood efficiently

Approximate this:

$$\mathcal{L}(\boldsymbol{\eta}; \mathcal{F}) = \prod_{f_j \in \mathcal{F}} \sum_{i=1}^M \Pr(t_i \mid \boldsymbol{\eta}) \Pr(f_j \mid t_i)$$

sum over all alignments of fragment

product over all fragments

with this:

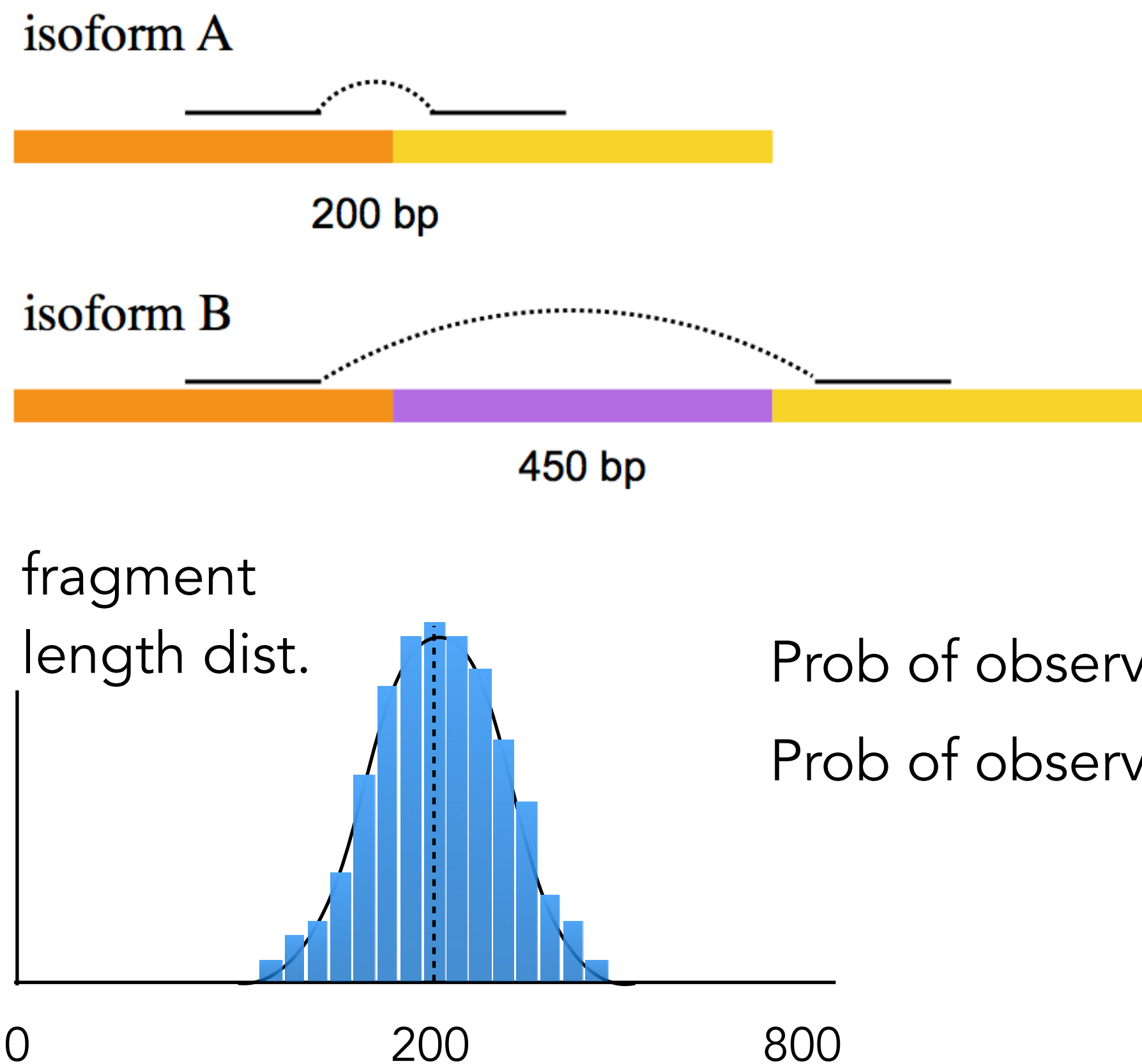
$$\mathcal{L}(\boldsymbol{\eta}; \mathcal{F}) \approx \prod_{\mathcal{F}^q \in \mathcal{C}} \left(\sum_{\langle i, t_i \rangle \in \Omega(\mathcal{F}^q)} \Pr(t_i \mid \boldsymbol{\eta}) \cdot \Pr(f \mid \mathcal{F}^q, t_i) \right)^{N^q}$$

sum over all transcripts labeling this eq. class

product over all equivalence classes

Why might $\Pr(f_j | t_i)$ matter?

Consider the following scenario:



Conditional probabilities can provide valuable information about origin of a fragment! ***Potentially different for each transcript/fragment pair.***

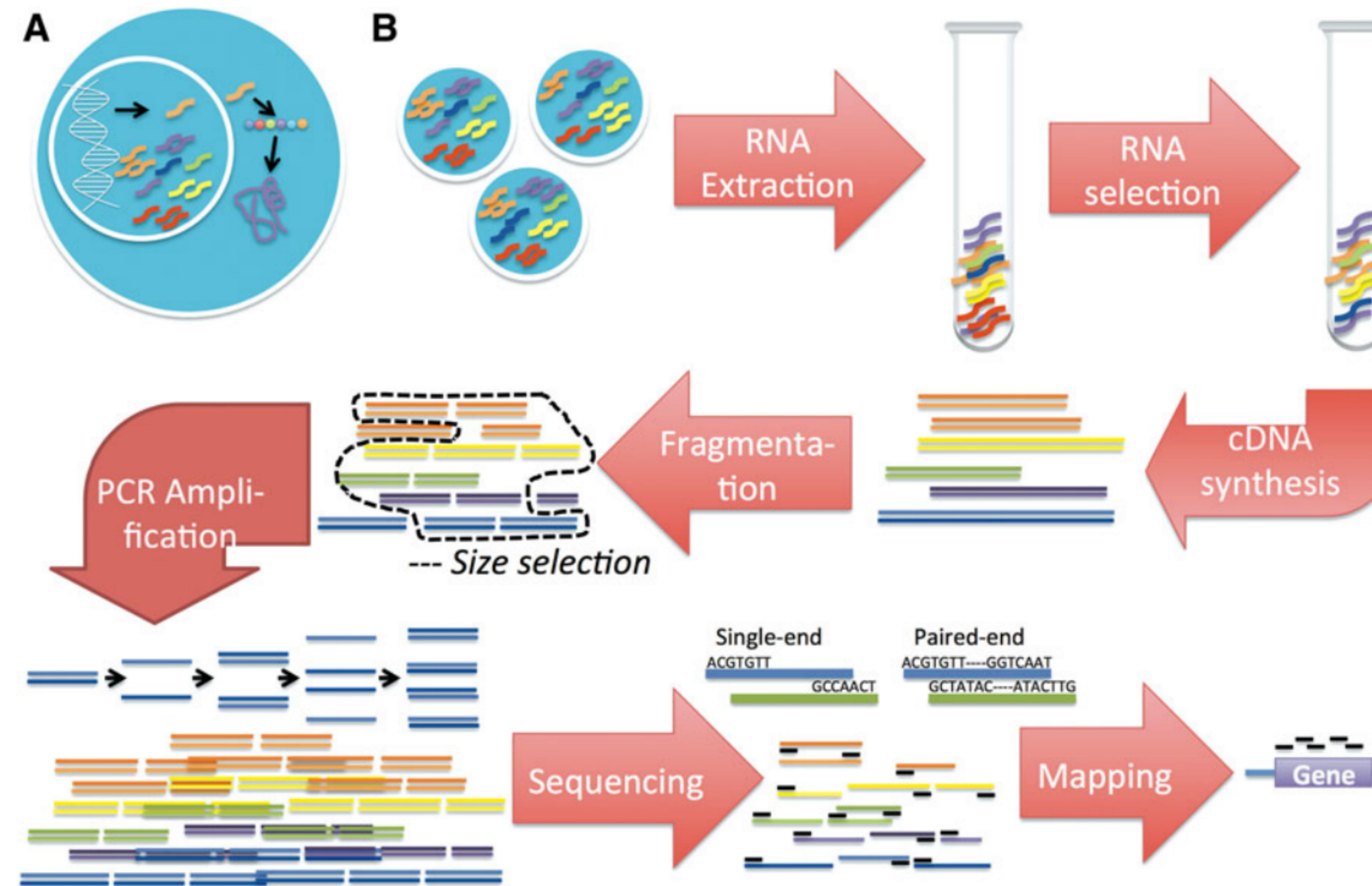
Prob of observing a fragment of size ~ 200 is **large**

Prob of observing a fragment of size ~ 450 is **small**

**Many terms can be considered in a general “fragment-transcript agreement” model¹.
e.g. position, orientation, alignment path etc.**

¹ “Salmon provides fast and bias-aware quantification of transcript expression”, Nature Methods 2017

Actual RNA-seq protocols are a bit more “involved”



There is **substantial** potential for biases and deviations from the *basic* model — indeed, we see quite a few.

Biases abound in RNA-seq data

Biases in prep & sequencing
can have a significant effect on the
fragments we see:

Fragment gc-bias¹—

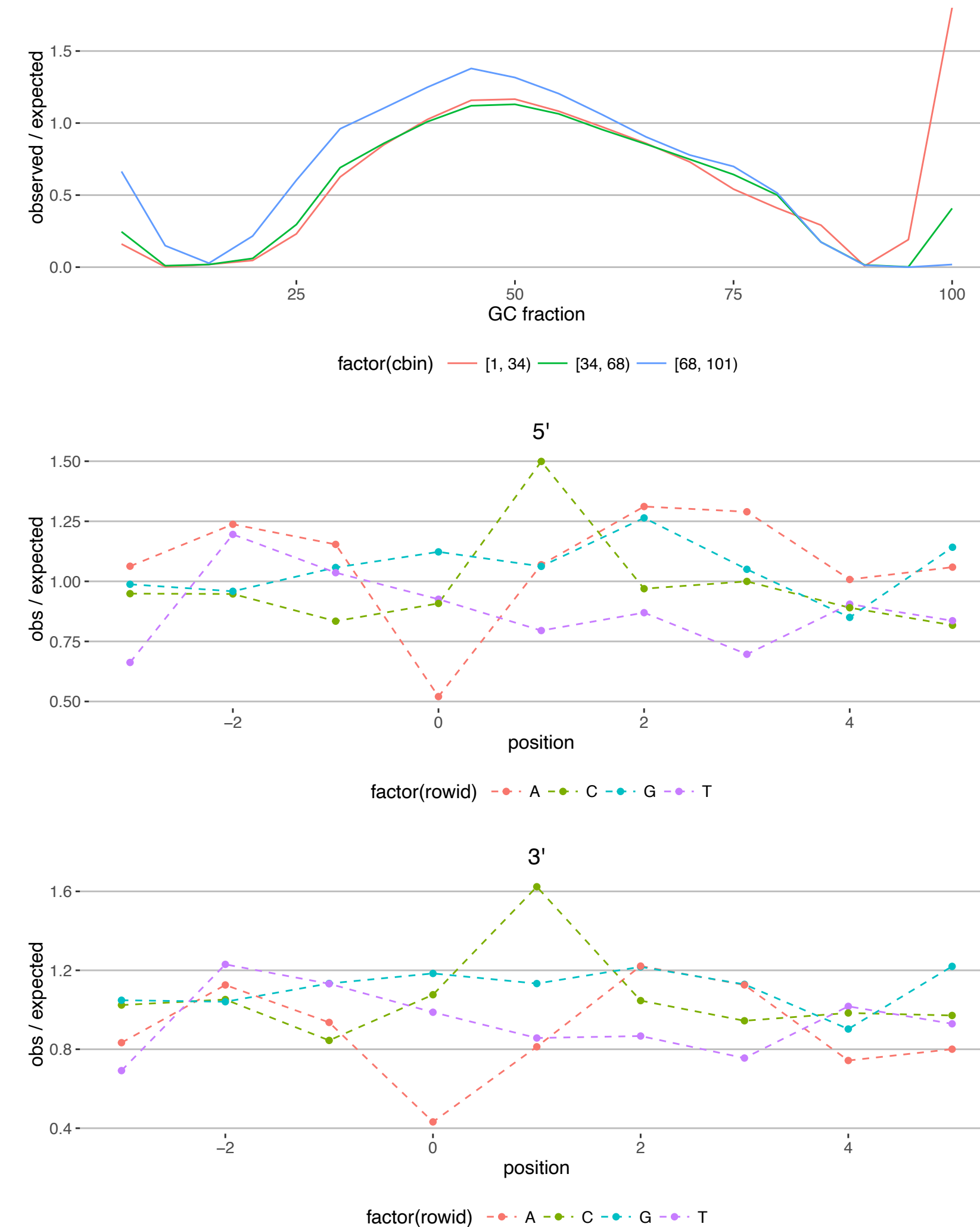
The GC-content of the fragment
affects the likelihood of sequencing

Sequence-specific bias²—

sequences surrounding fragment
affect the likelihood of sequencing

Positional bias²—

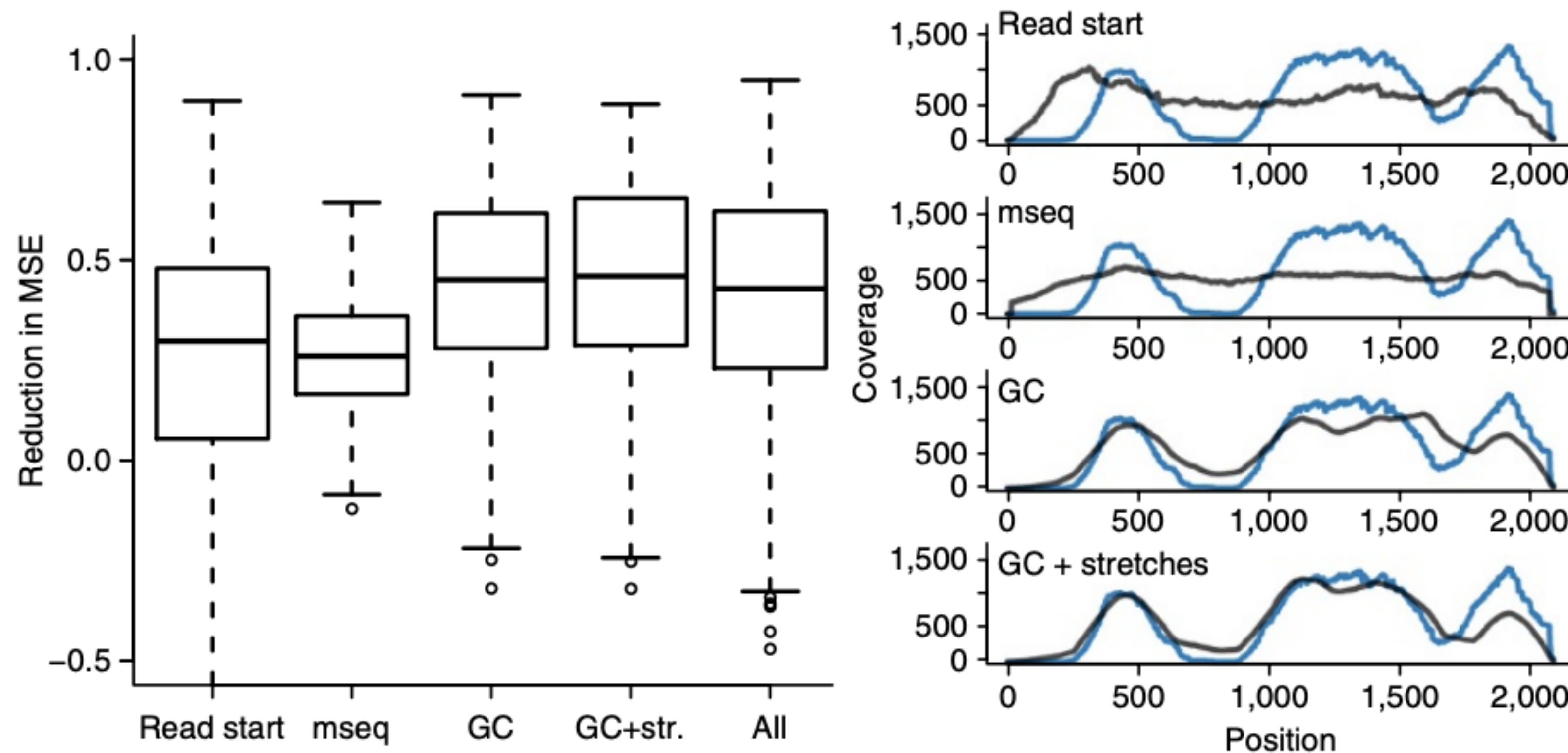
fragments sequenced non-uniformly
across the body of a transcript



1:Love, Michael I., John B. Hogenesch, and Rafael A. Irizarry. "Modeling of RNA-seq fragment sequence bias reduces systematic errors in transcript abundance estimation." bioRxiv (2015): 025767.

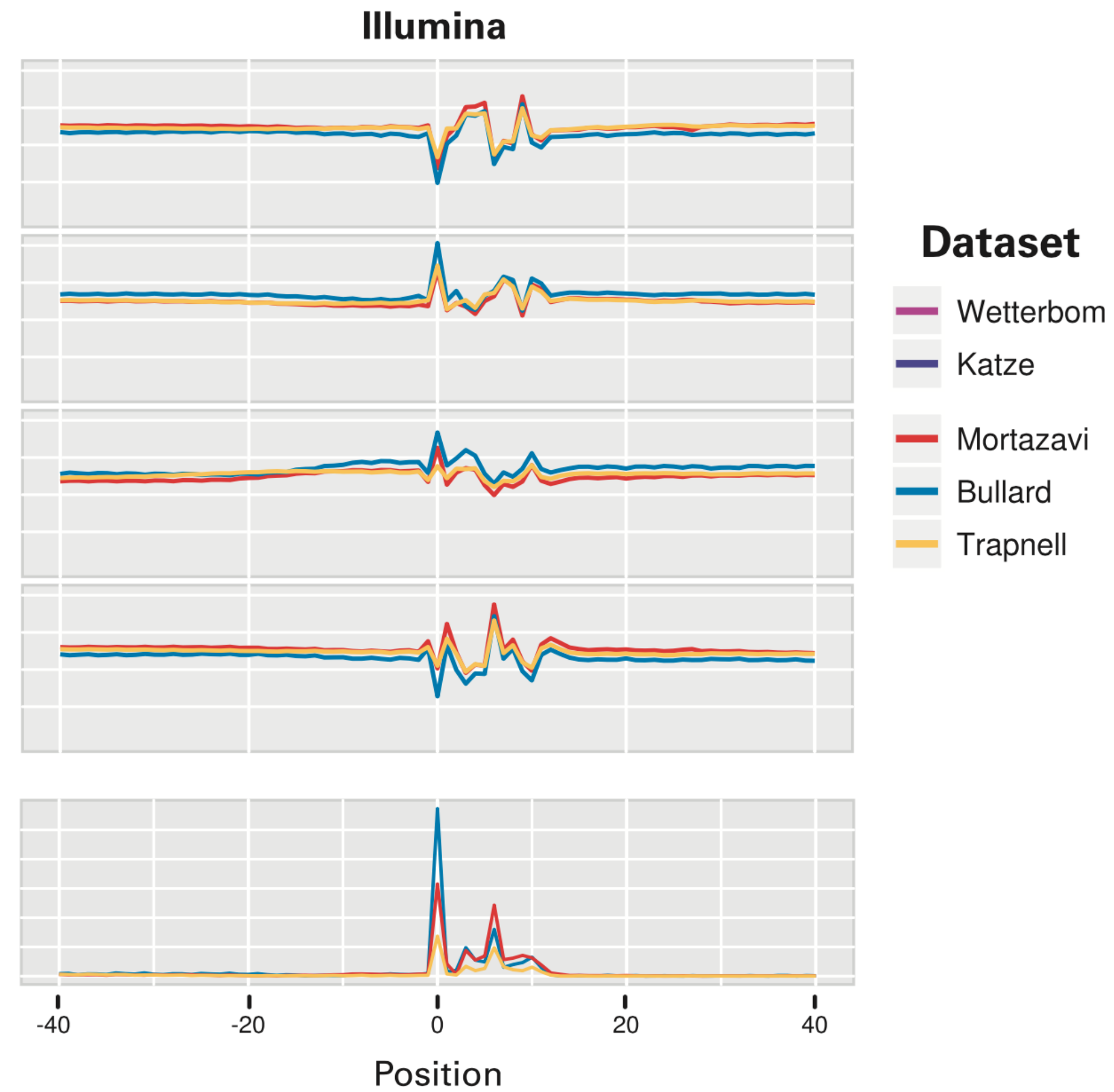
2:Roberts, Adam, et al. "Improving RNA-Seq expression estimates by correcting for fragment bias." Genome biology 12.3 (2011): 1.

Biases abound in RNA-seq data



Fragment GC-bias is often the most extreme

Priming bias is sample & sequence-specific



Biases abound in RNA-seq data

Basic idea (1): Modify the “effective length” of a transcript to account for changes in the sampling probability. This leads to changes in soft-assignment in EM -> changes in TPM.

Basic idea (2): The effective length of a transcript is the sum of the bias terms at each position across a transcript. The bias term at a given position is simply the (observed / expected) sampling probability.

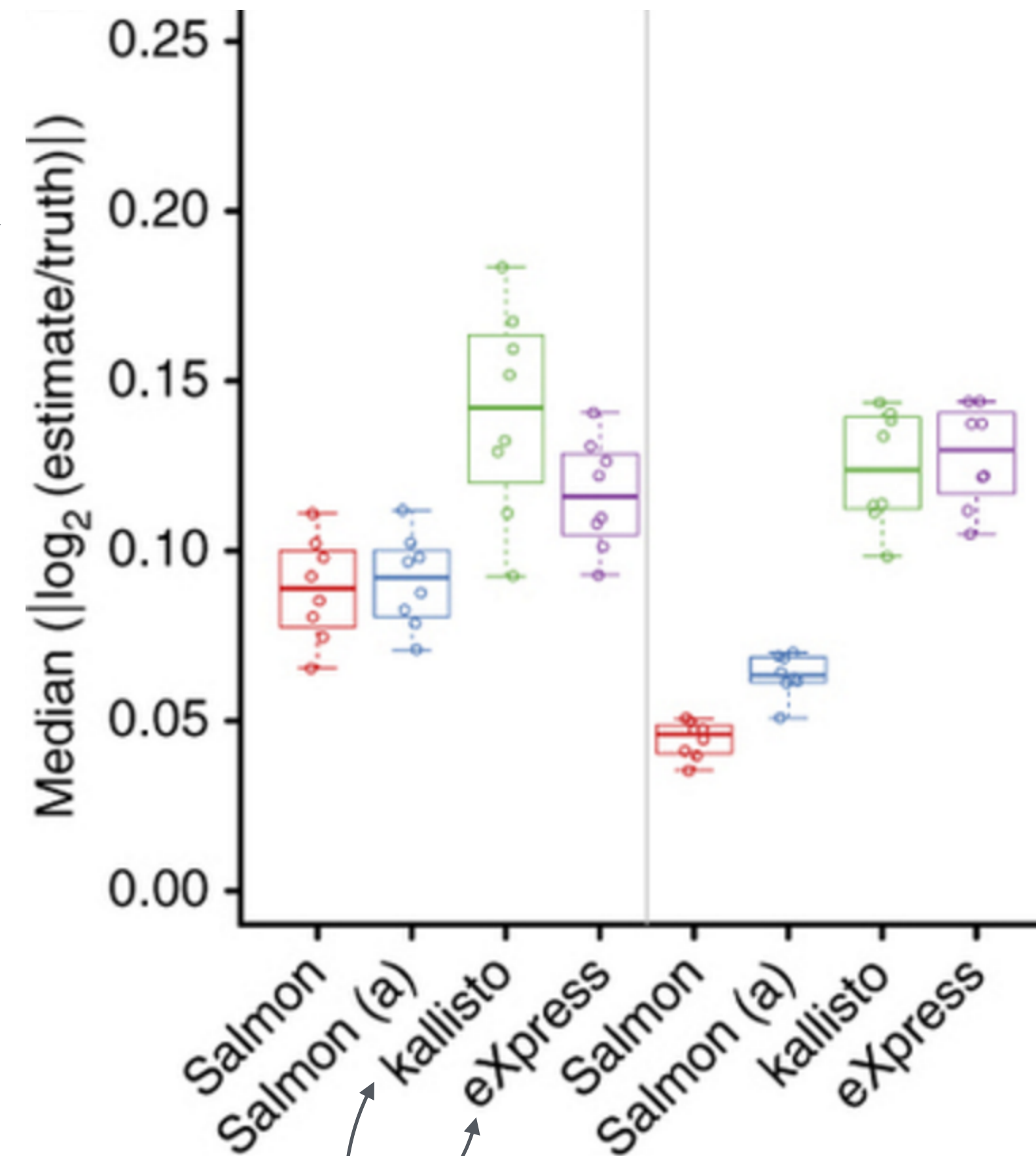
The trick is how to define “expected” given only biased data.

Accuracy difference can be larger with biased data

Simulated data:
2 conditions; 8 samples each

- Simulated transcripts across entire genome with known abundance using Polyester (modified to account for GC bias)
- How well do we recover the underlying relative abundances?
- How does accuracy vary with level of bias?

Lower is better



Sequence-bias models don't account for fragment-level GC bias

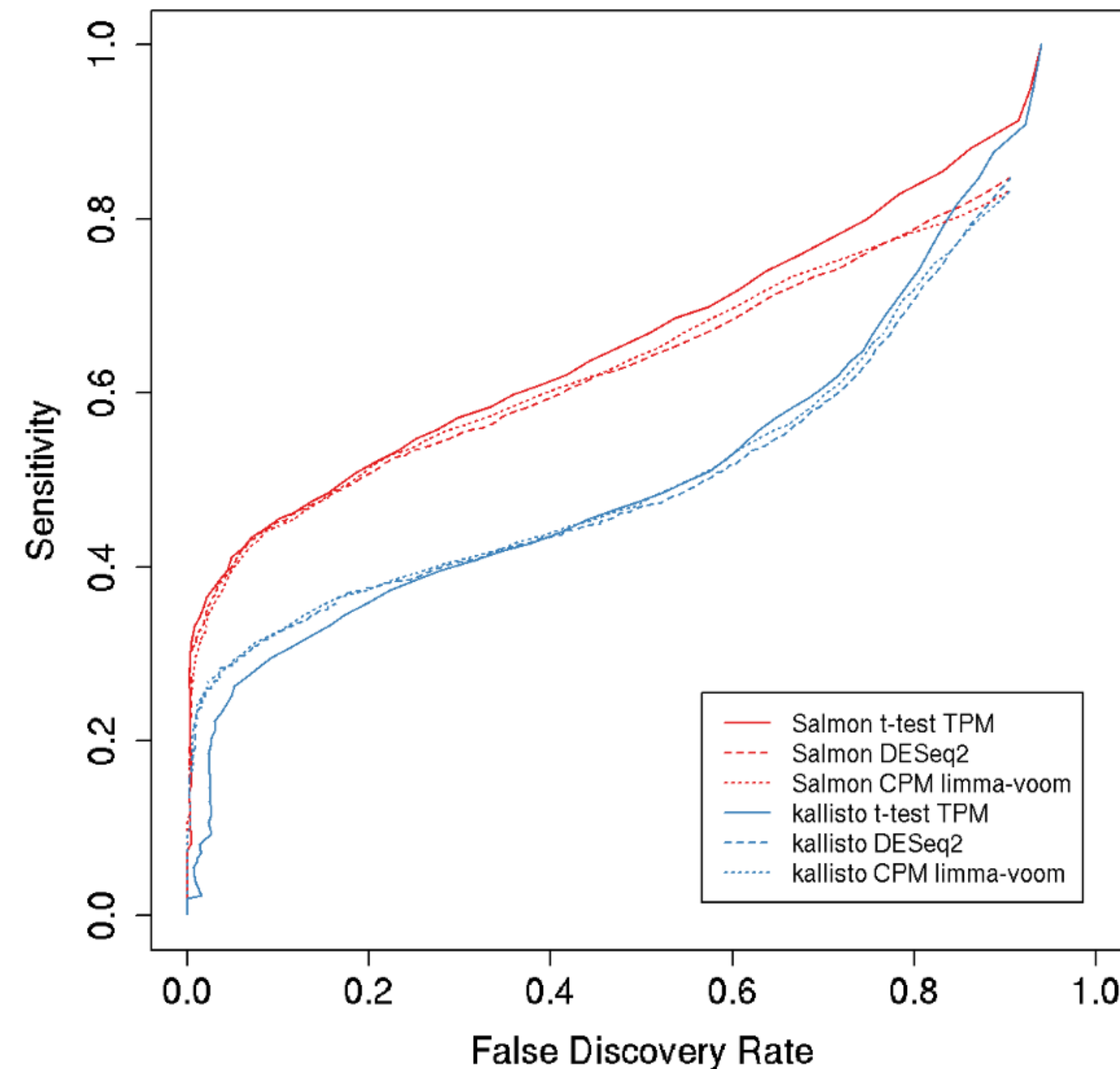
Mis-estimates confound downstream analysis

Simulated data:

2 conditions; 8 replicates each

- set 10% of txps to have fold change of 1/2 or 2 — rest unchanged.
- How well do we recover true DE?
- Since bias is systematic, effect may be even worse than accuracy difference suggests.

Recovery of DE transcripts



Importance with **experimental** data

30 samples from the GEUVADIS study:

15 samples from UNIGE sequencing center

15 samples from CNAG_CRG sequencing center

Same human population, expect few-to-no *real* DE

Randomized condition assignments result in $\ll 1$ DE txp

DE of data between centers (FDR < 1%) (TPM > 0.1)

	Salmon	Kallisto	eXpress
All transcripts	1,183	2,620	2,472
Transcripts of 2 isoform genes	228	545	531

Bias and batch effects are *substantial*, and must be accounted for.

Importance with **experimental** data

30 samples from the GEUVADIS study:

15 samples from UNIGE sequencing center

15 samples from CNAG_CRG sequencing center

Effects seem **at least as extreme** at the gene level

DE of data between centers (FDR < 1%) (TPM > 0.1)

	Salmon	Kallisto	eXpress
All genes	455	1,200	1,582
Transcripts of 2 isoform genes	224	545	531

Bias and **batch effects** are *substantial*, and must be accounted for.

Estimating Posterior Uncertainty

One “issue” with maximum likelihood (ML)

The generative statistical model is a principled and elegant way to represent the RNA-seq process.

It can be optimized efficiently using e.g. the EM / VBEM algorithm.

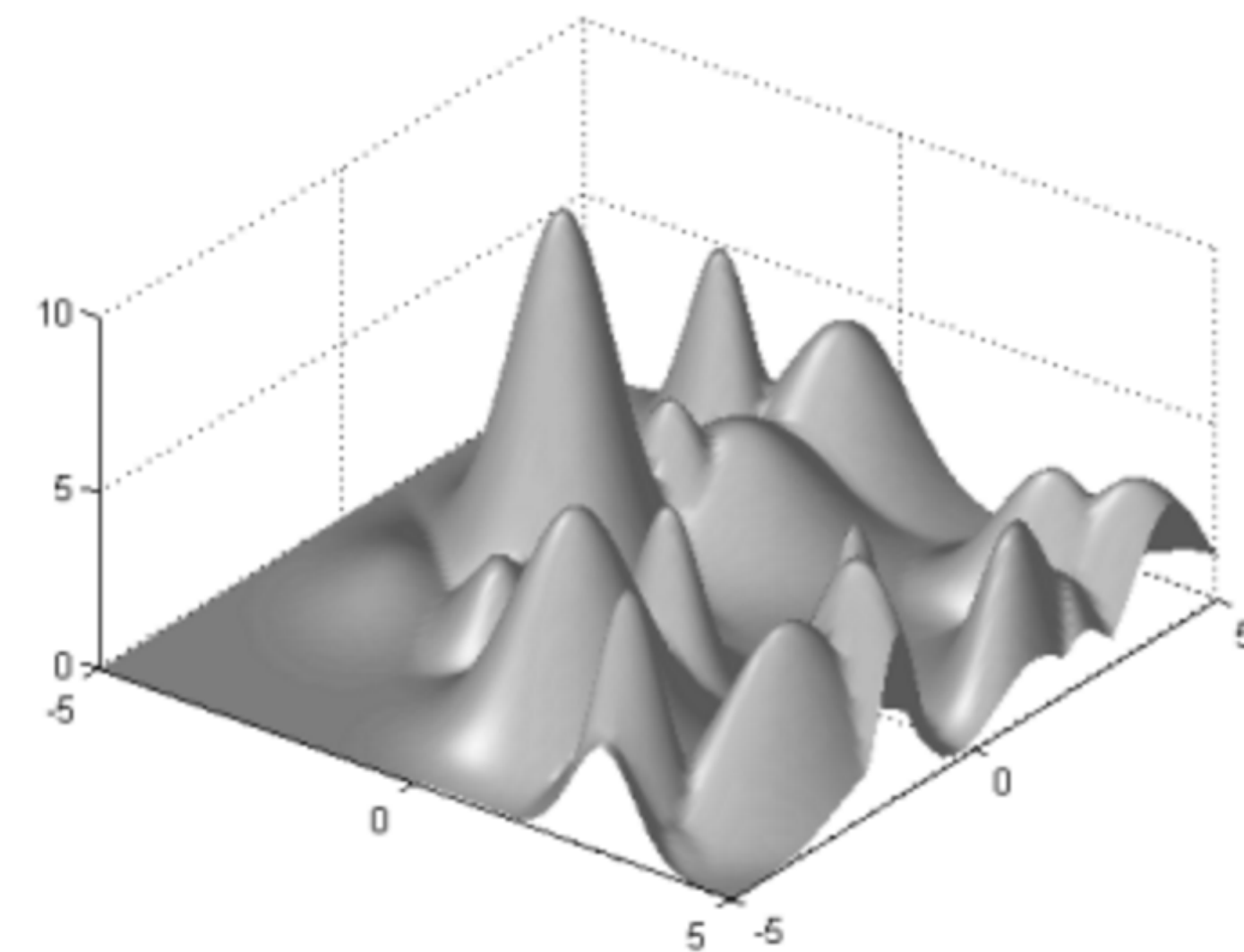
but, these efficient optimization algorithms return “point estimates” of the abundances. That is, there is no notion of how *certain* we are in the computed abundance of transcript.

One “issue” with maximum likelihood (ML)

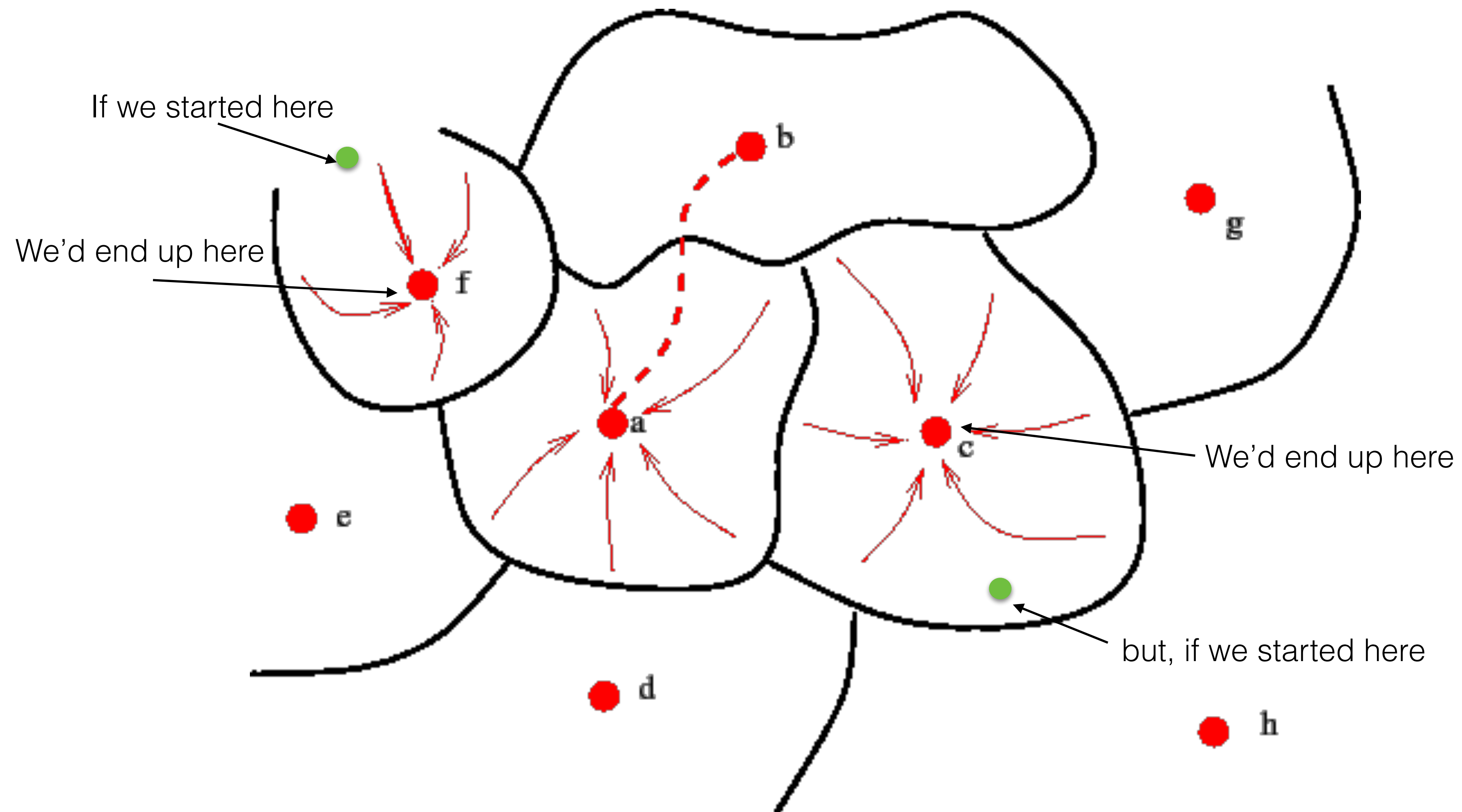
There are multiple sources of uncertainty e.g.

- Technical variance : If we sequenced the *exact* same sample again, we’d get a different set of fragments, and, potentially a different solution.
- Uncertainty in inference: We are almost never guaranteed to find a unique, globally optimal result. If we started our algorithm with different initialization parameters, we might get a different result.

We’re trying to find the *best* parameters in a space with 10s to 100s of thousands of dimensions!



One “issue” with maximum likelihood (ML)



Assessing Uncertainty

There are a few ways to address this “issue”

Do a fully Bayesian inference¹:

Infer the entire posterior distribution of parameters, not just a ML estimate (e.g. using MCMC) — too slow!

✓ Posterior Gibbs Sampling^{2,3}:

Starting from our ML estimate, do MCMC sampling to explore how parameters vary — if our ML estimate is good, this can be made *quite fast*.

✓ Bootstrap Sampling⁴:

Resample (from range-factorized equivalence class counts) with replacement, and re-run the ML estimate for each sample. This can be made reasonably fast.

1: BitSeq (with MCMC) actually does this. It's very accurate, but very slow. [Glaus, Peter, Antti Honkela, and Magnus Rattray. "Identifying differentially expressed transcripts from RNA-seq data with biological variation." *Bioinformatics* 28.13 (2012): 1721-1728.]

2: RSEM has the ability to do this, and it seems to work well, but each sample scales in the # of reads. [Li, Bo, and Colin N. Dewey. "RSEM: accurate transcript quantification from RNA-Seq data with or without a reference genome." *BMC bioinformatics* 12.1 (2011): 1.]

3: MMSEQ can perform Gibbs sampling over shared variables (i.e. equiv classes), producing estimates from the mean of the posterior dist. Turro, Ernest, et al. "Haplotype and isoform specific expression estimation using multi-mapping RNA-seq reads." *Genome biology* 12.2 (2011): 1.

4: IsoDE introduced the idea of bootstrapping counts to assess quantification uncertainty. [Al Seesi, Sahar, et al. "Bootstrap-based differential gene expression analysis for RNA-Seq data with and without replicates." *BMC genomics* 15.8 (2014): 1.], but it was first made practical / fast in kallisto by doing the bootstrapping over equivalence classes.

This uncertainty matters

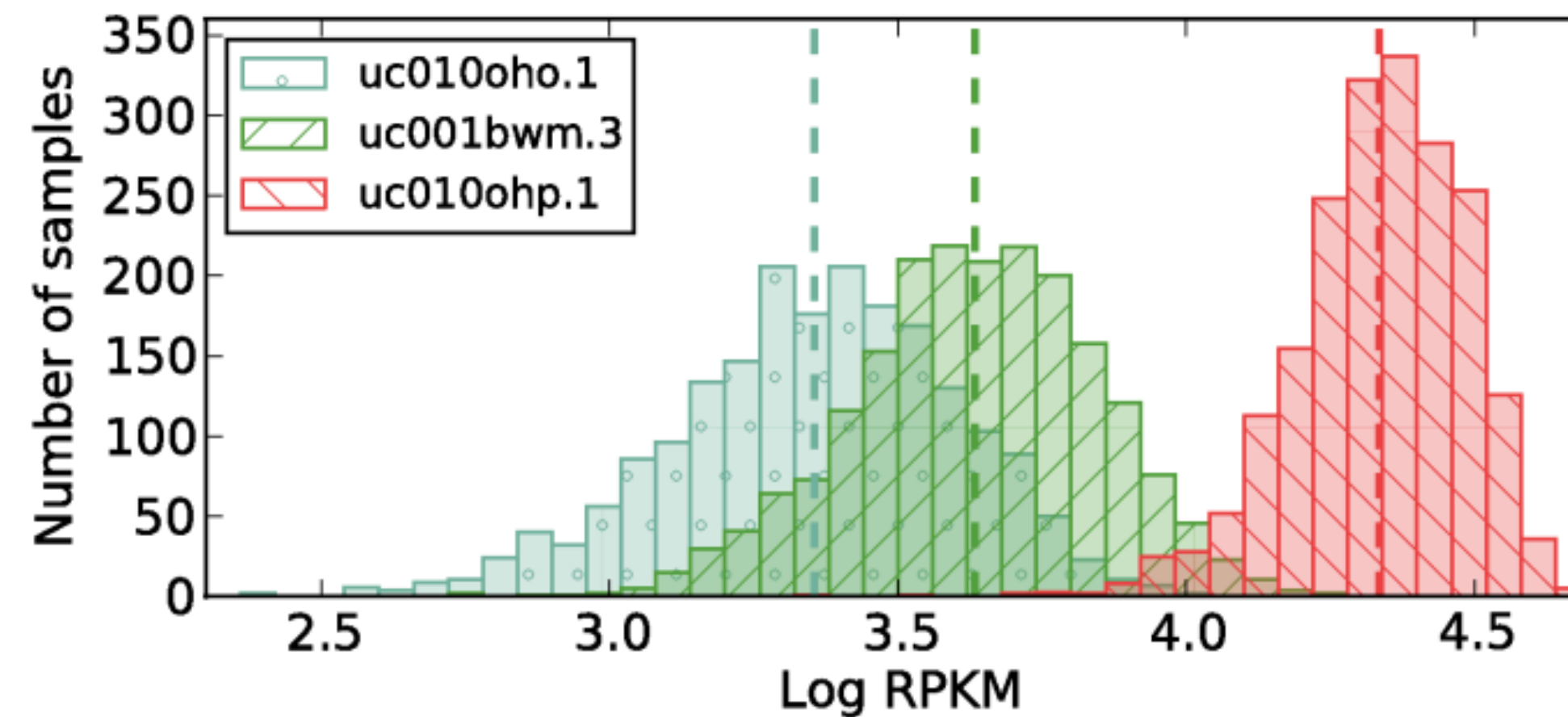
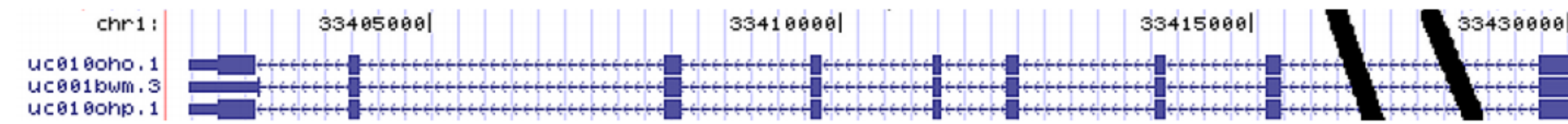
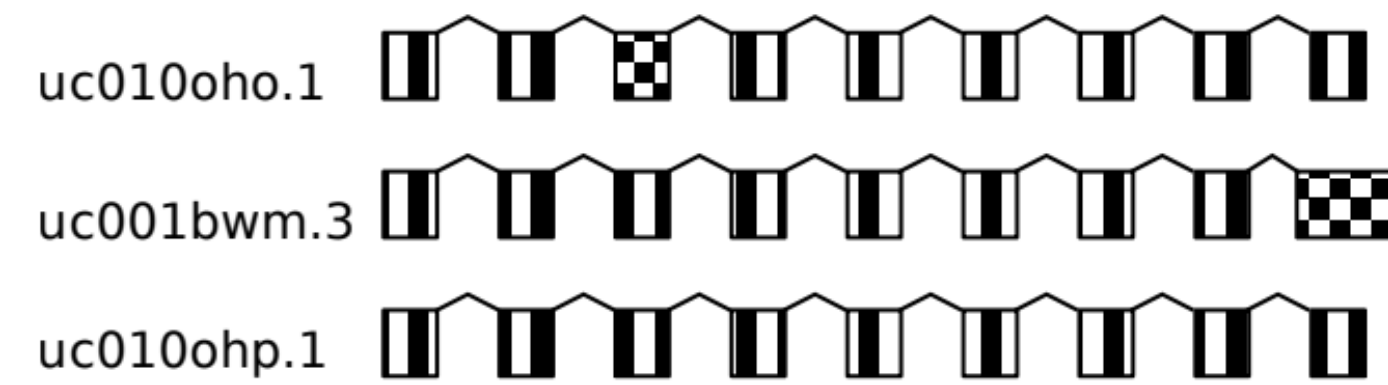


Figure 2.10: **Posterior distribution of expression levels of three transcripts of gene Q6ZMZ0.** The posterior distribution is represented in form of a histogram of expression samples converted into Log RPKM expression measure. The dashed lines mark the mean expression for each transcript.

This uncertainty matters

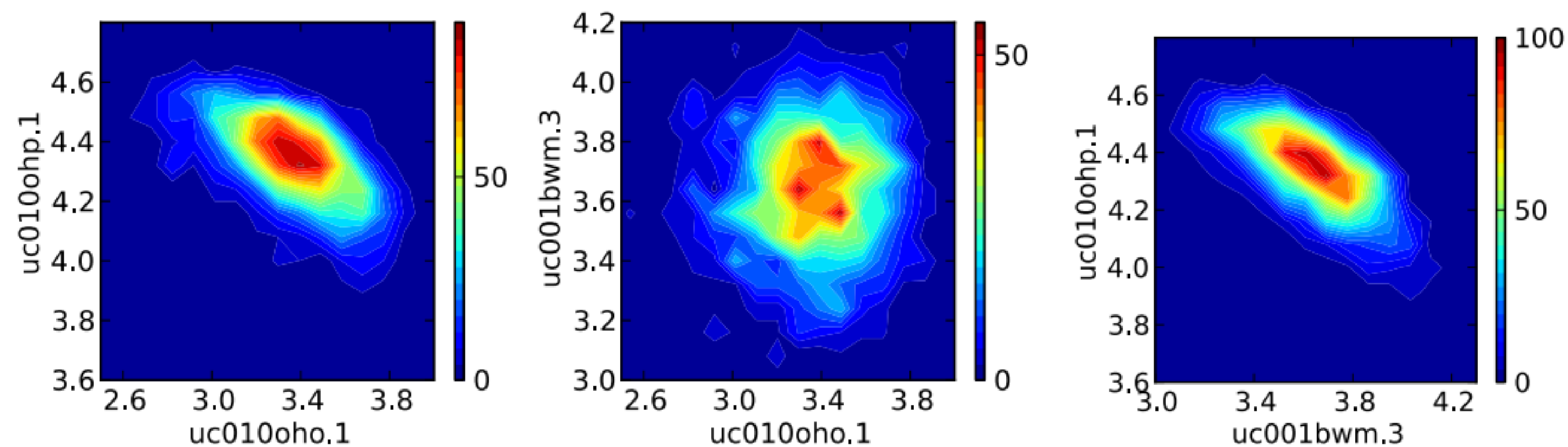


(a) Transcript sequence profile.



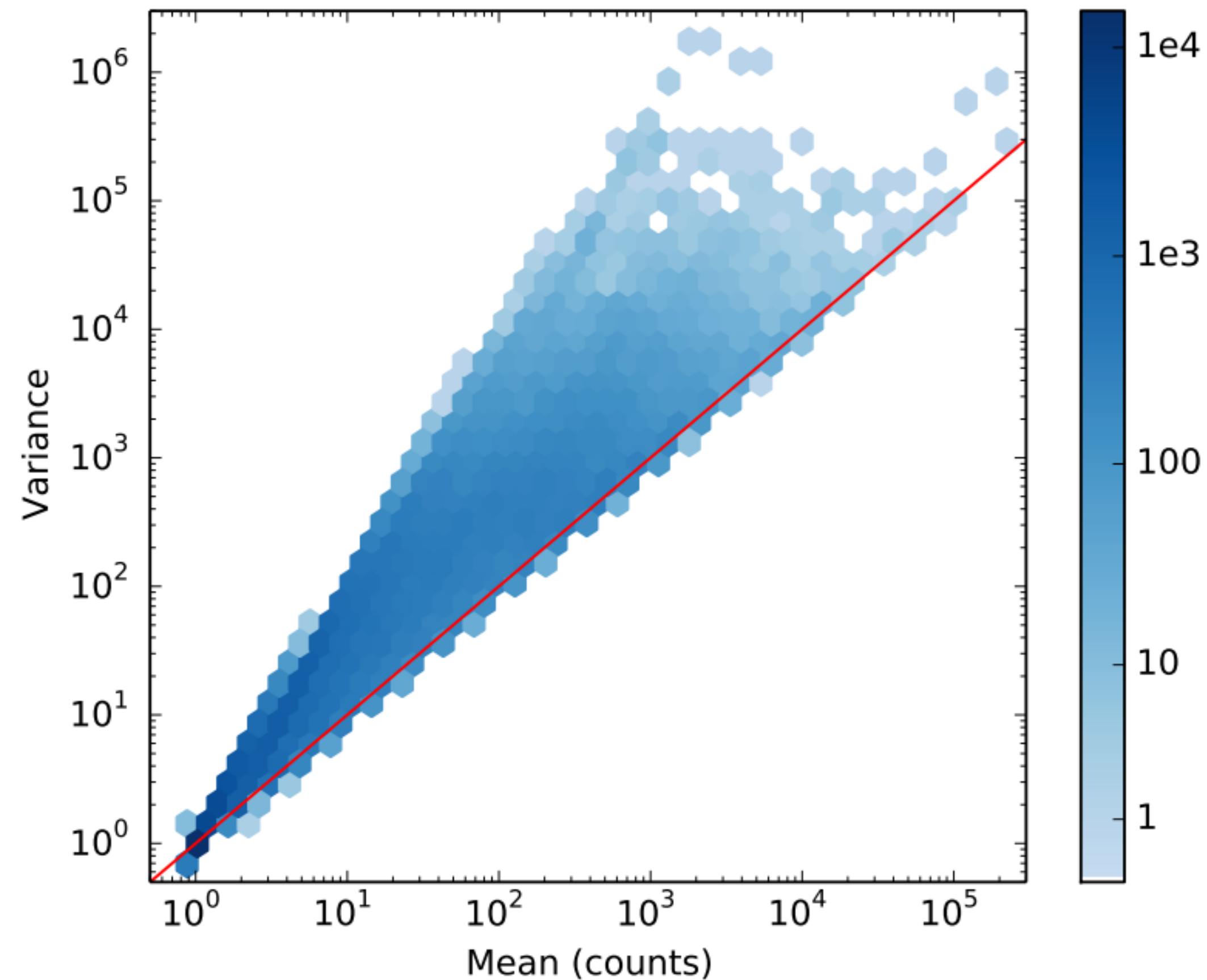
(b) Splice variant model.

Figure 2.12: **Exon model of transcripts of gene Q6ZMZ0.** (a) transcript sequence profile obtained from the UCSC genome browser (Kuhn et al., 2013). In this annotation, transcript uc001bwm.3 has different 3' untranslated region and transcript uc010oho.1 has extra nucleotides at the end of second exon. As the second change cannot be distinguished in the UCSC genome browser diagram, we provide schematic splice variant model highlighting the differences (b).



This uncertainty matters

We observe considerably increased variance due to read mapping ambiguity



If we know this increased uncertainty, we can propagate it & use it in downstream analysis (differential expression)!

Accounting for uncertainty in downstream analysis

Graphical abstract

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<https://doi.org/10.1093/nar/gkad1167>
Advance access publication date: 7 December 2023
Methods



Dividing out quantification uncertainty allows efficient assessment of differential transcript expression with edgeR

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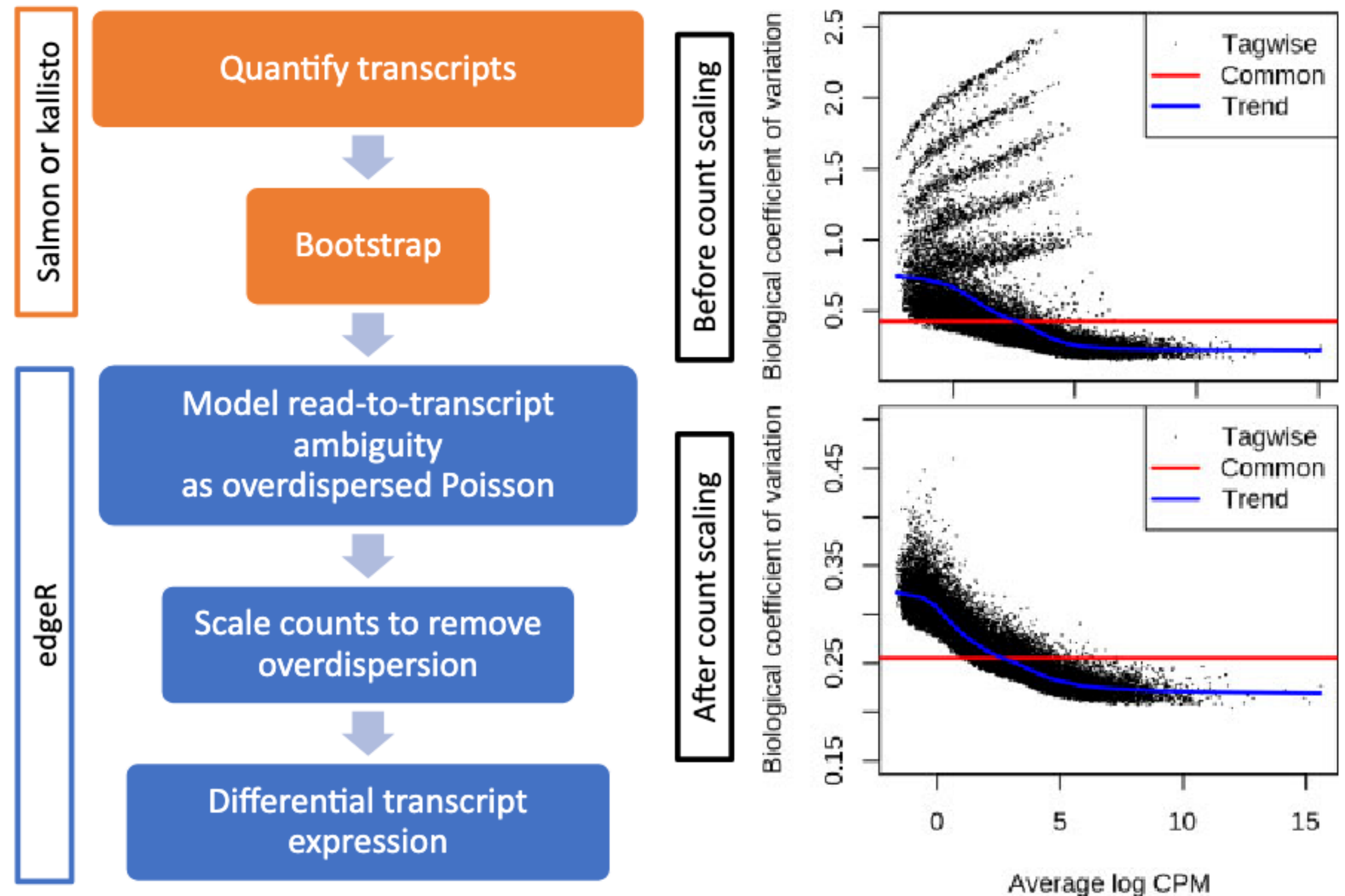
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[†]The first two authors should be regarded as Joint First Authors.



Accounting for uncertainty in downstream analysis

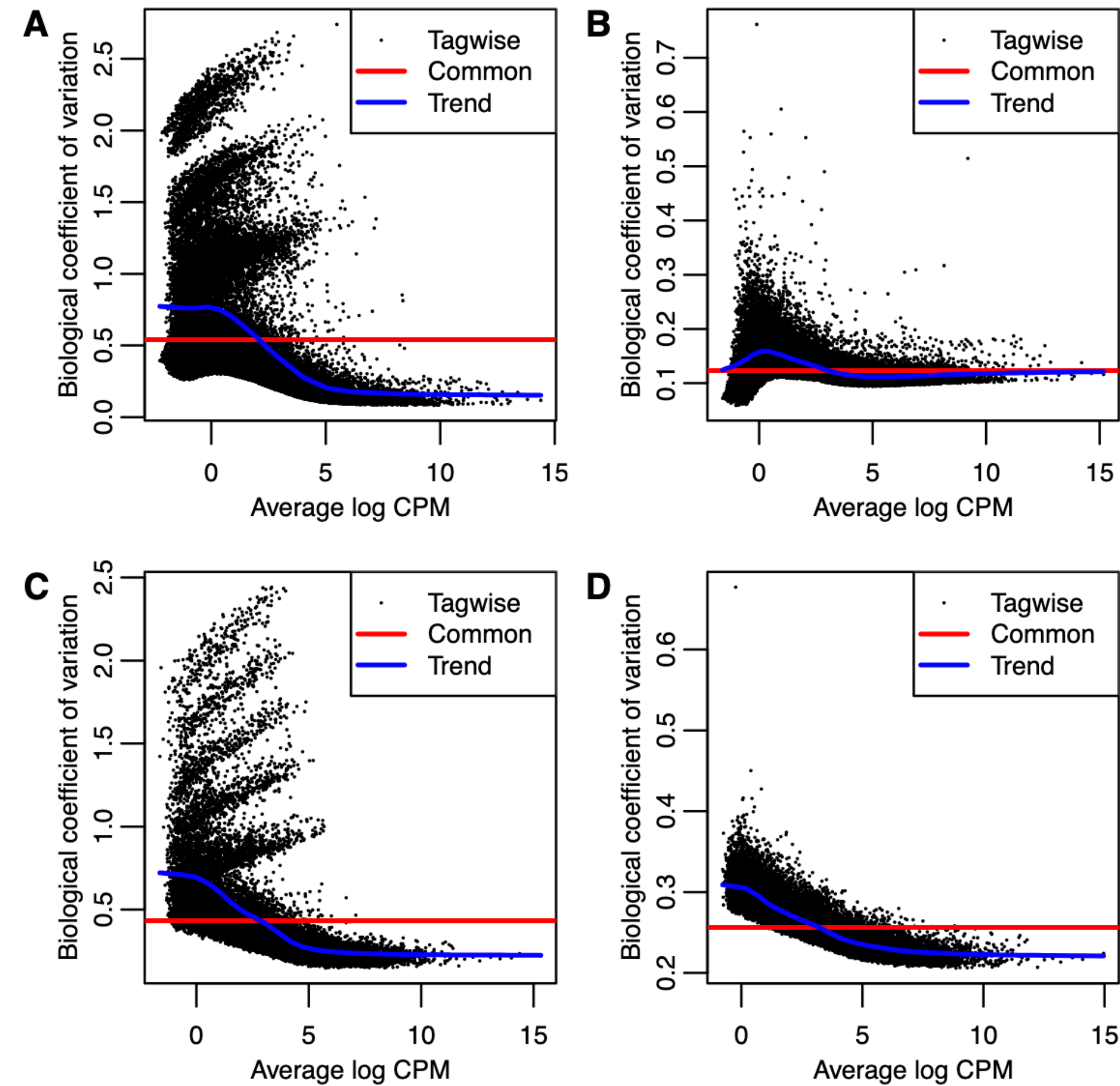


Figure 2. BCV plots of transcript-level counts from RNA-seq data. Panels (A) and (B) show results from the Illumina short paired-end reads RNA-seq experiment of the lung adenocarcinoma cell lines before and after count scaling, respectively. Panels (C) and (D) show results from a simulated mouse RNA-seq experiment before and after count scaling, respectively. Simulated data was generated with 100 bp paired-end reads, unbalanced library sizes, five samples per group, and all reference transcripts expressed. For both real and simulated data, experiments were quantified with *Salmon*.

Bonus Material

Bias Modeling

Bias correction works by adjusting the effective lengths of the transcripts:
The effective length becomes the sum of the per-base biases

$$\tilde{\ell}'_i = \sum_{j=1}^{j \leq \ell_i} \sum_{k=1}^{k \leq f_i(j,L)} \frac{b_{gc+}(t_i, j, j+k)}{b_{gc-}(t_i, j, j+k)} \cdot \frac{b_{s+}^{5'}(t_i, j)}{b_{s-}^{5'}(t_i, j)} \cdot \frac{b_{s+}^{3'}(t_i, j+k)}{b_{s-}^{3'}(t_i, j+k)} \cdot \frac{b_{p+}^{5'}(t_i, j+k)}{b_{p-}^{5'}(t_i, j+k)} \cdot \frac{b_{p+}^{3'}(t_i, j+k)}{b_{p-}^{3'}(t_i, j+k)} \cdot \Pr\{X = j\}$$

Fragment GC bias model:

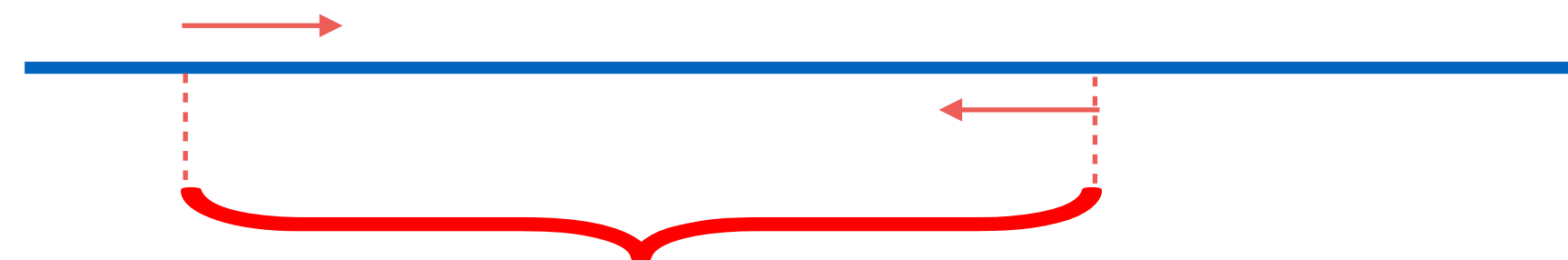
Density of fragments with specific GC content,
conditioned on GC fraction at read start/end

Foreground:

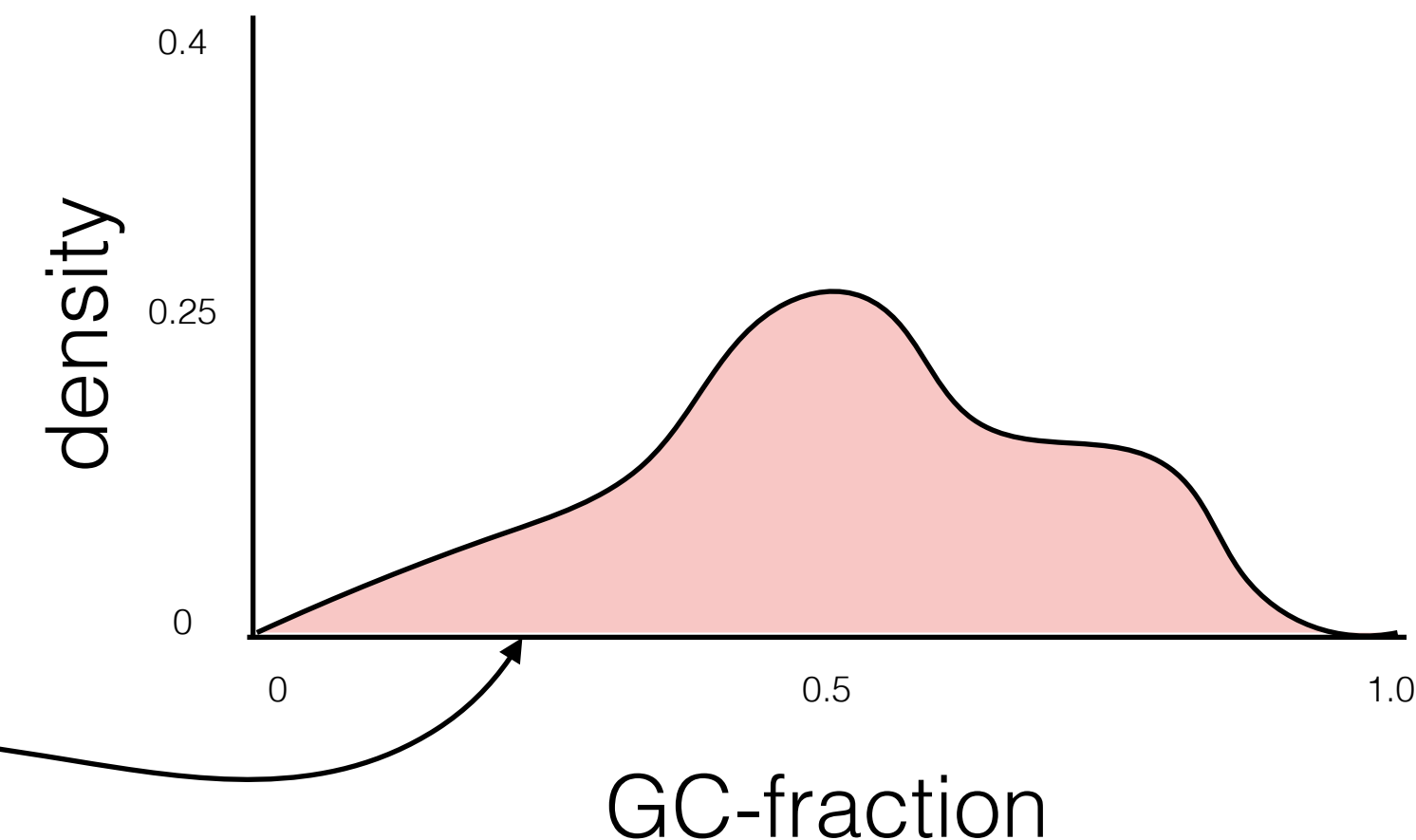
Observed

Background:

Expected given est. abundances



GC-fraction of fragment



Bias Modeling

Bias correction works by adjusting the effective lengths of the transcripts:
The effective length becomes the sum of the per-base biases

$$\tilde{\ell}'_i = \sum_{j=1}^{j \leq \ell_i} \sum_{k=1}^{k \leq f_i(j,L)} \frac{b_{gc+}(t_i, j, j+k)}{b_{gc-}(t_i, j, j+k)} \cdot \frac{b_{s+}^{5'}(t_i, j)}{b_{s-}^{5'}(t_i, j)} \cdot \frac{b_{s+}^{3'}(t_i, j+k)}{b_{s-}^{3'}(t_i, j+k)} \cdot \frac{b_{p+}^{5'}(t_i, j+k)}{b_{p-}^{5'}(t_i, j+k)} \cdot \frac{b_{p+}^{3'}(t_i, j+k)}{b_{p-}^{3'}(t_i, j+k)} \cdot \Pr\{X = j\}$$

Seq-specific bias model*:

VLMM for the 10bp window surrounding the 5' read start site and the 3' read start site

Foreground:

Observed

Background:

Expected given est. abundances



Add this sequence to training set with weight =
 $P\{f \mid t_i\}$

Same, but independent
model for 3' end

Bias Modeling

Bias correction works by adjusting the effective lengths of the transcripts:
The effective length becomes the sum of the per-base biases

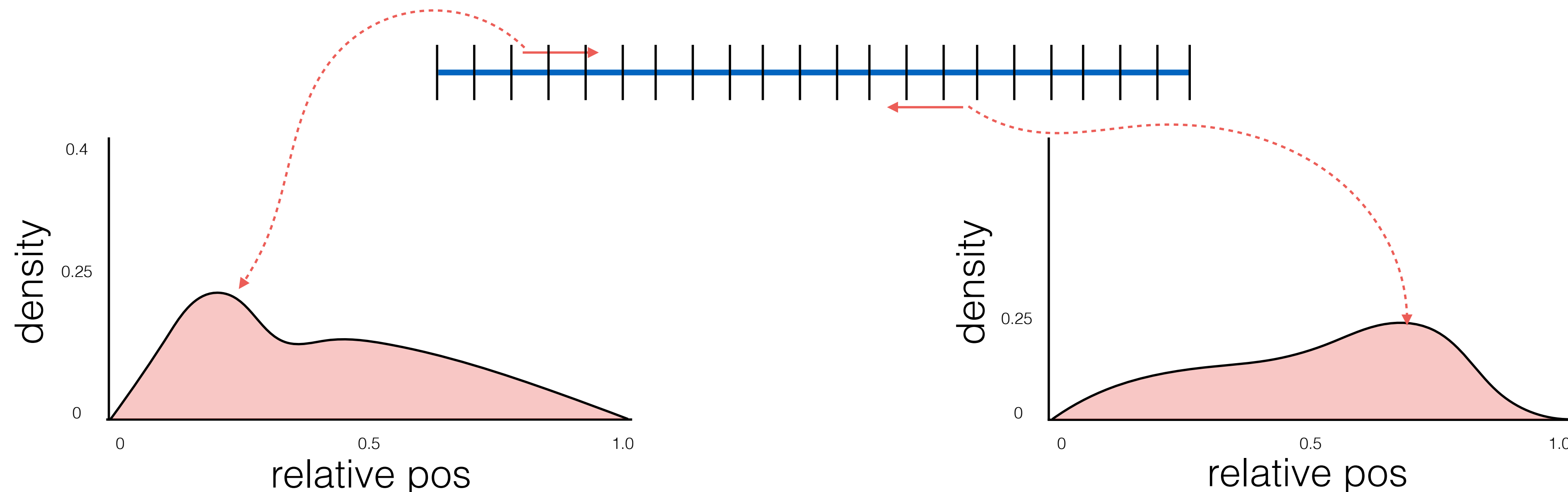
$$\tilde{\ell}'_i = \sum_{j=1}^{j \leq \ell_i} \sum_{k=1}^{k \leq f_i(j,L)} \frac{b_{gc+}(t_i, j, j+k)}{b_{gc-}(t_i, j, j+k)} \cdot \frac{b_{s+}^{5'}(t_i, j)}{b_{s-}^{5'}(t_i, j)} \cdot \frac{b_{s+}^{3'}(t_i, j+k)}{b_{s-}^{3'}(t_i, j+k)} \cdot \frac{b_{p+}^{5'}(t_i, j+k)}{b_{p-}^{5'}(t_i, j+k)} \cdot \frac{b_{p+}^{3'}(t_i, j+k)}{b_{p-}^{3'}(t_i, j+k)} \cdot \Pr\{X = j\}$$

Position bias model*:

Foreground:
Observed

Density of 5' and 3' read start positions —
different models for transcripts of different length

Background:
Expected given est. abundances



*Roberts, Adam, et al. "Improving RNA-Seq expression estimates by correcting for fragment bias." Genome biology 12.3 (2011): 1.

A few ways to implement Gibbs Sampling for this problem

The model of MMSeq

$$X_{it} \mid \mu_t \sim \text{Pois}(bs_i M_{it} \mu_t), \quad (12)$$

$$\mu_t \sim \text{Gam}(\alpha, \beta). \quad (13)$$

The full conditionals are:

$$\{X_{i1}, \dots, X_{it}\} \mid \{\mu_1, \dots, \mu_t\}, k_i \sim \text{Mult} \left(k_i, \frac{M_{i1}\mu_1}{\sum_t M_{it}\mu_t}, \dots, \frac{M_{in}\mu_n}{\sum_t M_{it}\mu_t} \right), \quad (14)$$

$$\mu_t \mid \{X_{1t}, \dots, X_{mt}\} \sim \text{Gam} \left(\alpha + \sum_i X_{it}, \beta + bl_t \right). \quad (15)$$

Again, the s_i are not needed as they are absent from the full conditionals.

A few ways to implement Gibbs Sampling for this problem

The model of BitSeq

$$P(I_n|\boldsymbol{\theta}, \theta^{act}, R) = \text{Cat}(I_n|\boldsymbol{\phi}_n), \quad (10)$$

$$\phi_{n0} = P(r_n|\text{noise})(1 - \theta^{act})/Z_n^{(\phi)},$$

$$m \neq 0; \phi_{nm} = P(r_n|I_n)\theta_m\theta^{act}/Z_n^{(\phi)},$$

$$P(\boldsymbol{\theta}|\mathbf{I}, \theta^{act}, R) = \text{Dir}(\boldsymbol{\theta} | (\alpha^{dir} + C_1, \dots, \alpha^{dir} + C_M)), \quad (11)$$

$$P(\theta^{act}|\mathbf{I}, \boldsymbol{\theta}, R) = \text{Beta}(\theta^{act} | \alpha^{act} + N - C_0, \beta^{act} + C_0), \quad (12)$$

$$C_m = \sum_{n=1}^N \delta(I_n = m).$$

A few ways to implement Gibbs Sampling for this problem

The model of BitSeq (collapsed sampler)

$$P(I_n | I^{(-n)}, R) = \text{Cat}(I_n | \phi_n^*), \quad (9)$$

$$\phi_{n0}^* = P(r_n | \text{noise})(\beta^{act} + C_0^{(-n)}) / Z_n^{(\phi^*)},$$

$$m \neq 0; \phi_{nm}^* = P(r_n | I_n)(\alpha^{act} + C_+^{(-n)}) \frac{(\alpha^{dir} + C_m^{(-n)})}{(M\alpha^{dir} + C_+^{(-n)})} / Z_n^{(\phi^*)},$$

$$C_m^{(-n)} = \sum_{i \neq n} \delta(I_i = m),$$

$$C_+^{(-n)} = \sum_{i \neq n} \delta(I_i > 0),$$

with $Z_n^{(\phi^*)}$ being a constant normalising ϕ_n^* to sum up to 1, and $\alpha^{dir} = 1, \alpha^{act} = 2, \beta^{act} = 2$.